DOCUMENT RESUME

ED 081 850 TM 003 211

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TITLE An Analysis of a Fleet Post-Training Performance

Measurement Technique.

INSTITUTION Naval Personnel Research and Development Lab.,

Washington, D.C.

SPONS AGENCY Office of Naval Research, Washington, D.C. Personnel

and Training Research Programs Office.

REPORT NO WTR-73-47
PUB DATE Jun 73
NOTE 180p.

EDRS PRICE MF-\$0.65 HC-\$6.58

DESCRIPTORS *Measurement Techniques: Military Personnel:

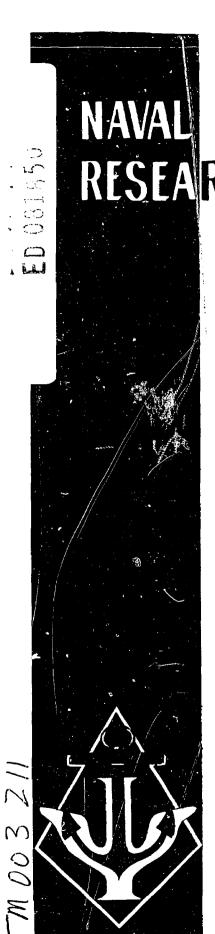
*Military Training; Performance Tests; *Statistical Analysis; *Task Performance; Technical Occupations;

Technical Reports; Test Validity

ABSTRACT

The purpose of this research effort is to validate the utility and effectiveness of a unique human performance measurement technique developed under ONR contract (N0001467C0107). Performance data on eight Navy ratings was collected from ships of LANTFLT and PACFLT. This report is the last in a series of technical reports on the statistical analysis of that data. A statistical analysis is provided on performance related data for electronic maintenance personnel sampled from 21 ships. Four different performance estimators, as functions of critical incidents, were evaluated. A detailed explanation of the distributional properties of the performance estimators is presented, and an explanation of the factors that lead to the adoption of a curvilinear regression analysis for analysis of the data is discussed. The results of the statistical analysis indicated that a certain combination of the performance data possessed moderate validity for appraising the absolute level of technician on-the-job performance on the EM, ET, FT, and IC ratings. Application of the technique to technicians in the RM, ST, and TM ratings was tenuous, but still appropriate, while none of the performance estimators seemed to be applicable to technicians in the RD rating. For this reason, it would seem that the appropriateness of application of this technique to other ratings warrants investigation, perhaps by the approach employed in this report. In any event it has been observed that the technique possesses sufficient merit to be recommended for more widespread use within the U. S. Navy. (Author)





NAVAL PERSONNEL RESEARCH AND DEVELOPMENT LABORATORY

WTR 73-47

June 1973

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AN ANALYSIS OF A
FLEET POST-TRAINING PERFORMANCE
MEASUREMENT TECHNIQUE

Bernard A. Rafacz Paul P. Foley

US DEPARTMENT OF HEALTH EDUCATION & WELFARE NATIONAL INSTITUTE OF

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WASHINGTON NAVY YARD WASHINGTON, D.C. 20390

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WTR 73-47

June 1973

AN ANALYSIS OF A FLEET POST-TRAINING PERFORMANCE MEASUREMENT TECHNIQUE

Bernard A. Rafacz Paul P. Foley

Principal Investigator: Paul P. Foley

Supported by
Personnel and Training Research Programs,
Office of Naval Research,
under Project Order Number
2-0046, NR 150-336

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FOREWORD

This research effort is being performed with the support of the Personnel and Training Research Programs Office, Office of Naval Research, under Project Order No. 2-0046, Work Unit Number NR 150-336 entitled PERSONNEL TECHNOLOGY: Relating Individual Performance Effectiveness to Unit and Ship Effectiveness. This is a final report on the analysis of the performance data collected in the course of the project.

Appreciation is expressed for the cooperation and assistance provided by Commander, Cruiser-Destroyer Force, Atlantic Fleet, Commander, Cruiser-Destroyer Force, Pacific Fleet, and Commander, Cruiser-Destroyer Flotilla NINE for providing the ships and men that took part in the data collection effort of this project.

The authors are indebted to Mr. William A. Sands of the Naval Personnel Research and Development Laboratory for his critical reading of the manuscript and invaluable suggestions on presenting the results of the data analysis.

The assistance of Dr. Arthur I. Siegel of Applied Psychological Services, Inc., Wayne, Pennsylvania, is also appreciated for providing some of the computer programs which were employed in analyzing the performance related data.

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NPRDL:71:PPF:sjh 3930 Ser 341

19 June 1973

From: Commanding Officer
To: Distribution List

Subj: Personnel Technology: Relating Individual Performance Effectiveness to Unit and Ship Effectiveness

Encl: (1) "An Analysis of a Fleet Post-Training Performance Measurement Technique," WTR 73-47, June 1973

- 1. Enclosure (1) is an evaluation of four different performance estimators relative to a criterion measure of absolute technician performance. The technicians were involved in fleet electronic maintenance activities in one of the ratings EM, ET, FT, IC, RD, RM, ST, and TM.
- 2. The report is forwarded for information. Comments and recommendations are invited.

A. L. BLANKS

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SUMMARY AND CONCLUSIONS

Problem

The advent of a more streamlined Navy operating under reduced manning levels and heightened operational requirements imposes the need for accurate human-performance evaluation of ship personnel systems. On the personnel systems level, electronic technician reliability measurement is a necessary and integral part in the evaluation of particular combat systems of which technicians are components. The objective is then to develop and evaluate human-performance reliability estimates so as to be able to effectively alter the personnel system in order to maximize the overall performance of the system. The purpose of this project is to validate the utility and effectiveness of a unique human performance measurement procedure developed under a prior Office of Naval Research project and designed to improve upon existing performance measuring techniques in a systems environment.

Background and Requirements

Human reliability performance estimation can be accomplished by considering the individuals being evaluated as components of a personnel system. This consideration allows the use of much of the theory already applied to equipment reliability estimation to be modified to human-performance estimation. After this theory is applied to evaluate the performance of human components in a personnel system, appropriate combinations of the individual performance estimators will provide a performance or reliability estimate of the personnel system itself.

In order to improve upon existing performance estimators of the human component in a personnel system, Dr. Arthur I. Siegel and his associates of Applied Psychological Services, Inc., Wayne, Pennsylvania, developed fleet post-training performance criteria for electronic maintenance personnel with the support of the Office of Naval Research. The cumulation of these efforts resulted in the development of unique human performance measurement techniques, closely allied with equipment reliability estimation techniques. Siegel also developed procedures for combining the technician performance estimates in appropriate ways in order to estimate team, ship or squadron performance.

An outgrowth of the prior research effort was the suggestion that the techniques be introduced on a limited basis to determine how they may be modified or elaborated upon. The Naval Personnel Research and Development Laboratory has been validating the utility and effectiveness of these techniques. The main technical objectives of this validation effort are to determine the validity of the performance measurement techniques, identify modifications required to maximize their possibility for implementation and to comment on the statistical properties of those techniques as related to their effectiveness in an operational context.

Approach

In order to realize an efficient and timely data collection effort, optical scanning instruments were utilized similar to those employed by Applied Psychological Services in prior research efforts.

The main data collection instruments were:

1. Job Performance Questionnaire (JPQ) ANSWER SHEET - this form records supervisory estimates of the total number of a technician's uncommonly effective (ΣUE) and uncommonly ineffective performance (ΣUI) that the supervisor has observed during a specified time period.



- 2. <u>Technical Proficiency Checkout Form</u> (TPCF) this form records the level of technical complexity at which a man is able to perform without direct supervision.
- 3. Personnel Identification Information Form (PIIF) this form records demographic data on the technician being evaluated.

On each of the above instruments an individual in one of the electronic maintenance ratings EM, ET, FT, IC, RD, RM, ST, and TM was evaluated by his supervisor. On the basis of the total number of uncommonly effective (Σ UE) and the total number of uncommonly ineffective (Σ UI) incidents of performance recorded on the Job Performance Questionnaire (JPQ), three different performance estimators were developed previously by Applied Psychological Services. These estimators are functions of the total number of uncommonly effective (Σ UI) incidents of performance observed by the supervisor on each of eight job dimensions characteristic of electronic maintenance activities. The three estimators of human reliability are:

- 1. Series Reliability Estimate (SRE)
- 2. Series-Parallel Reliability Estimate (PRE)
- 3. Geometric Mean Reliability Estimate (GRE)

In addition, a fourth measure of technician on-the-job performance developed in the course of this research effort was:

4. Weighted-Average Reliability Estimate (WRE).

By adopting the Technical Proficiency Checkout Form (TPCF) as a criterion measure of technician on-the-job performance, the degree of association between each of the four performance estimators and criterion measure was developed for various locations and each rating. Furthermore, a curvilinear regression analysis was applied to determine the best linear relationship between those variables.

While the purpose of this project was not to evaluate the criterion measure (TPCF), conditional and joint frequencies of the job tasks by rating revealed the modifications needed on the TPCF to make it more current. Furthermore, from the conditional and joint frequencies it was possible to develop a competency level for each rating and permit an in-depth analysis of the job task structure for those ratings.

Comparisons between ships and ratings were made by employing the WRE since it was identified as a more promising performance estimator. Initially a two-way Analysis of Variance was employed on the sample data. However, because of the significant interaction that was found to exist between ships and ratings, comparisons between ships (ratings) were made for a fixed rating (ship).

Finally, product-moment correlation coefficients were developed between various performance variables and several demographic variables. In particular, Basic Test Battery scores (GCT, ARI, MECH, and CLER scores), usually employed to predict actual school success, were investigated in order to determine the degree of association between those scores and the measures of on-the-job performance developed in this research effort.

Findings

In order to determine the appropriateness of standard statistical techniques or tests employed for various purposes in this research effort, initial findings were concerned with the results of an analysis of the distributional properties of the predictor and criterion variables. From an application of appropriate goodness-of-fit tests for



normality to the sample data, only the predictor variable WRE could be termed normally distributed. This result was also generally true when individual ratings were similarly studied. These characteristics of the sample data necessitated the use of a curvilinear regression analysis. An emphasis on only the least-squares analysis resulting from an application of that technique was employed in order to determine the relative merits of each of the predictor variables.

A comparison of the multiple correlation coefficients resulting from the curvilinear regression analysis applied across all ratings revealed that a straight-line was the best fit to the sample data. The product-moment correlations between the predictor variables (SRE, PRE, GRE, and WRE) and the selected criterion variable suggested moderate promise on the part of the WRE for appraising the on-the-job performance of individuals involved in electronics maintenance activities. However, the relatively low multiple correlation coefficients suggested a significant degree of unexplained criterion variance. As such the analysis was applied by rating. Relevant findings by rating revealed that in almost every rating the WRE demonstrated more promise for appraising on-the-job technician performance than the other estimators. The WRE was considered a more promising estimator in the sense that the sample product-moment correlation coefficients were generally of a larger magnitude and, by some fit to the data, the WRE seemed to account for more criterion variance. For example, the WRE indicated product-moment correlations of .492, .445, .430, and .434 for the EM, ET, FT, and IC ratings, respectively, while the next more promising estimator (GRE) demonstrated product-moment correlations of .301, .508, .374, and .387 for the EM, ET, FT, and IC ratings, respectively. It is to be noted that while the WRE may not necessarily be a statistically significantly different estimator in terms of producing higher correlations with the criterion variable and explaining more criterion var ance, the sample data results did tend to demonstrate that the WRE was the more promising estimator in that consistently the sample results did produce higher product-moment and multiple correlation coefficients for the WRE. Altogether the WRE demonstrated very promising results on the EM, ET, FT, and IC ratings and fair results on the RM, S', and TM ratings. None of the performance estimators were at all promising in the RD rating.

Relevant findings from an analysis of the Technical Proficiency Checkout Form revealed that it was a very instructive instrument for determining technician proficiency in one job task in relation to another. Only one job task - the calibrating of the equipment used by the technicians - seemed to be out of place in the hierarchial order of the job tasks represented by this instrument.

Findings of the multiple comparison of ships and ratings included the result that significant interaction exists between ship-rating combinations. This resulted in the development of multiple comparisons of ships (ratings) for a fixed rating (ship). From this analysis it was found that no pairwise significant difference exists between ships. On some ships a pairwise significant difference exists between some ratings, but no pattern emerged across ships as to which rating(s) demonstrated a higher or lower mean performance level.

The product-moment correlation coefficients developed between the demographic variables and the performance variables were of the same magnitude as those which are usually found to exist between predictor variables and measures of on-the-job performance. Promising results were found in this research effort on the relationship of demographic variables to on-the-job performance.

Conclusions

Employing the TPCF as a criterion measure of technician on-the-job performance it may be stated that the following list represents an accurate portrayal of the performance measurement technique that was researched:

1. The distribution of the predictor variables SRE, PRE, and GRE are generally skewed in one direction or another, while the criterion variable derived from the TPCF is negatively skewed. The WRE is normally distributed. These conclusions are for each rating and across ratings.



- 2. The WRE is a more promising type of performance estimator. It has greatest utility when applied to technicians in the EM, ET, FT, and IC rating and fair promise for application in the RM, ST, and TM ratings.
- 3. None of the performance estimators is appropriate for use upon technicians in the RD rating.
- 4. The Technical Proficiency Checkout Form (TPCF) demonstrated significant promise for appraising the job task structure and proficiency of electronics maintenance personnel without restriction to a particular rating.
- 5. In all ratings a more current factor analytic task analysis would be desirable before implementation of the technique. This would involve a revision of the job activity factors on the JPQ ANSWER SHEET and job task descriptions on the TPCF by rating.
- 6. It is recommended that a validation of the performance variables (SRE, PRE, GRE and WRE) be completed before the technique is applied to other than those ratings researched in this report.

In conclusion it is to be noted that the performance measurement technique that was researched is of significant merit to be considered for practical application in the U.S. Navy (particularly in the EM, ET, FT, and IC ratings). At the present stage in the art of developing performance measurement techniques, the technique that was researched is probably the best performance measurement procedure presently available for application within the U.S. Navy.



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INTRODUCTION

Current and expanded commitmed is of a modern sophisticated Navy will require greater operational effectiveness of fleet personnel systems. The Navy will have to operate fleet personnel systems at optimal effectiveness levels and maintain fleet readiness. Together with these requirements, the advent of an all-volunteer force and smaller ship systems with reduced manning levels make the problem of optimizing, and evaluating, personnel system effectiveness a complex and critical problem.

While performance assessment serves a multitude of purposes, it has been seen as an especially valuable tool when applied to the areas of optimizing personnel system performance, providing feedback on naval school effectiveness, the interpretation of man-machine interaction, and as a factor in the optimal assignment of men to jobs. It is in these applications that personnel systems performance measurement will be able to address some of the more critical present-day Navy problems. For example, many ship systems are operating with increasingly sophisticated and complex equipment. However, is it necessarily true that technicians of comparable sophistication in training and mental ability need to be employed to operate, maintain, and repair that equipment in order that the ship complete its mission? While this report does not address that specific question, future research employing individual, and system, performance measurement may reveal that the Navy could effectively utilize personnel in those positions who may now be rejected for some reason related to their projected on-the-job performance in those positions. As such there is a definite need for valid and reliable individual system performance assessment.

In order to achieve a performance appraisal of personnel systems, recent research has been directed towards viewing personnel system performance estimation as analogous to equipment reliability estimation. This particular approach views individuals (in the system) as "components" in the personnel system. This viewpoint, and an application of the techniques already employed in equipment performance estimation, reduces personnel system performance estimation to an evaluation of the performance of the individuals in the system. Once the individual performance estimates have been made, meaningful combinations of these estimates can provide estimates of personnel system performance. As such the initial problem reduces to that of finding accurate and valid measurements of individual performance.

Much recent research in the area of individual performance measurement has been directed towards examining individual performance as a function of the extremes of behavior (critical incidents) of an individual's performance. These functions of behavior can provide estimates of individual performance. By appropriate combinations of the individual performance estimates, estimates of personnel system performance can be developed. This report will address the validity, application, and implications of a particular procedure that employs the critical incidents technique to estimate the performance of electronics maintenance personnel.

BACKGROUND

The purpose of this section is to give the reader a logical development of the performance measurement techniques employed by Applied Psychological Services and others. Fundamentally these researchers employed a critical incidents technique in deriving estimates of human performance.



Generally the main approach is to estimate the performance of a particular personnel system as a function of the performance of individuals that are a part of the system. This necessarily reduces personnel system performance estimation to a discussion of estimators of individual performance where individuals are the components of the system. Combining the individual estimates will provide estimates of personnel system performance.

Personnel Performance Estimation

Let UE (UI) represent an uncommonly effective (uncommonly ineffective) incident of performance observed by a rater in a certain time period on some individual under observation. Furthermore, let Σ UE (Σ UI) represent the total number of uncommonly effective (uncommonly ineffective) incidents of performance observed. Using these functions of critical incidents, Whitlock [21] demonstrated that there is a definite straight line or curvilinear relationship between Σ UE (or the ratio Σ UE/ Σ UI) and corresponding performance evaluations. Prior results such as this provided significant evidence that the application of a critical incidents technique to performance evaluation is a valid and useful approach.

Following upon the results of Whitlock, for example, Applied Psychological Services further developed and applied the above mentioned techniques to the post-training performance evaluation of individuals in various avionic or electronic ratings in the U.S. Navy. In particular Siegel and Pfeiffer [18], utilizing estimates of uncommonly effective and uncommonly ineffective performances, showed that these estimates possess merit as useful indicators of overall personnel proficiency. The researchers employed magnitude estimates of the number of uncommonly effective and uncommonly ineffective performances relative to a short prior period for avionic personnel. They derived a performance index from the ratio of the sum of uncommonly effective performance (SUE) to the sum of uncommonly effective plus the sum of uncommonly ineffective performance (Σ UI), namely (Σ UE/[Σ UE + Σ UI]). Siegel and Pfeiffer [18] concluded that: (1) magnitude estimates of uncommonly effective and ineffective performance yielded useful data which could form the basis for a personnel subsystem reliability index; (2) the ratio of the sum of uncommonly effective to the sum of uncommonly effective plus the sum of uncommonly ineffective performance yields an index which discriminates in the anticipated direction; and, (3) the obtained avionic personnel subsystem index could be utilized for post-training performance appraisal, personnel placement, and squadron evaluation purposes.

Job Performance Questionnaire

Eight job activity factors descriptive of naval avionic electronic maintenance jobs were isolated by Siegel and Schultz [19] and are shown in Appendix A, page A-7, along with their definitions. These factors formed the basis of the Job Performance Questionnaire (JPQ), an instrument for recording the frequency of critical incidents for each of the job activity factors. Siegel and Federman [16] demonstrated the utility and practicality of a Job Performance Questionnaire (Appendix A, page A-3) for technicians in the eight electronic maintenance ratings EM (electrician's mate), ET (electronics technician), FT (fire control technician), IC (interior communications electrician), RD (radarman), RM (radioman), ST (sonar technician), and TM (torpedoman's mate). From an evaluation of 499 technician in those ratings, the researchers found that the JPQ yields an estimate of the total number of uncommonly effective and uncommonly ineffective incidents of behavior on the eight identified job activity factors.



¹All numbers enclosed in brackets refer to corresponding numbers of documents and publications listed under REFERENCES.

Specifically, for each job activity the reliability ratio ($\Sigma UE/[\Sigma UE + \Sigma UI]$) yields an estimate of the probability of effective performance for the individual technician on the particular job activity considered. Then the reliability ratios are compounded to provide estimates of individual effectiveness or reliability of on-the-job performance across the job activities. It was reported by Siegel and Federman [16] that estimates of uncommonly effective and of uncommonly ineffective behavior along all eight dimensions of job activities could be combined into a meaningful measure of technician effectiveness. Moreover, they indicated that the individual technician effectiveness values can be further treated to form effectiveness values for ratings, ships, and squadrons.

Estimation of Technician Reliability

Employing the reliability ratio concept ($\Sigma UE/[\Sigma UE + \Sigma UI]$), Siegel and Federman [16] have developed the following three reliability estimates:

1) <u>Series Reliability Estimate</u> (SRE)
The series reliability measure of total effectiveness for an individual is derived by multiplying individual job activity reliability ratios to yield a total reliability score, i.e.,

$$R_{s} = r_{1} \times r_{2} \times ... \times r_{8}$$

where R_s = series reliability², and

 r_i = (SUE/[SUE + SUI]) is the reliability ratio for the i^{th} job activity.

It is to be noted that use of the series reliability estimate requires the assumption that performance reliability on each job activity is independent of performance reliability on other job activities.

2) Series Parallel Reliability Estimate (PRE)
Siegel and Federman [16] reported that "... the series and series-parallel reliabilities provide measures of personnel proficiency relative to performance on the entire job (i.e., all eight job activities).", (p. 46). The series-parallel estimate of individual proficiency is defined as:

$$R_p = R_s \times (2 - r_1) \times ... \times (2 - r_8)$$

where R_s and $r_1(i = 1, ..., 8)$ are defined in 1) above.

- a) for each of the i = 1, ..., 8 job activities
 - $0 \le r_1 \le 1$, and, therefore,
- b) $0 \le R_s \le 1$
- c) $R_s \leq smallest r_i$.



²It is helpful to note that the series reliability estimate possesses the following properties:

This particular estimate tends to provide a more optimistic estimate of individual performance. However, the content validity and derivation of this estimate deserves further development.

3) Geometric Mean Reliability Estimate (GRE) Let r_1^* , r_2^* , r_3^* , and r_4^* be the four highest job activity reliability

ratios of the eight reliability ratios for a technician being evaluated. The geometric mean reliability for the technician is defined as:

$$R_g = 4 / r_1^* \times r_2^* \times r_3^* \times r_4^*$$

This particular estimate is an estimate of individual performance that stresses the strong points of an individual's performance. However, it also tends to ignore his weak points and, therefore, should be used with caution.

In addition to the three performance estimators (SRE, PRE, and GRE) previously introduced, this report will also discuss an estimate that weights the importance of each job activity in determining a technician's overall performance.

- 4) Weighted Average Reliability Estimate (WRE) To develop this estimate let:
 - NJ = number of job activities on which the technician actually worked;
 - i = index for the sum over the job activities on which the technician actually worked;
 - r_i = the reliability ratio for the ith job activity;
 - w_i = weight denoting the importance of the i^{th} job activity in estimating the technician's overall performance.³

The Weighted-Average Reliability Estimate (WRE) of technician effectiveness is then defined as:

$$R_{W} = \sum_{i=1}^{NJ} r_{i} \times w_{i}/NJ$$

Validation of Performance Estimators

In addition to performance data collected on the JPQ, performance data were collected by Siegel and Federman [16] by means of an evaluative instrument called the Technical Proficiency Checkout Form (TPCF) (Appendix A, page A-5). The TPCF consists of eight job tasks listed in hierarchial order from easiest to most difficult. The eight tasks meet the Guttman requirements for scalability (see for example, Guttman [11]). Siegel, Schultz, and Lanterman [20], 1964, employed the scale underlying the eight tasks to determine the cutting points for placing avionic petty officers, third class and strikers in one of three levels of technical proficiency. The procedure for placing a technician in one of three levels of technical proficiency is accomplished by means of a Technical Proficiency (TP) score developed from the TPCF.



 $^{^{3}\}text{The procedure for deriving the weights is given in Appendix K.}$

Technical Proficiency (TP) Score

Define the function F_i (i = 1, ..., 8) as:

$$F_{i} = \begin{cases} 1 \text{ if the technician is CHECKED OUT on the } i^{th} \\ \text{task of the TPCF} \\ 0 \text{ if the technician is NOT CHECKED OUT on the } i^{th} \\ \text{task of the TPCF.} \end{cases}$$

The TECHNICAL PROFICIENCY (TP) score for a technician is then defined as:

TP score =
$$\sum_{i=1}^{8} F_i$$
.

Technical Proficiency Checkout (TPC) Level

Three TECHNICAL PROFICIENCY CHECKOUT (TPC) levels are:

Level 1: above desirable

Level 2: below desirable but at least minimally acceptable

Level 3: below minimally acceptable.

Siegel, Schultz, and Laterman [20] based this trichotomous division of the TPCF on supervisor's judgments of the performance level required for achieving the objectives given in Appendix A, page A-8.

The procedure for determining the TPC level was reported on in the previously mentioned report and is given by:

- a) add 0.5 to TP score for an individual. Let TP* be the resultant score.
- b) if TP* < 3.92, the TPC level \approx 3
- c) if $3.92 \le TP* \le 5.63$, the TPC level = 2
 - d) if TP* > 5.63, then TPC level = 1.

Siegel and Fischl [17] correlated technicians TPC levels with the technicians total scores on a performance test. Employing a triserial correlation coefficient (see, for example, Jaspen [13]) as an estimate of the product-moment correlation, they found a triserial correlation of .40. When corrected for the lack of perfect reliability in the performance test criterion, the correlation became .74. On the basis of their investigation of the then concurrent validity of the TPCF, they concluded that the Technical Proficiency Checkout Form, "... previously shown to be reliable and practical, may now be considered to possess a substantial degree of validity for appraising the absolute proficiency level of avionics technicians in the fleet.", (p. 46). Finally, Siegel and Federman [16] recorded a triserial correlation of .38 between the TPC level of the technicians evaluated and their Series Reliability Estimate (SRE), concluding "... there is some basis to believe that the JPQ results correlate with on-the-job performance.", (p. 62).



Main Results of Prior Studies

Important conclusions of prior reports relative to the merits of the <u>Series Reliability Estimate</u> (SRE), <u>Series-Parallel Reliability Estimate</u> (PRE), <u>Geometric Mean Reliability Estimate</u> (GRE), and the <u>Technical Proficiency Checkout Form (TPCF)</u> are as follows:

- 1. Reliability ratios of the form $\Sigma UE/(\Sigma UE + \Sigma UI)$ indicate the probability of effective performance on a particular job activity for the technician being evaluated.
- 2. The JPQ is an instrument for providing magnitude estimates of ${^\Sigma UE}$ and ${^\Sigma UI}$ for each man being evaluated by his immediate supervisor.
- 3. The TPCF possesses a substantial degree of promise for appraising the absolute level of avionic technician proficiency.
- 4. There is some basis (triserial correlation of .38 with TPC level) to believe that the SRE is a reasonably good estimator of on-the-job performance.

PURPOSE

The purpose of this research effort is to report on the data collection effort and data reduction methods and analyses that have been performed for the Office of Naval Research (ONR) under the project entitled Personnel Technology: Relating Individual Performance Effectiveness to Unit and Ship Effectiveness (Project Order Number: PO 2-0046 NR 150-336). The goal of this research project is to provide an empirical basis for assessing the utility to the Navy of a performance measurement technique developed under a prior ONR contract. Under that contract Dr. Arthur I. Siegel, Philip J. Federman, and their associates of Applied Psychological Services, Inc., Wayne, Pa., developed fleet post-training performance evaluative measures which have potential value for eventual widespread implementation within the U. S. Navy. The results of their study were contained within the report - Development of Performance Evaluative Measures: Investigation into and Application of a Fleet Post-Training Performance Evaluative System [16]. An outgrowth of that effort was the suggestion that the technique be employed on a limited basis by a Navy laboratory to identify areas of modification upon operational testing. In response to that recommendation the Naval Personnel Research and Development Laboratory submitted a proposal to ONR to accomplish that task. Essentially the research effort undertaken by NAVPERSRANDLAB was accomplished by replicating the efforts of Siegel and Federman [16].

A second major objective of this research effort was to further develop the performance measurement techniques of Siegel and Federman [16]. Furthermore, similarly related performance measurement techniques were researched with a view towards possible implementation of those techniques within the U.S. Navy.

DATA COLLECTION

The procedures employed in data collection for this project closely paralleled those employed by Siegel and Federman [16], with some modifications in the research instruments. This procedure was adopted so that a similar statistical analysis on the same type of population would permit some comparisons to be made between the results of this research effort and the results obtained by Siegel and Federman [16].



Every effort had been made to minimize interfering with normal shipboard duties. For this reason the data collection procedure centered upon the efforts of ship liaison officers conducting the data collection aboard each ship. Appendix B contains a discussion on the procedures for the data collection reflecting various aspects of the effort that resulted in the orderly and successful completion of the task. 4

Data Collection Instruments

An example of the performance evaluation forms that were completed by each supervisor for each technician evaluated are given in Appendix A. The instruments are optical scanning forms, thus, making them machine-readable and more capable of being placed in an operational mode. In particular the forms were:

- 1) Job Performance Questionnaire (JPQ) ANSWER SHEET This form serves the same purpose as the JPQ discussed earlier, i.e., to record estimates of the total number of uncommonly effective (Σ UE) and uncommonly ineffective (Σ UI) performances the supervisor has observed on each of the eight job activities for each man he is evaluating.
- 2) <u>Technical Proficiency Checkout Form (TPCF)</u>
 This form is essentially identical to the TPCF used by Siegel and Federman [16].
- 3) Personnel Identification Information Form (PIIF)
 This form was concerned with the background data of the man being evaluated. It was completed in part by his supervisor with the Administrative Officer providing the remaining information.

Analyses Based on the Data Collection Effort

Employing the data collection instruments discussed in the previous section, performance data were collected with the assistance of men and ships of Commander; Cruiser-Destroyer Flotilla NINE (located at San Diego, California) and men and ships of Commander; Cruiser-Destroyer Force Atlantic Fleet (located at Newport, Rhode Island, and Boston, Massachusetts). The participating ships and type are shown in Table 1 along with the number of men evaluated by rating and ship for each location.

Analyses Based on the Job Performance Questionnaire

A descriptive analysis of each of the performance estimators (SRE, PRE, GRE, and WRE) derived from the Job Performance Questionnaire (JPQ) is presented in Figures 1, 2, 3, and 4 and in the form of histograms. Each of these histograms was developed on 949 technicians and based on the performance estimators which are continuous over the range of 0.0 to 1.0. Class intervals are numbered from 1 through 21 where a given class interval is of length .05. Class intervals corresponding to each of the numbered intervals are provided in Table 2.



For the interested reader Appendices C and D contain the instructions for the ship liaison officer and the technician supervisor, respectively.

⁵Henceforth in this report Location No. 1 will refer to ships at San Diego, Calif., and Location No. 2 will refer to ships at either Newport, R.I., or Boston, Mass.

TABLE 1

NUMBER OF MEN IN EACH RATING AND SHIP

CRUDESFLO: NINE (Location No. 1)

Ship	Туре	EM	ET	FT	IC	RD	RM	sτ	TM	TOTAL
USS AGERHOLM	DD-826	6	7	7	4	5	5	5	3	42
USS BROOKE	DEG-1	6	8	16	3	10	8	9	1	61
USS GRAY	DE - 1054	5	10	5	4	9	7	9	1	50
USS HORNE	DLG-30	4	12	11	6	5	5	8	2	53
USS HULL	DD-945	3	9	8	3	9	9	8	1	50
USS JOUETT	DLG-29	5	10	12	5	7	6	9	2	57
USS PRAIRIE	AD-15	6	29	12	6	0	10	C	13	76
USS RUPERTUS	DD-851	5	11	3	2	9	9	10	1	50
USS SHELTON	DD-790	4	5	5	1	5	5	5	0	30
USS SOUTHERLAND	DD-743	5	8	9	5	11	11	9	2	60
USS HENRY W. TUCKER	DD-875	5	9	0	3	13	וו וו	11	2	54
TOTAL		54	118	88	42	83	86	83	28	582

CRUDESLANT (Location No. 2)

Ship	Type	EM	ET	FŢ	IC	RD	RM	ST	TM	TOTAL
USS BASILONE	DD-824	4	4	5	0	0	5	5	0	23
USS DEALEY	DE-1006	3	6	5	2	3	4	6	1	30
USS DEWEY	DLG-14	6	6	6	4	6	5	6	2	47
USS FISKE (Boston)	DD-842	5	4	5	0	5	5	6	0	30
USS J. A. FURER	DEG-6	4	5	14	4	5	8	5	2	47
USS GARCIA (Boston)	DE -1040	4	6	6	2	6	8	10	2	44
USS GLOVER (Boston)	AGDE ~1	4	7	3	0	6	6	12	1	39
USS HUCH PURVIS	DD-709	6	7	10	0	15	0	9	2	49
USS TALBOT	DEG-4	5	5	9	2	5	5	5	0	36
USS JOHN WILLIS	DE-1027	2	5	3	2	5	5	5	1	28
TOTAL		43	55	66	16	56	51	69	11	367
TOTAL (BOTH LOCATIONS)		97	173	154	58	139	137	152	39	949



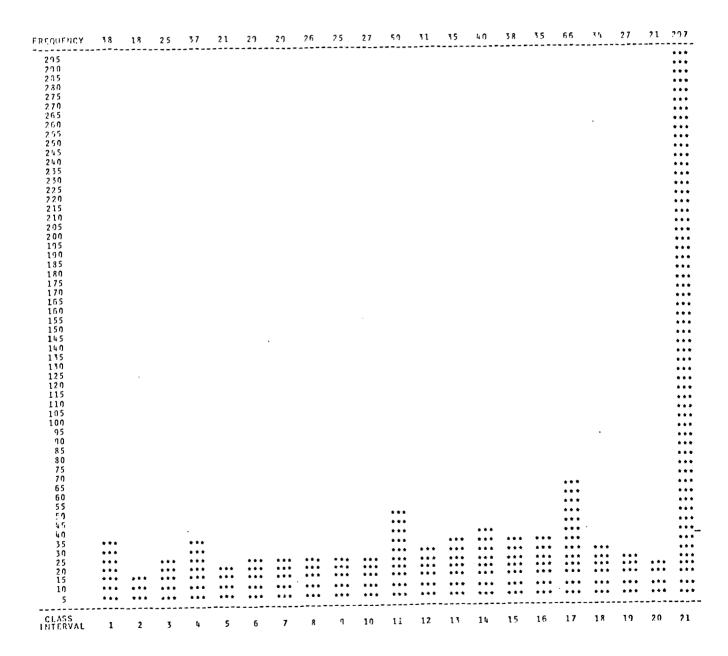


Fig. 1.--Histogram of Series Reliability Estimates (SRE)



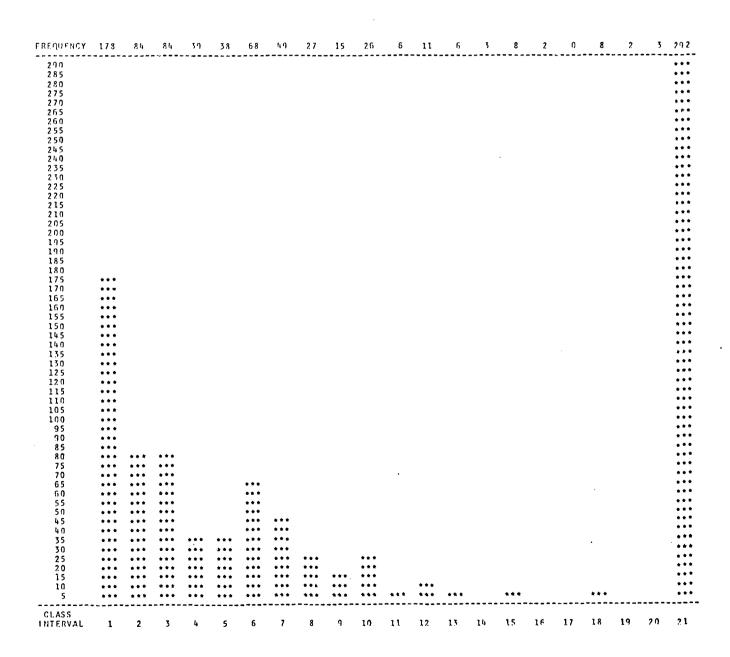


Fig. 2.--Histogram of Series-Parallel Reliability Estimates (PRE)



REQUENCY	519 	154	122	70	40	12	8	2	2	2	l	 0	0) 1) () (16
513	***											 		 	 			
504 4 95	***																	
436	***																	
477	***																	
468	***																	
459	***																	
450	***																	
441	* * *																	
4 5 2 4 2 3	***																	
1,14	***																	
405	***																	
396	***																	
337	***																	
378	***																	
369 360	***																	
35 l	***																	
342	***																	
333	***																	
324	***																	
315	* * *																	
306	***																	
297 288	***																	
279	***								•									
270	***																	
26 l																		
252	***																	
243	***																	
234	***																	
225	* * *																	
216 207	***																	
198	***																	
189	***																	
180	***																	
171	* * *																	
162	***																	
153 144	***	* * *																
135	***	***																
126	***	***																
117	***	***	***															
108	***		***															
91	***	***	* * *															
20	***	***	***															
8 l 7 2	***	***	***															
72 63	***	***	***															
54	***	***		***														
45	* * *	***	***	***														
3 G	***	***	***															
2.7	***	• -	***	***	***													
18	***	***	***	***	***													
g 		***		***	***							 		 	 			***
LASS														 	 		-	

Fig. 3.--Histogram of Geometric Mean Reliability Estimates (GRE)



FREQUENCY	24	25	46	62	78	79	75	94	92	96	70	54	ti ()	2.8	16	27	14	11	4		:
94							***														
92								* * *													
3.0							***	* * *	•••												
8.8							* * *	* * *	***												
86							* * *	* * *	***												
84							***	* * *													
82							***	* * *	***	* * *											
80							* * *	* * *	***	* * *											
78					* * *	* * *	* * *	* * *	***	***											
7.6					***	* * *	***	* * *	* * *	4 * *											
74					***	* * *	* * *	***	* * *	***											
72					***	* * *	***	* * *	* * *	* * *											
70					4 * *	* * *	* * *	* * *	* * *	***	* * *										
6.8					* * 3	* * *	***	* * *	* * *	• • •	***										
66					***	• • •	* * *	* * *	• • •	***	***										
64					***	* * *	***	• • •	* * *	* * *	* * *										
62				***	***	***	•••	***	* * *	* * *	* * *										
6 0				***	***	***	* * *	***	* * *	* * *	***										
58				***	* * *	• • •	* * *	* * *	* * *	* * *	**										
56 54				h * +	* • •	* * *	* * *	* * *	• • •	* * *	***										
5 2				***	* * *	***	* * *	***	* * *	* * *	***	* * *									
50				***	* * *	***	***	***	* * *	* * *	* * *	***									
48				***			***	* * *	* * *	* * *	* * *	* * *									
46				* 6 *			***	***	* * * *	***	***	***									
44				***		***	***				***	***									
4.2				2 * *			* • •		• • •		• • •	***									
40			***				***					***	***								
38			***	***		***	***		• • •	* * *	***	• • •									
36						• • •			• • •		***										
34			***	***	***		***	* * *			***										
32			* * *			* * *					***	***									
30			***	***	***	***		***			***										
28				* * *	***	***	* * *		***		***		* * *	* * *							
26			***				***	+ + +		* * *		* * *	***	***		***					
24		***	***	***		***		***		***	***					***					
22	***		***	* * *		_* *	***	***	* * *			* * *		***.		***					
20	***	* * *	***	***		***	* * *			***	***	* * *	* * *			* * *					
18	***	***	* * *	***		* * *	***	***	* * *	***	***	* * *		* * *		* * *					
16	* * *	* # *	***	***	* * *	***	* * *	* * *				***	* * *	***	***	* * *					
14		***	***	***	***	***	***		***	***	* * *	***	***	**	**	***	***				
12	***	***	***	***	***	***	***	***	***	* * *	* * *	***	* * *	* * *	***	***	***				
10	***	***	***	• • •	* * *	***	* * *	***	***	* * *	* * *	* * *	* * *		* * *	***	***	***			
8	***	**	***	***	* * *	* * *	***	* * *	***	* * *	***	***	***	***	***	***	***				
6	* * *	***	***	***	* * *	***	***	• • •	***	***	* * *	* * *	***	* * *	* * *		***	***			
4	***	***	•••	***	***	• • •	***	***	***	***	• • •	* * *	* * *	* * *	* * *	* * *	***	***	***		
2	***	* * *	***	***	* * *	***		* * *	• • •	* * *	* * *	* * *	* * *	* * *	* * *	***	***	***	***	***	**
CLASS																					

Fig. 4.--Histogram of Weighted-Average Reliability Estimates (WRE)



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TABLE 2
CLASS INTERVALS FOR HISTOGRAMS

Class Interval Number	Lower Boundary	Upper Boundary
1	0.96	1.00
2	0.91	0.96
3	0.86	0.91
4	0.81	0.86
5	0.76	0.81
6	0.71	0.76
7	0.66	0.71
8	0.61	0.66
9	0.56	0.61
10	0.51	0.56
11	0.46	0.51
12	0.41	0.46
13	0.36	0.41
14	0.31	0.36
15	0.26	0.31
16	0.21	0.26
17	0.16	0.21
18	0.11	0.16
19	0.06	0.11
20	0.01	0.06
21	0.00	0.01



Further analysis of results on the JPQ ANSWER SHEET indicated a high frequency of nonresponse in some job activities and ratings. The type of nonresponse that resulted was for the case in which the man being evaluated did not work at the particular job activity under consideration. Furthermore, there was also a significantly high proportion of men who, while they worked at the job activity being considered, received $\Sigma UE = 0$ and $\Sigma UI = 0$ from their supervisors. These observations required the consideration of two important areas relative to the JPQ.

<u>Problems in Calculating Performance Estimates</u>. As discussed in the Background section of this report, reliability ratios of the form ($\Sigma UE/[\Sigma UE + \Sigma UI]$) were derived for each man on each of eight job activities and these ratios were combined to form the SRE, PRE, and GRE. However, the following two cases require the adoption of some convention in order to calculate the reliability ratios:

- 1) the technician did not work at that job activity, or
- 2) the technician received $\Sigma UE = 0$ and $\Sigma UI = 0$ by the supervisor, implying that the reliability ratio 0 = 0 is undefined.

By observing the frequency with which such cases occur across all 21 ships participating in the project, it is possible to determine the extent to which any convention for estimating performance in those cases would effect individual SRE, PRE, and GRE values. A complete discussion of this effect is given in Appendix E. Summarizing, the above two cases can have a dramatic affect upon the individual performance estimates and these estimates will be greatly influenced by the convention that is adopted.

A Convention for Estimating Performance in Certain Job Activities. Siegel and Federman [16] employed "... the average value for his rating on his ship ...", (p. 28), on those job activities which the technician did not work at or received $\Sigma UE=0$ and $\Sigma UI=0$ by his supervisor. Unfortunately the results of the data collection effort at Location No. 1 (Destroyer Flotilla NINE) and at Location No. 2 (Cruiser-Destroyer Force Atlantic Fleet) indicated that this technique was not feasible. 6

In order to overcome this problem, the convention adopted in this report was to employ a composite reliability value across all shirs at a location for each job activity and rating. Appendix F discusses the procedure for deriving the composite reliability values, as employed in this report.

Analyses Based on the Technical Proficiency Checkout Form

Table 3 represents the numbers of men at each of the three TPC levels by rating and ship and across each rating and ship at Location No. 1. Table 4 reflects the same information for Location No. 2. It will be remembered that level 1 reflects an "above desirable" proficiency level while level 3 reflects a "below minimally acceptable" proficiency level.

In addition to the TPC levels, TP scores were developed. A histogram of the resulting TP scores for the 949 technicians evaluated is presented in Figure 5. Almost identical histograms of TP scores were obtained for data collected at the two locations. Hence, only one histogram is presented.



 $^{^{6}}$ On every ship sampled at those locations there were ratings for which in some job activities those two cases occurred for all men in that rating. Appendix F provides a detailed account of this problem for the interested reader.

TABLE 3

NUMBER OF MEN AT EACH TECHNICAL PROFICIENCY CHECKOUT (TPC) LEVEL

Location No. 1

						:	SHIP						EACH
RATING	TPC LEVEL	1	2	3	4	5	6	7	8	9	10	11	RATING
EM	1	4	5	4	2	1	4	5	3	2	5	3	38
	2	1	1	0	1	1	1	1	2	2	0	2	12
	3	1	0	1	1	1	0	0	0	0	0	0	4
ET	1	4	6	7	9	7	8	17	8	5	8	7	86
	2	2	2	3	1	2	1	9	2	0	0	1	23
	3	1	0	0	2	0	1	3	1	0	0	1	9
FT	1	5	10	3	10	8	8	6	2	2	5	0	59
	2	2	5	0	0	0	3	3	0	3	0	0	16
	3	0	1	2	1	0	1	3	1	0	4	0	13
IC	1	2	2	2	5	1	3	4	2	1	2	2	26
	2	1	0	1	1	1	2	1	0	0	3	0	10
	3	1	1	1	0	1	0	1	0	0	0	1	6
RD	1	0	0	0	0	0	0	0	0	0	1	0	1
	2	1	1	1	0	1	0	0	8	0	3	0	15
	3	4	9	8	5	8	7	0	1	5	7	13	67
RM	1	1	3	7	0	0	4	4	1	5	0	0	25
	2	1	1,	0	0	8	2	3	5	0	2	8	33
	3	3	1	0	5	1	0	3	3	0	9	3	28
ST	1	4	7	7	6	3	6	()	4	5	5	5	52
	2	1	2	0	1	2	0	()	6	0	4	4	20
	3	0	0	2	1	3	3	()	0	0	0	2	11
TM	1	2	1	0	2	1	2	4	1	0	2	2	17
	2	1	0	0	0	0	0	3	0	0	0	0	4
	3	0	0	1	0	0	0	6	0	0	0	0	7
EACH	1	22	34	30	34	21	35	40	21	20	28	19	304
	2	10	15	5	4	17	9	20	23	5	12	15	133
	3	10	12	15	15	14	12	16	6	5	20	20	145



TABLE 4 ·

NUMBER OF MEN AT EACH TECHNICAL PROFICIENCY CHECKOUT (TPC) LEVEL

Location No. 2

THE SHIP									EACH			
RATING	TPC LEVEL	1	2	3	Ļ	5	6	7	8	9	10	RATING
ЕМ	1 2 3	2 2 0	2 0 1	3 1 2	4 1 0	3 1 0	0 2 2	2 1 1	5 1 0	4 1 0	1 1 0	26 11 6
ET	1 2 3	3 1 0	5 1 0	5 1 0	3 0 1	5 0 0	5 1 0	7 0 0	7 0 0	2 2 1	3 1 1	45 7 3
FT	1 2 3	2 1 2	4 0 1	4 2 0	1 0	12 2 0	4 2 0	1 2 0	7 3 0	8 1 0	2 1 0	48 15 3
IC	1 2 3	0 0	0 1 1	2 2 0	0 0 0	1 1 2	0 0 2	0 0 0	0 0	2 0 0	1 0 1	6 4 6
RD	1 2 3	0 0 0	0 0 3	0 6 0	0 4 1	0 0 5	1 1 4	2 4 0	0 6 9	0 0 5	0 5 0	3 26 27
RM	1 2 3	3 2 0	3 1 0	4 1 0	1 4 0	7 1 0	2 6 0	0 3 3	0 0 0	0 4 1	0 0 5	20 22 9
ST	1 2 3	5 0 0	2 2 2	4 0 2	3 2 1	4 0 1	5 1 4	2 0 10	9 0 0	1 0	2 1 2	40 7 22
ТМ	1 2 3	0 0 0	0 1 0	0 1 1	0 0 0	2 0 0	1 0 1	0 1 0	1 1 0	0 0 0	1 0 0	5 4 2
EACH SHIP	1 2 3	15 6 2	16 6 8	22 14 5	15 12 3	34 5 8	18 13 13	14 11 14	29 11 9	20 9 7	10 9 9	193 96 78



OUENCY		טפי 	 64	(0	91	136	133	701	257
255									
2 50									••••
245									••••
240		:							
235									••••
230									• • • •
225									
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145									****
140									••••
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135							• • • •		••••
130						••••			••••
125						••••	••••		••••
150						••••	••••		****
115						••••	••••		••••
110 105							****		****
						••••	••••	****	••••
100					••••	••••	••••	••••	
95						••••	••••	••••	••••
90					••••	••••	••••	••••	
85						••••	••••	••••	••••
80					••••	••••	••••	••••	••••
75						••••	****	••••	••••
70									
55			••••	••••	••••	••••	• 0 • •	****	••••
60			•••	a • • •	••••	••••	••••	••••	••••
55			••••	***					****
50		••••	••••				••••		••••
45			••••	***	• • • •	••••	• • • •	••••	••••
40		•••	••••	0000	••••		••••	••••	••••
75		4004			••••		••••	••••	••••
30			••••	••••	****		••••	••••	••••
25			••••						••••
50								••••	****
15	0000					• * • •		••••	••••
10		***	••••	••••	••••	****	••••	••••	••••
5	••••	••••	••••	****	••••	••••	****	••••	••••

Fig. 5.--Histogram of Technical Proficiency Scores (N = 949)



APPROACH

An approach to the statistical analysis of the data involved the selection of appropriate analyses within four general areas:

1. The validity of the performance estimators SRE, PRE, GRE, and WRE with respect to a selected criterion measure.

In order to determine the validity of the four performance estimators (SRE, PRE, GRE, and WRE) for predicting on-the-job performance of the electronics maintenance personnel involved in the research effort, the results of the TPCF were used as a criterion measure of absolute technician proficiency. The belief that the TPCF reflects the on-the-job performance of electronics maintenance personnel must rest to a large degree upon related results of prior research efforts. In particular, from references that were cited in the section Validation of Performance Estimators.

Initially, in order to determine the degree of association between the performance estimators (SRE, PRE, GRE, and WRE) and the selected criterion variable (TPC level), triserial correlation coefficients were developed by location. Due to the extreme skewness of the distribution of the underlying continuum (represented by TP score, see Figure 5), a test of normality of TP score was executed to determine the appropriateness of triserial correlation. This resulted in the choice of a curvilinear regression analysis as a better approach for validating the performance estimators by location and subsequently by rating with TP score as the continuous criterion measure.

2. An evaluation of technician job competency as determined or implied by the TPCF.

The appropriateness of the job tasks represented on the TPCF was approached by developing a frequency table of men CHECKED OUT and NCT CHECKED OUT by rating on each job task. The agreement of each job task to the hierarchial classification of the tasks provided some indication of the extent to which it was still applicable to electronic maintenance activities. A more detailed analysis of the TPCF that included the development of sample conditional and joint frequency tables allowed the development of a procedure for determining technician job competency within a rating. Furthermore, these analyses revealed areas of suggested modifications of the TPCF prior to its implementation.

 Multiple comparisons between ratings or ships with respect to their average performance levels.

The approach employed in this report to develop comparisons between ratings and between ships was an additive model suggested by a two-way fixed effects analysis of variance with interaction. This required the selection of the appropriate variable, from among SRE, PRE, GRE, WRE, and TP score, upon which to base the comparisons. The variable selected was the one which best met the statistical requirements, i.e., normality of the variable, homogeniety of variances over the main effects, and independence of ship, rating, and cell observations.

4. Degree of association between various demographic variables and the performance variables.



From the demographic information collected on the Personnel Identification Information Form (PIIF), product-moment correlations between the demographic variables, and the performance estimators (SRE, PRE, GRE, WRE, and TP score) were developed. This same approach was applied to each of the eight job activities in order to determine if any job activity related to a particular demographic variable.

These areas are the most rewarding in the sense that they would provide some insight into the merits of the performance measurement technique being researched and of the implications for its application within the U.S. Navy.

RESULTS AND DISCUSSIÓN

Validity of the Performance Estimators

Triserial Correlation Analyses

Initially sample mean reliabilities were developed for each of the four performance estimators (SRE, PRE, GRE, and WRE) for each TPC level. The mean reliability values are the average values of the performance estimators in the TPC levels, therefore, the mean values would be expected to be smaller for a lower proficiency level. The results of this phase of the statistical analysis are presented in Table 5 for technicians at those ships sampled at Location No. 1, Location No. 2, and combined locations.

Employing the results of Table 5, triserial correlation coefficients were developed between each of the performance estimators (SRE, PRE, GRE, and WRE) and Technical Proficiency Checkout (TPC) level for technicians sampled at each location and combined locations. Table 6 presents the resulting triserial correlations.

Comparing the locations, particularly with respect to the SRE, PRE, and GRE, only the triserial correlation coefficients for data collected at Location No. 2 agree with Siegel and Federman's prior results [16]. This observation required a consideration of the appropriateness of triserial correlation to the data collected in this project.

Application of the triserial correlation coefficient involves the following requirements:

- a) the segmented variable is basically continuous and normally distributed; and,
- b) all the segments which together would form a whole normal distribution are present.

Consider again the histogram in Figure 5. Recall that the variable, Technical Proficiency score (TP score), was segmented into one of three levels of technician proficiency. This histogram represents the entire distribution of the segmented variable, which may be taken as continuous and is clearly negatively skewed. A



TABLE 5
MEAN RELIABILITIES

		Location No.	. 1						
		Mean Re	<u>liabilitie</u>	s in Each	TPC Level				
TPC LEVEL	N	SRE	PRE.	GRE	WRE				
1	303	.351	.597	.931	.657				
2	134	.331	.553	.919	.566				
3	145	.427	.625	.923	.557				
		Location No	. 2						
		Mean Re	<u>liabilitie</u>	s in Each	TPC Level				
TPC LEVEL	N	SRE	PRE	GRE	WRE				
1	193	.361	.629	.947	.642				
2	96	.213	.380	.928	.557				
3	78	.141	.338	.878	.502				
			•						
	Combined Locations								
	Mean Reliabilities in Each TPC Level								
TPC LEVEL	N	SRE	PRE	GRE	WRE				
1	496	.355	.608	. 937	.651				
2	230	.283	.483	.923	.562				
3	223	, 328	.524	.907	.538				



TABLE 6
TRISERIAL CORRELATION ANALYSES

		Performance Estimators						
	N	SRE	PRĒ	GRE	WRE			
Location No. 1	582	-0.090*	-0.015	0.031	0.256*			
Location No. 2	367	0.340*	0.335*	0.235*	0.335*			
Combined Locations	949	0.061	0.122*	0.099*	0.283*			

^{*}Significantly different from zero at the α = .05 level.

goodness of fit test for normality [5] was applied to the distribution of TP scores of the 949 technicians. This test statistic will be called the g_1 test statistic.

When the g_1 test statistic was applied to the sample data of TP scores, the resulting test statistic values were

$$g_1 = -.5123$$
, implying $z = 6.4632$.

Therefore, the assumption of normality for TP scores must be rejected for the sample data collected on the population of electronics maintenance personnel (949 technicians in the sample).

⁷The g₁ test statistic is given by g₁ =
$$\frac{\sqrt{N}\sum(X-\overline{X})^3}{[\sum(X-\overline{X})^2]^{3/2}}$$

where X represents an observation, \overline{X} the sample mean, and N is the sample size. If the null hypothesis is that the underlying distribution is normal,

then it has been shown [5] that
$$z = g_1 \sqrt{\frac{(N+1)(N+3)}{6(N-2)}}$$

is approximately normal with mean zero and variance one. In fact a test of the hypothesis that the underlying distribution is normal (at the α = .05 level of signficance) is given by:

reject the null hypothesis of normality if z is greater than 1.96 or less than -1.96.

This particular goodness of fit test has several advantages over the usually applied Kolmogorov-Smirnov tests or the well known Chi-square tests in that, in particular, the population mean and standard deviation need not be known and the test need not be applied just to large samples. Furthermore, this test is more sensitive to departures from normality due solely to skewness than the other two tests [5].



The g_1 test statistic was also applied to the distribution of TP scores at either location. The test statistic values were $g_1 = -.4930$, z = -4.8810 and $g_1 = -.5473$, z = -4.3156 at Location No. 1 and Location No. 2 respectively. Therefore, at both locations the assumption of normality of TP scores must be rejected. These results were verified by the histograms of TP scores for those locations. In both cases these histograms demonstrated the negative skewness of the distribution of TP scores.

Curvilinear Regression Analysis

Essentially due to the non-normality of the TP scores, an alternate analysis was employed in order to determine the degree of association between the predictor variables (SRE, PRE, GRE, and WRE) and criterion variable (TP score). The particular procedure to be employed to achieve this end was a curvilinear multiple regression procedure outlined in Cooley and Lohnes [3]. A few remarks on this subject for the purposes of this report have been provided in Appendix G.

Appendix H provides the results of the analyses of these predictor and criterion variables for the total of 949 technicians sampled. Similar results as found in Appendix H were also developed for Location No. 1 and Location No. 2. Although those printouts are not presented in this report, the essential information from those printouts is given in Table 7^8 along with the essential information from Appendix H.

Consider Table 7 and the evaluation of SRE as a predictor of TP score for the two locations combined. The product-moment correlation between SRE and TP score is .055 (not significantly different from zero at the $\alpha=.05$ level). In attempting to fit a linear, quadractic, and cubic model to the data of SRE values and TP scores, the multiple correlation coefficient (R²) values were .003, .007, and .043 respectively. However, in view of the fact that the residual mean square does not change from the linear to cubic model, it would be just as well to chose the linear model (particularly since R² for the cubic equation is not significantly larger than .007). Therefore, from Appendix H, the best regression equation is

TP score = 5.219 + 0.397 SRE.

Because SRE and TP score are essentially independent, the best estimate of SRE will always be the mean of the observed TP scores, regardless of the observed SRE. This result is further reflected in noticing that the sample mean TP score is 5.350, approximately equal to 5.219 - the TP score intercept of the regression line.

Observing the results of Table 7 for the predictor variables PRE and GRE on the two locations combined, only minimal improvement can be made with these estimates over the SRE. In fact the GRE is almost identical to the SRE and for practical purposes cannot be held to possess significant merit. The PRE is modestly better with a correlation of .1. However, for the PRE, the highest R^2 value does not even reach .05, a long way from a perfect fit with an R^2 value of 1.0. Based upon this analysis, PRE must be termed only slightly better than the SRE and GRE.

The WRE provides the most promising and consistent (over locations) estimator of the four predictors considered. It is most promising in the sense that it provides



⁸All of the computer printouts on the curvilinear regression procedure employed in this report are in terms of "centered data." This technique improves the computation of the printout values by minimizing roundoff errors. Therefore, when reviewing the results of Table 7, one must be concerned with the relative magnitude of the residual mean square in attempting to fit a linear model versus fitting a higher order model.

TABLE 7

CURVILINEAR REGRESSION ANALYSES BY LOCATION+

		Loc	ation N	o. 1		•	
				Type of	Curve	_	
Predictor		Lir	<u>near</u>	Quadr	atic	Cub	ic
Variable	r _{xy}	R ²	s ²	R^2	s ²	R^2	s ²
SRE	069	.005	.002	.007	.002	.072	.002
PRE	009	.000	.002	.043	.002	.047	.002
GRE	.024	.001	.002	.003	.002	.004	.002
WRE	.242*	.058	.002	.058	.002	.062	.002
		. 1.00	ation N	lo 2			
		LOC		Type_of	Curve		
Duaddatas		1 3 4				Cub	
Predictor Variable	r _{xy}	R ²	near_s ²	<u>Quadr</u> R ²	s ²	R ²	s ²
CDF		.083	.003	.083	.003	.090	.003
SRE	.288*						
PRE	.277*	.077	.003	.078	.003	.001	.002
GRE	.212*	.045	.003	.047	.003	.052	.003
WRE	.292*	.085	.003	.085	.003	.089	.003
		Loca	tions Co	m bin ed			
				Type of	Curve		
Predictor		Lin	near	Quadra		Cub	
Variable	r xy	R ²	s ²	R ²	s ²	R^2	s ²
SRE	.055	.003	.001	.007	.001	.043	.001
PRE	.100*	.010	.001	.036	.001	.036	.001
GRE	.087*	.008	.001	.008	.001	.009	.001

 $⁺R^2 = r_{xy}^2$ in the linear case.

WRE

.257*

.066

.066

.001

.071

.001



.001

^{*}Significantly different from zero at the α = .05 level.

 r_{xy} = product-moment correlation coefficient between the predictor variable x and the criterion variable y (TP score).

 R^2 = multiple correlation coefficient.

 s^2 = residual mean square

the highest product-moment correlation coefficient with the selected criterion variable, TP score. However, it cannot be said that it better fits the data than any of the other three estimators with respect to the three types of curves considered.

The curvilinear regression for the predictor variable WRE, Appendix H, page H-6, indicates that a linear curve is the best fit to the data. The regression equation is given by

TP score = 3.571 + 2.95 WRE.

In comparison to the other estimators the WRE can be said to possess moderate validity at best for appraising an absolute level of technician performance. As such its application to the population of electronics maintenance personnel is tenuous.

Analyzing the results at either location again points out the differences between the results obtained at either location. However, this is due mainly to a difference in rxy values and not to R² values for goodness of fit of the linear, quadratic, and cubic models from one location to the next. This may be due to the high degree of unexplained criterion variance in the data at either location and with respect to the locations combined. Scatterplots of the data in those three cases with respect to each predictor variable verified the high degree of dispersion in the data and the lack of any obvious pattern or functional relationship in those plots.

Within Rating Analyses

A factor that may influence the frequency with which a population does not work at a particular job activity is, of course, the appropriateness of the job activity to present-day electronic maintenance activities. The most homogeneous type of subpopulation that would reflect most members of the subpopulation working at the same job activities should be rating. It is mainly for this reason that the subpopulations considered in this report are ratings, and not, for example, ships within locations which should (and did) reflect results similar to those found in Appendix H. In order to perform the most general type of analysis to determine the validity of the performance estimates, a curvilinear regression analysis as previously discussed and employed in Appendix H was applied per rating for each of the four predictor variables - SRE, PRE, GRE, and WRE. Appendix I gives the resulting 32 printouts of the Cooley and Lohnes [3] curvilinear regression analysis. The object is to select for each rating, the best of three possible curves - linear, quadratic, and cubic - for each performance estimator which best fits the data in terms of significantly larger multiple correlation coefficients for smaller residual mean squares. Comparisons between the performance estimators may then be made by performing an appropriate test of hypothesis of the equality of two correlation coefficients (or multiple correlation coefficients). The test that is usually applied is the "Fisher's z" test which employes the asymptotic distribution of the sample correlation coefficient. However this test requires that the distributions of the underlying populations are bivariate normal (see Anderson [1], page 78). Because the distribution of TP scores and the SRE, PRE, and GRE were not normally distributed with respect to the individual ratings, the appropriateness of employing this test is in question. Furthermore the literature seems to be vacant of a discussion of the robustness of the test. Therefore the approach must be in terms of comparing the observed sample product-moment correlations, and, in particular, on the amount of criterion variance explained by the variables. This approach will not necessarily produce a statistically significantly different performance estimator but one which is a more promising estimator, in terms of the sample information.

The following outline represents the essential results of the curvilinear regression analyses presented in Appendix I. The results are presented by rating in Tables 8 through 13 together with observations and recommendations.



TABLE 8

EM CURVILINGAR REGRESSION ANALYSES (N = 97)

				Type of	f Curve		
Predictor Variable	r _{xy}	Line R ²	s ²	Quadi R ²	ratic_s ²	Cub.	ic s ²
SRE	.355*	.133	.009	.136	.009	.176	.009
PRE	.247*	.061	.010	.133	.009	.113	.009
GRE	.301*	.090	.010	.154	.009	. 221	.008
WRE	.492*	.242	.008	. 254	.008	.300	.008

^{*} Significantly different from zero at the α = .05 level.

1. EM rating - Clearly the estimator WRE demonstrates the best fit to the data. The product-moment correlation of .492 indicates at least a fair degree of association of the WRE with the technician's absolute performance level. (Although the R² value is significantly better for the cubic curve over the other two curves, the value of .300 in that case can only indicate a moderate fit of the cubic curve to the data.

TP score = 4.849 - 15.571 WRE + 44.319 WRE² - 26.561 WRE³.

The other estimators do not demonstrate a better fit to the data than the WRE. In view of this and other considerations (e.g., the examination of the scatterplots), the WRE is recommended for application in this rating for those situations in which the probability of effective performance of an individual in the EM rating is required.

TABLE 9
ET CURVILINEAR REGRESSION ANALYSES (N = 173)

		Type of Curve							
Predictor Variable	r _{xy}	Line R ²	s ²	Quadr R ²	ratic s ²	Cub R ²	s ²		
SRE	.318*	.101	.005	.142	005	. 142	.005		
PRE	.366*	.134	.005	.135	.005	.135	.005		
GRE	.508*	.258	.004	. 266	.004	.271	.004		
WRE	.445*	.198	.005	.209	.005	.210	.005		

^{*}Significantly different from zero at the α = .05 level.



2. ET rating - For this rating two significant estimators arise, the GRE and WRE. The WRE is still a promising estimator even though it has a lower product-moment correlation than the GRE (.445 versus .508) and does not fit the data as well. The GRE values may be spurious observations for this rating for it is not significantly higher in any other rating. However, without evidence of this fact it is recommended that either the GRE or WRE be employed in this rating. Furthermore, the values of the product-moment correlations for those estimators reflect at least a moderate degree of association with TPCF results.

Because the ${\sf R}^2$ values do not change appreciably from one type of curve to the next, it is recommended that the linear model be employed for prediction purposes in either case. The linear curves are:

TP score = 2.99 + 4.08 GRE, and TP score = 4.344 + 3.79 WRE.

TABLE 10

FT CURVILINEAR REGRESSION ANALYSES (N = 154)

			Type of Curve						
Predictor Variable	r _{xy}	Line R ²	a <u>r</u> s	Quadi R ²	ratic s ²	Cub R ²	ic s ²		
SRE	.346*	.119	.006	.120	.306	.120	.006		
PRE	.265*	.070	.006	.100	.006	.104	.006		
GRE	.374*	.140	.006	.148	.006	.165	.006		
WRE	.430*	.185	.005	.200	.005	.204	.005		

^{*}Significantly different from zero at the $\alpha = .05$ level.

3. FT rating - This rating again illustrates that the WRE is the most promising estimator. In fact it is almost identical to the results obtained for the ET rating and the WRE. The product-moment correlation of .43 reflects a moderate degree of association with the criterion variable. The R² value of near .2 still illustrates a significant degree of unexplained criterion variance in the data. It is again suggested that the linear model:

TP score = 3.366 + 4.635 WRE

be employed in this rating. The linear model fit is modest (R^2 = .185) but the WRE possesses sufficient promise to be called a good estimator of FT technician proficiency.



TABLE 11

IC CURVILINEAR REGRESSION ANALYSES (N = 58)

		Type of Curve						
Predictor Variable	r _{xy}	Line R ²	ear s ²	<u>Quadr</u> R ²	s ²	<u>Cubi</u> R ²	s ²	
SRE	.361*	.131	.016	. 131	.016	.133	.016	
PRE	.322*	.104	.016	.131	.016	.134	.016	
GRE	.387*	.149	.015	.154	.015	.169	.015	
WRE	.434*	.188	.014	.241	.014	.292	.013	

^{*}Significantly different from zero at the α = .05 level.

4. IC rating - Unfortunately the sample size (N = 58) for this rating is to3 low to place confidence in the obtained correlation coefficient. However, the WRE must be selected as the most valid estimator on all counts. The cubic model

TP score = $-7.57 + 68.26 \text{ WRE} - 113.95 \text{ WRE}^2 + 63.03 \text{ WRE}^3$

is suggested for use in this rating and is a moderate estimate of TP score. The residual mean square is high for this rating only because of the small sample size. In any event this rating reflects the general trend that we have been witnessing and, therefore, the WRE may be employed in this rating.

TABLE 12

RD CURVILINEAR REGRESSION ANALYSES (N = 139)

			Type of Curve							
Predictor		Line	ear	Quadr	atic_	Cub	ic			
Variable	rxy	R ²	s ²	R ²	s ²	R ²	s ²			
SRE	16 8	.028	.007	.093	.007	.131	.006			
PRE	215*	.046	.007	.046	.007	.093	.007			
GRE	.021	.000	. 007	.013	.007	.047	.007			
WRE	051	.003	.007	.021	.007	.041	.007			

^{*}Significantly different from zero at the α = .05 level.



5. RD rating - None of the four estimators that have been discussed in this report should be employed in the RD rating. In part the failure of the estimators in this case must be attributed to the significantly high frequency with which individuals in this rating do not work at the job activities that were used on the evaluation forms.

TABLE 13
RM, ST, AND TM CURVILINEAR REGRESSION ANALYSES

				Type of	f Curve_		
Predictor		Line	ear	Quadi	ratic_	Cut	
Variable	r _{xy}	R ²	s ²	R^2	s ²	R^2	s ²
RM Rating (N = 137)						
SRE	.211*	.044	.007	.057	.007	.068	.007
PRE	.156	.024	.007	.028	.007	.031	.007
GRE	.041	.002	.007	.140	.006	.139	.006
WRE	.366*	.134	.006	.153	.006	.154	.006
ST Rating (N = 152)						
SRE	.199*	.040	.006	.069	.006	.069	.006
PRE	.058	.003	.007	.003	.007	.102	.006
GRE	.099	.010	.007	.134	.006	.178	.006
WRE	.234*	.055	.006	.080	.006	.123	.006
TM Rating (N = 39)						
SRE	.198	.039	.026	.129	.024	.217	.022
PRE	.243	.059	. 025	.094	.025	.153	.024
GRE	.186	.035	.026	.035	.027	.085	.026
WRE	.403*	.162	.023	.177	.023	.224	.022

^{*}Significantly different from zero at the α = .05 level.



6. Rating RM, ST, and TM - The results for these ratings are almost identical and can be discussed as a group. In every case the WRE is the best estimator of TP score. The product-moment correlations are low for the RM and ST ratings, and moderate for TM's. All of the R² values demonstrate a poor fit to the data, indicating an even higher degree of unexplained criterion variance than that found in the EM, ET, FT, and IC ratings. It is recommended that the cubic curve be employed in the ST and TM rating and a linear curve for the RM rating. The corresponding equations may be found from the computer printout of the curvilinear regression for those ratings (pages I-26, I-30, and I-34 respectively). However, the utility of their use as prediction instruments has not been convincingly demonstrated.

Technician Job Competency Evaluation

As hypothesized in the discussion on the validity of the performance estimators SRE, PRE, GRE, and WRE, the TPCF could be employed as a criterion measure of the absolute proficiency level of electronics maintenance personnel. That the TPCF possesses that characteristic was evaluated by Siegel and Fischl [17] for avionics technicians and at that time the TPCF possessed a substantial degree of promise. Siegel and Federman [16] applied the TPCF when appraising the proficiency level of electronics maintenance personnel and in evaluating the SRE, PRE, and GRE.

Due to the high degree of reliance on the TPCF as a criterion measure, it is well to investigate the TPCF and consider whether some of its properties seem to hold up for the data collected in this project. The appropriateness of the job tasks on the TPCF for present-day electronic maintenance activities can be considered, by rating. Furthermore, are the job tasks listed in the same hierarchial order as when originally developed? While it may not be possible to completely answer questions of this type, it is possible to give at least a partial answer based on the present format of the TPCF. By using the properties of the TPCF it is possible to develop a competency level for technicians in each rating and determine whether a particular rating may be performing at a certain proficiency level.

Job Task Analyses

The reader will recall that the job tasks listed on the TPCF are ordered with Job Task No. 1 being the easiest and Job Task No. 8 the most difficult. The following is a list of the job tasks as found on the TPCF:

Task Description

(easiest) :.

- : Capable of employing safety precautions.
- 2. Capable of <u>replacing</u> most of unit's equipment.
- 3. Capable of removing most of unit's equipment.
- 4. Capable of following block diagrams.
- 5. Capable of knowing <u>relationship of equipment to other</u> <u>related equipment</u>.



- 6. Capable of calibrating most of unit's equipment.
- Capable of trouble-shooting/isolated malfunction(s).

(most difficult)

8. Capable of maintenance.

Table 14 is the frequency of occurrence of men CHECKED OUT and NOT CHECKED OUT by rating over the eight job tasks without reference to their performance on other job tasks. For example, of the 97 men in the EM rating, 95 were CHECKED OUT on Job Task No. 1, 87 were CHECKED OUT on Job Task No. 2, ..., and 55 were CHECKED OUT on Job Task No. 8. Similarly, for the other ratings.

From Table 14 the hierarchial difficulty of tasks is generally represented in the EM, ET, FT, IC, ST, and TM ratings. This is evidenced by the fact that one would expect progressively fewer men to complete a more difficult task. Except for the ST rating, Job Task No. 6 - Capable of calibrating most of this unit's equipment with which his rating is concerned - seems out of place in that fewer individuals are CHECKED OUT on this task than the next more difficult task, Job Task No. 7. In fact the better position for Job Task No. 6 would be after Job Task No. 8. In that case the job tasks would follow more closely a hierarchial classification.9

In any event it can be said that Job Task No. 6 is no longer of such prominence as originally thought. This is a result of the introduction of more sophisticated electronic equipment aboard ships in recent years. This equipment frequently consists of integrated circuitry and compact electronic modules requiring little if any calibration. Furthermore, test instrumentation calibration is also in demise because less test equipment is being employed. Normally a defective component is replaced en toto without regard aboard ship for finding the particular fault in the component. Those functions are beyond normal shipboard electronic maintenance activities.

In the RD rating there are a significant number of individuals NOT CHECKED OUT on many of the tasks. This is probably due more to a nonapplicability of those tasks to that rating than to a lack of sufficient training. This is to a degree verified by the findings in Appendix E where it was demonstrated that for most job activities there are a very high proportion of individuals in that rating who either do not work at the job activities or received $\Sigma UE = 0$ and $\Sigma UY = 0$ from their supervisor. It is possible that the nonapplicability of many of the tasks to the RD rating did give erroneous TP scores for this rating and depressed the validity coefficients (see Table 12). A more detailed study than the present one would be required to come to a definite conclusion on this issue.

As shown in Table 14 and the RM rating, Job Tasks 2 and 4 are not in agreement with the underlying order of job tasks, but Job Task No. 6 is in a proper position for this rating. It seems that a reordering of the first five job tasks on the TPCF would allow this instrument to be more applicable to this rating. It would probably require a minor research effort to revise the tasks in the RM rating.



⁹It is important to note that none of the conclusions of previous sections would be altered with this modification because the TP score (or TPC level) is independent of the hierarchial classification. Those variables are dependent only upon the number of job tasks an individual was CHECKED OUT on, and not upon the order of the CHECKED OUT tasks.

TABLE 14

JOB TASK ANALYSES

		_						_		
			EM	Frequenc	les (N = 9	97)		_		
Job Task	tio.	1	2	3	4	5	6	7	8	
Checked	Out	95	87	87	77	75	43	63	55	
Not Checked	Out*	2	10	10	20	22	54	34	42	
			ET	Frequenci	les (N = 1	(73)				
Job Task	No.	1	2	3	4	5	6	7	8	
Checked	Out	168	147	164	165	126	96	129	132	
Not Checked	Out	5	26	9	8	47	77	44	41	
			FT	Frequence	les (N = 1	1541				
Job Task	Na.	1	2	3	4	5	6	7	8	
Checked		148	125	134	143	116	93	100	102	
Not Checked		6	29	20	11	38	61	54	52	
		_		•	les (N = 5	-				
Job Task		1	2	3	4	5	6	7	8	
Checked		57	45	49	44	41	26	33	27	
Not Checked	Out	i	13	9	14	17	32	25	31	
			RD	Frequenct	es (N = 1	39)				
Job Task	No.	7	2	3	4 .	5	6	7	8	
Checked	Out	136	27	38	50	97	8	8	26	
Not Checked	Out	3	112	101	89	42	131	131	113	
			RM	Frequenci	es (N = 1	37)				y
Job Task	No.	1	2	3	4	5	6	7	8	
Checked	Out	135	88	100	88	113	46	27	13	
Not Checked	Out	2	49	37	49	24	91	110	124	
		•	ST	Frequenci	es (N = 1	52)				
Job Task	No.	1	2	3	4	5	6	7	8	
Checked		144	118	119	117	121	103	77	80	
Not Checked	Out	8	34	33	35	31	49	75	72	
			_,		4					
lab Taxt	u.	,			es (N = 3		_	_		
Job Task		1	2	3	4	5	5	7	8	
Checked		38	31	32	31	34	14	18	8	
Not Checked	UUL	1	8	7	8	5	25	21	31	

^{*}The reader is cautioned that a score of NOT CHECKED OUT on a task does not differentiate between whether a technician really cannot be trusted with doing the task on his own without direct supervision or whether he was given a score of NOT CHECKED OUT because he does not work at that task.



Conditional Frequencies for Job Tasks

Appendix J provides tables of a more detailed analysis of some observations on the TPCF. In particular the Job Task Conditional Frequency is given by rating; i.e., for a given task, the number of individuals who were CHECKED OUT on each easier task given they were at most CHECKED OUT on the given task. Those tables should provide more insight into the hierarchial classification of the job tasks and also a relative level of technician competency by rating.

To illustrate the use of the tables prepared in Appendix J, consider Table 15. This table represents the Job Task Conditional Frequencies for the EM rating as can be found in Appendix J, page J-4.

TABLE 15

EM RATING CONDITIONAL FREQUENCIES (N = 97)

-			Job	Task			
8	7	6	5	4	3	2	1
55	46	33	50	50	50	52	55
	17	4	13	14	17	17	17
		6	6	5	5	4	6
			6	5	6	6	6
				3	3	3	3
					6	5	6
						0	. 0
							2

The number of men that were not checked out on any job task is 2.

From the first row of Table 15, 55 of the 97 EM's were CHECKED OUT at most on Job Task No. 8. Of those 55 men, 46 were also CHECKED OUT on Job Task No. 7 (the next easier task), 33 were CHECKED OUT on Job Task No. 6, ..., and finally all 55 men were CHECKED OUT on Job Task No. 1 (the easiest task). Continuing, from the second row of the Conditional Frequency table, 17 of the 97 EM's were CHECKED OUT at most on Job Task No. 7 (i.e., none of those 17 men were CHECKED OUT on Job Task No. 8). Again, of those 17 men, 4, 13, 14, 17, 17, and 17 were also CHECKED OUT on Job Task No.'s 6, 5, 4, 3, 2, and 1 respectively. Finally, only 2 men were CHECKED OUT on Job Task No.'s No. 1, i.e., none of those 2 men were CHECKED OUT on a more difficult task. Also, there were 2 men in the EM rating that were not CHECKED OUT on any of the eight tasks. Clearly then one would expect any row of a Conditional Frequency table to contain almost identical entries, i.e., if the job tasks are truly hierarchial and representative of tasks in that rating.



Consider now Appendix J, page J-4, and the Conditional Frequency table. From the first two rows of that table, most of the EM's CHECKED OUT at most on Job Tasks 7 and 8 were also able to complete the easier tasks, except for Job Task No. 6. Again, that task is very much out of place in the order of job tasks. There is a low degree of incompetence present for 13 percent of the EM's could be CHECKED OUT at most on no more than Job Task No. 4. However, it must be realized that this result could also be due to the possibility that many of those individuals simply do not work at a more difficult task than Job Task No. 4.

The ET, FT, ST, and TM ratings (Appendix J, pages J-5, J-6, J-10, and J-11 respectively) demonstrate practically the same results for the Conditional Frequency tables as the EM's. For those ratings 8, 10, 9, and 13 percent were CHECKED OUT at most on no more than Job Task No. 4 respectively. The IC, RD, and RM ratings demonstrated 21, 30, and 17 percent. These remarks illustrate the practicality of the TPCF as an instrument for evaluating the competency level of particular ratings and for isolating those job tasks requiring further training by electronic maintenance technicians.

It is to be noted that as in the EM rating, Job Task No. 6 is inconsistent in most of the other ratings with respect to its proper order in the eight job tasks.

Joint Frequencies for Job_Tasks

Appendix J also provides the Job Task Joint Frequencies by rating. For each rating this is the number of individuals who were CHECKED OUT on any two given tasks. Use of the Joint Frequency tables is illustrated by Table 16.

TABLE 16
EM RATING JOINT FREQUENCIES (N = 97)

1	2	3	4	5	6	7	8
95	87	87	77	75	43	63	55
	87	84	73	70	40	63	52
		87	71	69	40	62	50
			77	69	40	56	50
				75	42	56	50
					43	35	33
						63	46
	•						55

From the first row of Table 16, 95 of the 97 EM's were CHECKED OUT on Job Task No. 1, 87 (of the 97 EM's) were CHECKED OUT on both Job Task No.'s 1 and 2, 87 were CHECKED OUT on Job Task No.'s 1 and 3, ..., and 55 (Of the 97 EM's) were CHECKED OUT on Job Task No.'s 1 and 8. Similarly, from row 2 of the Joint Frequency table, 87 (of the 97 EM's) were CHECKED OUT on Job Task No. 2, while 84, 73, 70, 40, 63, and 52 (of the 97



EM's) were CHECKED OUT on both Job Task No. 2 and Job Task No. 3, 4, 5, 6, 7, and 8 respectively. It is well to note that the diagonal entries of this table corresponds exactly to the frequency entries for the CHECKED OUT case given in Table 14.

Just as in Table 15 on the Conditional Frequencies one would expect that any row of that table to have almost identical frequencies, in a Joint Frequency table one would expect any column to contain identical frequencies. For example, completion of Job Task No. 5 should also insure completion of Job Tasks 1, 2, 3, and 4. Likewise, the rows of the Joint Frequency tables should illustrate decreasing frequencies for being CHECKED OUT on a task does not insure being CHECKED OUT on a more difficult task.

From the Joint Frequency tables (Appendix J, pages J-4, J-5, J-6, J-7, and J-11) for the EM, ET, FT, IC, and TM ratings respectively, the previously mentioned patterns that are illustrated by a Joint Frequency table are demonstrated except for Job Task No. 6. This is consistent with the analysis on the Conditional Frequency tables. The ST rating conforms almost exactly to the expected patterns for a Joint Frequency table. Unfortunately there are a few instances in which the Conditional Frequency table for that rating deviates from the expected pattern. In any event the task descriptions on the TPCF are very descriptive of the ST rating job tasks.

The RD rating Joint Frequencies (page J-8) illustrate the general lack of patterns characteristic of a Joint Frequency table. The RM rating Joint Frequencies (page J-9) more closely conform to the desired patterns but still leave something to be desired.

Multiple Comparisons of Ships and Ratings

From either a research or applied point of view it is often necessary to compare the performance levels of various ships or ratings. Such comparisons may be employed to compare ships (or ratings) relative to mean performance levels. However, such comparisons can also offer a basis upon which to formulate a decision as to whether a particular ship configuration of men and/or equipment is, or is not, detrimental to ship/rating performance. In a similar fashion, policy making decisions also can be evaluated with respect to their influence on the performance of ships or ratings affected by that policy.

In this section the main subject of concern will be the development and justification of a technique for performing pairwise comparisons between ratings and between ships on the basis of mear performance levels with respect to some variable. The variables to be considered are the SRE, PRE, GRE, WRE, and the selected criterion variable TP score. In general this subject falls under the area of multiple comparisons arising from an Analysis of Variance (AOV) of ship/rating effects. The particular model to be considered will be a two-way fixed effects design with an unequal number of observations for the ship/rating

$$Y_{ijk} = \mu + \alpha_i + \beta_j + \delta_{ij} + e_{ijk}$$

where e_{ijk} is an error term associated with the above model. In this situation it is a two-way fixed effects design with an unequal number of observations in the i-jth cell.



 $^{^{10}{\}rm It}$ is assumed that each observation y_{ijk} on the k^{th} individual in the i^{th} ship in the j^{th} rating is due to an overall effect (µ), a ship effect (α_i), and a rating effect (β_j). Furthermore, a general model addresses the question of the existence of significant row-column interaction (δ_{ij}) upon each observation. In particular this report addresses the additive model of the form

combinations (see Table 1). Choice of this model requires the verification of normality of the selected performance variable over the main effects, homogeneity of the variances of the main effects and independence across and between the main effects. The main effects are "ships" and "ratings." That performance variable to be selected for performing the multiple comparisons will, most naturally, be the one which best conforms to the requirements for use of the model being considered and is most associated with technician on-the-job performance.

Distributional Properties of the Performance Estimators

The g_1 test statistic for normality was applied to the distribution of SRE, PRE, GRE, and WRE values for the 949 technicians in the sample (combined locations). Table 17 illustrates the results of this analysis.

TABLE 17
g1 TEST STATISTIC VALUES FOR THE PERFORMANCE ESTIMATORS

	SRE	PRE	GRE	WRE
gl	.5807	4565	-4.7092	.7939
z	7.3265	-5.7599	-59.4129	~.5823*

^{*}Because WRE scores (Figure 4) are almost normally distributed, a test statistic [4] similar to the g₁ test statistic but more sensitive to departures from normality due to kurtosis was applied to the WRE scores.

Recall that rejection of normality results if $|z| \ge 1.96$.

Clearly then the only estimator which appears to be normally distributed is the WRE. This same analysis was applied by location and rating with practically the same results. In each case only WRE scores proved to be normally distributed, except for isolated incidents of the other estimators. This characteristic of WRE scores is not surprising as this situation is essentially an application of the Central Limit Theorem as found in probability theory, see [12]. The lack of normality on the part of the other estimators is not necessarily an undesirable feature, but it is true that this exercise does point out yet one more desirable feature of the WRE, namely, its general normality.

Choice of a Variable

As previously noted it is necessary to choose an appropriate variable upon which to base the multiple comparisons of ships and ratings. From the previous section - Distributional Properties of the Performance Estimators - it was observed that WRE scores possessed the unique characteristic among the performance variables (SRE, PRE, GRE, and TP score) of being normally distributed. This characteristic of WRE scores is most helpful particularly since homogeneity of variances in the ships, ratings, and



 $^{^{11}}$ A search of the literature failed to locate a particular test to employ for independence in the case of unequal numbers for the ship/rating combinations. See Anderson [1] for the case of equal numbers.

An identical AOV was executed as above but for the four ratings (EM, ET, FT, and IC) for which the performance measurement technique was most promising. Table 19 reflects the results of this AOV.

TABLE 19
AN ANALYSIS OF VARIANCE FOR 21 SHIPS AND 4 RATINGS

Source	Degrees of Freedom	Sum of Squares	Mean Square	F Value
Total	482	192.0419		
Between Ships	20	2.669	.1334	3.884*
Between Rating	3	.5621	.1874	5.4598*
Ship by Rating Interaction	60	4.0437	.0674	2.2975*
Residual	398	11.6747	.0294	

^{*}Indicates significance at the α = .05 level.

In either of the above two AOV tables there exists significant ship/rating interaction. This situation will not permit the comparing of any two of the main effects under a two-way design because of the confounding of the main effects with the interaction terms. The only recourse is to test for significant differences in each ship, across ratings, or in each rating, across ships. If the interaction terms were not significant, it would be an easy matter to compare ratings and to compare ships by employing the entire set of data represented in the AOV's of Table 18 or Table 19, see again Scheffe [15].

Testing for Significant Ship (Rating) Effects for a Particular Rating (Ship)

In order to develop comparisons between ships (for a fixed rating) and to develop comparisons between ratings (for a fixed ship), this report will employ an AOV model 13 resulting from a one-way fixed effects design with unequal numbers of observations across the main effects. Homogeneity of the variances across the main effects must be assumed in this case for the usual tests are large sample tests and inapplicable for ship/rating numbers less than ten (see Layard [14] and Table 1).

For present purposes only the AOV's, and subsequent multiple comparisons, will be developed on the 21 ships participating in the project and the four ratings EM, ET, FT, and IC. This will involve 25 different AOV's to be constructed and subsequently developed for multiple comparison of the main effects if the AOV indicates significant main effects.

$$Y_{ijk} = \mu + \alpha_i + e_{ijk}$$
 (j is fixed and i=1, ..., 21)

for testing for significant ship effects, and a mode of the form

$$Y_{ijk} = \mu + \beta_i + e_{ijk}$$
 (i is fixed and j=1, ..., 4)

for testing for significant rating effects. The subscript k is an index representing the $k^{\mbox{th}}$ man in that main effect.



¹³This report will employ a model of the form

Initally the hypothesis of equality of mean WRE scores was tested for the four ratings EM ,ET, FT, and IC for each of the 21 ships. This hypothesis was accepted for 12 of the 21 ships. On the remaining 9 ships Scheffe's multiple comparison technique was applied in searching for a possible significant difference between any two (of the four) ratings on the basis of mean WRE scores. Table 20 reflects the results of these multiple comparisons, each derived from a separate Amalysis of Variance. (The resulting 21 AOV tables are too extensive to be presented in this report).

Finally the hypothesis of equality of mean WRE scores was tested for the 21 ships for each of the four ratings EM, ET, FT, and IC. This hypothesis was rejected for each of the four ratings. No significant differences were found between any pair of the ratings using Scheffe's multiple comparison technique. This is not an inconsistent result, for this could be a result of trying to perform too many multiple comparisons (therefore, increasing the comparison error rate) and/or the effect referenced in the footnote of Table 20.

TABLE 20
MULTIPLE COMPARISON OF FOUR RATINGS ON NINE SHIPS*

Ship	Highest Mean WRE			Lowest Mean WRE
a	<u>EM</u>	FT	ET	IC
b	FT	ET	EM	10
С	EM	<u>ET</u>	IC	FT
d	ET	FT	EM	<u>IC</u>
е	<u>IC</u>	ЕМ	ET	FT**
f	IC	ET	FT	<u>EM*</u> *
g	FT	<u>IC</u>	EM	ET
h	EM	ET	IC	
1	ЕМ	ET	FŢ	

^{*}For each ship a line under two ratings signifies no significant difference between the ratings.



^{**}It is important to note that while no differences in rating means were detected, the rejection of the null hypothesis of the equality of rating means is not in error. This only implies that some contrast other than those contrasts testing rating differences lead to that rejection, see Ferguson [8, pp. 279-283].

Association Between Demographic and Performance Variables

Usually it is of interest from either a research or applied point of view to estimate school success, final grade, etc. of an individual in some training program or school by employing such predictor variables as the Basic Test Battery (BTB) scores - GCT, ARI, MECH, and CLER scores or combinations of those scores. Typically BTB scores are used to predict actual school success which in turn is viewed as being the best available measure of potential job success. Seldom, however, are direct measures of on-the-job performance available as those developed in the current study. Accordingly the relation-ship between the BTB scores - GCT, ARI, and Skill (GCT + ARI) - and measures of on-the-job performance - SRE, PRE, GRE, WRE, and TP score - was investigated.

Besides the BTB scores, other available predictor variables were also considered under the general heading of demographic variables. In this report the demographic variables to be considered are: months known by supervisor, months on current job assignment, GCT score, ARI score, SKILL (GCT + ARI) score, and A, B, and C school final grade. The performance variables that will be employed to relate to on-the-job technician performance are the eight reliability ratios relative to the job activity factors on the JPQ ANSWER SHEET, and the variables SRE, PRE, GRE, WRE, and TP score.

Utility of Demographic Variables in Performance Prediction

In order to determine the utility or effectiveness of the demographic variables previously introduced to relate or be associated with on-the-job technician performance, product-moment correlation coefficients were developed between those demographic variables and the performance variables. From this analysis it is possible to infer the extent which the demographic variables may be viewed as linear predictors of the demographic variables.

Demographic Variable Prediction on Combined Locations

Table 21 reflects the product-moment correlation coefficients that resulted when the 949 technicians were considered. In that table there are many product-moment correlation coefficients significantly different from zero, but they are of such low magnitude. Now from a linear prediction standpoint, the multiple correlation coefficient (R²) is related to the product-moment correlation coefficient (r_{xy}) by $R^2 = r_{xy}^2$. (x is one of the demographic variables and y is a performance variable.) Therefore, from Table 21, even for the largest r_{xy} value observed, 14 namely 0.217, the R^2 value for this case is still a very low 0.047 for the degree of fit of the linear model to the data (x is SKILL score and y is TP score). However, even with the low correlations represented by the BTB scores, those scores do offer some hope of predicting on-the-job performance as measured by the TPCF. The correlation coefficients reflect the degree of association that is typically found between predictor variables and on-the-job performance. However, these results show considerable promise for further research in the area of the development of on-the-job performance measures.

Demographic Variable Prediction in the EM, ET, FT, and IC Ratings

In a similar manner product-moment correlation coefficients were developed between the demographic and performance variables for the technicians sampled from the four more promising ratings EM, ET, FT, and IC. In Table 22 are the results of this analysis. From this table the same observations as above can be made for this case since no product-moment correlation is greater than 0.217, the largest previously observed. Also, as in the above, some hope is offered by BTB scores for predicting on-the-job performance as measured by the TPCF. The correlation coefficients, though suppressed, are still of sufficient magnitude to warrant some promise for further research.

¹⁴Excluding B school final score which has such a low sample size represented.



TABLE 21

DEMOGRAPHIC CORRELATIONS ON EIGHT RATINGS

		Reliability Ratios							
	N	Ref Mat	Equip Op	Cir An a l	Pers Rel	Elec Safe	Inst	Elec Rep	Elec Cog
Mo known by super	948	0.032	0.068*	-0.006	-0.020	0.038	0.001	-0.069*	-0.036
Mo on current assign	948	0.029	0.106*	-0.005	-0.034	9.057	-0.013	0.012	-0.009
GCT scores	909	-0.027	-0.C34	-0.048	-0.078*	-0.014	-0.096*	-0.050	-0.076*
ARI scores	909	0.018	-0.014	-0.030	-0.086*	-0.031	-0.076*	-0.039	-0.022
SKILL scores	909	-0.006	-0.029	-0.046	-0.095*	-0.027	-0.100*	-0.052	-0.057
A school final grade	774	0.089*	0.090*	0.046	0.038	0.017	0.082*	0.104*	0.101*
B school final grade	16	-0.022	0.379	0.227	0.231	0.178	0.323	-0.044	0.300
C school final grade	264	0.016	0.077	0.006	0.000	-0.054	-0.009	0.119	0.079
•			Perf <u>o</u> m	nance Estima	tors				
	N	SRE	PRE	GRE	WRE	TP Score			
Mo known by super	948	-0.029	-0.038	0.055	0.011	0.086*			
Mo on current assign	948	0.067*	0.036	0.044	0.070*	0.191*			
GCT scores	909	-0.119*	-0.100*	-0.040	0.018	0.186*			
ARI scores	909	-0.076*	-0.066*	-0.013	0.037	0.193*			
SKILL scores	909	-0.111*	-0.095*	-0.034	0.032	0.217*			
A school final grade	774	0.142*	0.096*	0.084*	0.098*	0.070*			
B school final grade	16	0.339	0.387	0.167	0.265	-0.173			
C school final grade	264	0.041	0.027	0.047	0.092	0.091			

^{*}Significantly different from zero at the α = .05 level.



TABLE 22
DEMOGRAPHIC CORRELATIONS ON FOUR RATINGS

			<u>_</u>		ity Ratios				
	N	Ref Mat	Equip Op	Cir Anal	Pers Rel	Elec Safe	Inst	Elec Rep	Elec Cog
Mo known by super	481	-0.043	-0.021	-0.097*	-0.055	0.025	-0.085	-0.062	-0.084
Mo on current assign	482	-0.035	0.093*	-0.034	-0.044	0.064	0.032	0.081	-0.019
GCT scores	459	-0.045	-0.001	0.026	-0.072	-0.047	-0.013	0.008	-0.007
ARI scores	459	-0.002	0.014	0.026	-0.094*	-0.035	0.006	-0.008	0.032
SKILL scores	459	-0.026	0.007	0.030	-0.093*	-0.046	-0.003	0.001	0.014
A school final grade	383	0.086	0.107*	0.047	0.009	0.057	0.063	0.067	830.0
B school final grade	10	-0.342	0.322	0.103	0.129	-0.201	0.328	-0.225	0.250
C school final grade	164	0.004	0.051	-0.057	0.016	-0.103	-0.097	-0.018	0.085
			_						
			Perform	ance <u>Estima</u>	tors				
	N	SRE	PRE	GRE	WRE	TP Score			
Mo known by rater	481	-0.083	-0.096*	-0.005	-0.040	0.089*			
Mo on current assign	482	0.017	0.014	0.088*	0.062	0.214*			
GCT scores	459	-0.050	-0.081	0.008	0.065	0.172*			
ARI scores	459	-0.033	-0.056	0.016	0.048	0.144*			
SKILL scores	459	-0.047	-0.077	0.014	0.064	0.178*			
A school final grade	383	0.071	0.049	0.112*	0.125*	0.167*			
B school final grade	10	0.028	0.213	-0.010	0.027	-0.265			

C school final grade 164 -0.003 -0.045 -0.019 0.080 0.135



^{*}Significantly different from zero at the α = .05 level.

SUMMARY AND CONCLUSIONS

Summary of Data Analyses

Evaluation of the Performance Estimators

When evaluating a performance estimator it is initially of interest to consider the degree to which that estimator desc in fact measure an individual's on-the-job performance. In order to address this particular aspect of the analyses it was necessary to choose an appropriate criterion measure of on-the-job performance. The variable selected as the criterion measure in this report was Technical Proficiency 'P' score (or Technical Proficiency Checkout (TPC) level) as derived from the Technical Proficiency Checkout Form (TPCF). Prior research efforts demonstrated that this variable was a viable criterion variable. Initially triserial correlation coefficients were developed between the TPC level and each of the performance estimators (SRE, PRE, GRE, and WRE) being evaluated. Essentially due to the lack of normality of the continuous variable TP score, triserial correlation had to be considered inappropriate for the data collected in the course of this project. Finally a curvilinear regression analysis was executed between each of the performance estimators and TP score with an emphasis upon a least-squares interpretation of the corresponding results.

When the entire sample of 949 technicians was considered, relatively low multiple correlation coefficients corresponding to a linear, quadratic, and cubic model were obtained. Furthermore, low product-moment correlation coefficients were obtained for some estimators. Due to these results it was decided that a more appropriate approach to analyzing the data would be an analysis by rating. That is also a more appropriate approach for job tasks or 33b activities would be more homogeneous within a particular rating.

When the individual ratings were considered it could be concluded that of the four performance estimators - SRE, FRE. GRE, and WRE - the WRE possesses the greatest degree of premise in all ratings. The only exceptions occur for the ET and RD ratings in which the GRE is slightly better and none of the estimators apply, respectively. The WRE is a moderate estimate of absolute technician performance in the EM, ET, FT, and IC ratings and a modest estimate of absolute performance in the RM, ST, and TM ratings. It is recommended that the WRE be applied, in particular, upon the EM, ET, FT, and IC ratings for purposes of comparing individuals or groups of individuals in those ratings. In fact, it is suggested that in most analyses where the performance of an individual, or group of individuals in those ratings is needed, (i.e., an estimate of the probability of effective performance is required) that the WRE be employed. Usually these analyses would involve a classification or discrimination of such individuals on the basis of their present-day on-the-job performance. However, use of the WRE in the RM, ST, and TM ratings is not as justified and, the efore, if used in those ratings, it should be only with utmost caution with regard to its low validity, as determined by the TPCF.

The Criterion Measure of On-the-Job Performance

Because the TPCF was employed as a criterion measure of technician on-the-job performance, an analysis of the extent to which its properties were verified by the data collected in this project was considered. To achieve this end Conditional and Joint Frequency tables were developed by rating to determine whether the hierarchial classification of the job tasks was in effect and whether the job tasks were appropriate



to present-day electronic maintenance activities. Results from those tables revealed that only one job task - capable of <u>calibrating</u> most of this unit's equipment with which his rating is concerned - was more dated than the other job tasks. With this exception the job tasks seemed to follow a hierarchial classification in all but the RM rating, and to a lesser degree in the RD rating. From these analyses it is suggested that the following alterations be completed by a potential user before the TPCF is employed either in an operation or research context:

- 1. An updating of the job task descriptions by rating. It is not likely that a lint of task descriptions can be developed that are characteristic of the population of electronics maintenance personnel without being too general and not specific enough for their intended use in any rating.
- 2. A verification of the hierarchial classification of the updated task descriptions by rating.
- 3. Revision of the TPCF to include an answer column that would reflect whether the technician actually works at the task description being considered.
- 4. Development of alternative test statistics that would reflect the hierarchial classification of the job tasks which TP score does not do. Suggested random variables to consider are:
 - X = the highest job task the technician is CHECKED OUT on, or
 - Y = (X + TP score)/2.

Use of X in a rating is most warranted if the hierarchial classification is correct; if not, use of Y will "average" the effects of X and TP score.

5. A validation by rating that the TPCF is highly predictive of on-the-job performance for that rating.

In conclusion it can be said that the TPCF represents the most promising of the performance instruments evaluated. It is certainly the most easily administered and offers the least amount of confusion to the evaluator. Furthermore, the possibility of deriving competency levels from the Conditional Frequency tables offers an alternative procedure for comparing ratings or groups of technicians within a rating on the basis of the proportion of individuals CHECKED OUT at most on a certain task. Although such comparisons were not attempted in this report, it certainly is an appropriate topic for future research on the TPCF.

Comparisons Between Ships and Ratings

Multiple comparisons between ratings and between ships was accomplished in this report by means of appropriate fixed-effect additive Analysis of Variance (AOV) models. Use of these models required the homogeneity of variances across the main effects (ships and/or ratings), normality of the variable upon which the comparisons are to be based, and independence of the observations across ships and ratings. The performance variable selected, from among SRE, PRE, GRE, WEE, and TP score, was that one which best conformed to the above requirements and could be most associated with technician on-the-job performance. Initially this involved the investigation of the distributional properties of each of the performance variables. The distributions of the variables SRE, PRE, GRE, and TP score were all shown to be non-normal for the subpopulations considered. This



characteristic of those variables is only detrimental in AOV analyses for the case of unequal number of observations for the ship-rating combinations, which was the case for the data collected in this report (see Table 1). Essentially the difficulty arises when one attempts to apply a test of homogeneity of variances for the variable considered. The standard tests for homogeneity of variance require normality of the variable considered.

Only the variable WRE could be termed normally distributed for all the subpopulations considered. Furthermore, it was shown to be the most promising type of performance estimator, particularly in the EM, ET, FT, and IC rating. In view of these and other related considerations it was the most optimal choice of a performance variable upon which to base the multiple comparisons. Unfortunately, not all the requirements for use of the AOV models to be employed were satisfied by the WRE, however, the techniques employed to that end demonstrate the procedures that one should go through in order to validly apply the AOV models.

Initially a two-way fixed effects design was applied to the data. However, because significant interaction was found to exist between ships and ratings, this necessarily forced multiple comparisions between ships (ratings) to be performed for a fixed rating (ship). Twenty-five AOV's were executed in that event in order to detect significant mean WRE scores between the four more promising ratings (EM, ET, FT, and IC) for each of the 21 ships in the project and between the 21 ships for each of these four ratings. No particular pattern seemed to emerge across ships as to which rating is more proficient (in terms of mean WRE scores). Furthermore, no pairwise significant difference was detected between any two ships (for any of the four ratings).

In conclusion it would seed that the WRE is the most appropriate performance variable upon which to base the multiple comparison of ships and ratings. However, it would be a valuable exercise for a user to also include the other candidate performance variables and choose that variable most appropriate to the characteristics of the data collected. Likewise, the application of the appropriate variable for particular ship-rating configurations may yield insignificant ship-rating interaction, thus eliminating some of the difficulties encountered in this report. In particular it is the user's responsibility to test the requirements and various configurations before deciding on a particular technique for reforming multiple comparisons. It is hoped that the procedure employed in this report offers some guidelines to the potential user for application of the performance measurement technique being researched.

Analysis of Demographic Data

In addition to the performance data collected on the TPCF and JPQ ANSWER SHEET, demographic data were collected by means of the Personnel Identification Information Form (PIIF). The demographic variables that were subsequently studied were: months known by rater, months on current job assignment, GCT and ARI scores, and A, B, and C school final grades. Product-moment correlation coefficients were developed between those variables and the SRE, PRE, GRE, WRE, TP score and each of the eight job activity reliability ratios. Although the resulting correlations were low, some hope was offered by GCT and ARI scores for prediciting on-the-job performance as measured by TP scores. It is felt that further research in this area may reveal that Basic Test Battery scores have some utility for predicting job performance levels.

Observations on the Use of Composite Reliability Values

It will be observed that many of the analyses chosen for analyzing the data rest upon the particular characteristics observed in the data, e.g., distributional properties



of the data. This is a reasonable and preferred procedure to follow in order to form an opinion based on the data collected. However, there was one necessary, and artificial, revision of the fundamental characteristics of the original data that was required in order to arrive at a numeric estimate of performance based on the estimators SRE, PRE, GRE, and WRE. That revision, of course, involved the adoption of a convention to estimate technician performance for those cases in which the technician either does not work at a job activity or received $\Sigma \mathsf{UE} = \Sigma \mathsf{UI} = \mathsf{O}$. The choice of a composite reliability value (Appendix F) to estimate technician performance in those cases cannot be viewed as the most optimal choice. However, it is certainly a reasonable and efficient means of employing the original data in order to achleve performance estimates in those cases. Reasonable in the sense that this estimate is derived as the individual job activity reliability ratios are derived, i.e., by totaling the number of UE's and UI's observed and forming the ratio Σ UE/ $[\Sigma$ UE + Σ UI]. Therefore, it should be no less objectional than are the reliability ratios. Furthermore, it is efficient in the sense that no complicated mathematical procedure is required in order to derive the composite reliability values.

The composite reliability values were derived by location (Table F-2) and it was those values which were subsequently employed throughout the remainder of the analyses. A better procedure would have been to derive a composite reliability value for the total of N = 949 technicians and employ those values for subsequent analyses rather than by the use of "per location" composite reliability values. However, as can be observed in Table F-2, the composite reliability values do not differ significantly by location except in two cases: Job Activity No. 7 in the RD rating and Job Activity No. 3 in the TM rating.

Effect of the Convention on the Performance Estimates SRE, PRE, and GRE

The Convention in the RD Rating

For the RD rating the composite reliability value is 0.0 for Job Activity No. 7 at Location No. 2 (Table F-2). However, at Location No. 2 the RD's received SRE = PRE = 0.0 (see p. 3 and footnote 2) because all (but one) RD at that location either did not work at Job Activity No. 7 or received $\Sigma UE = \Sigma UI = 0$ (see Table E-1). In addition to the low SRE and PRE scores resulting from the use of this convention at Location No. 2, it must be pointed out that low scores were likewise recorded for the TPCF at that location (Table 4). This effect would result in a high degree of association between the predictor and criterion variables for the RD rating at Location No. 2 because low predictor scores are corresponding to low criterion scores. This situation, however, is unique to the RD rating at Location No. 2.

At Location No. 1 the situation of a high degree of association between the predictor and criterion variables was not illustrated for the RD rating. As at Location No. 2, most of the RD technicians are in the lowest TPC levels (Table 3) and every RD at Location No. 1 did not work at Job Activity No. 7 or received $\Sigma UE = \hat{\Sigma} UI = 0$ (Table E-1). However, all the composite reliability values are high for Location No. 1 (see Table F-2) implying that most likely the SRE and PRE are different from zero and of a large magnitude. This characteristic of the SRE and PRE would imply an inconsistency with the low TPC levels recorded at Location No. 1. Namely, an inverse relationship has been created at Location No. 1 essentially due to the convention that was adopted. That is, high predictor scores correspond to low TPCF scores in the RD rating, whereas at Location No. 2, low predictor scores are associated with low TPCF scores. It may be said that this effect at Location No. 1 has significantly contributed to the low correlation values that were subsequently calculated (see Tables 6 and 7). However, in terms of composite reliability values, the correlations are most correct at Location No. 1. Only at Location No. 2, where as previously noted, there is an inconsistency in the composite reliability values, were the performance estimators, SRE, PRE, and GRE able to correspond to the low TPC levels and, thus, incidentally, produce higher correlations. Therefore, the correlations obtained at Location No. 1 probably reflect the most accurate portrayal of the use of composite reliability values.



The Convention in Other than the RD Rating

Now except for the extreme case mentioned above in the RD rating, al! other composite reliability values in Table F-2 are high. This is in close agreement with the TPCF results (Tables 3 and 4) for the EM, ET, FT, and IC ratings. The distribution of TPC levels for those ratings demonstrate a high concentration of individuals in the highest level of technical proficiency. On the other hand, in the RM, ST, and TM ratings, the distribution of TPC levels is uniform or demonstrate a higher concentration of technicians in the lowest level of technical proficiency, but the composite reliability values are high.

Now it is difficult to generalize as to how the above observations are reflected in the estimators SRE, PRE, GRE, and WRE for association with TPCF results. This depends on the frequency with which the composite reliability values are used (Table E-1). However, it seems true that the greater the frequency of use of the composite reliability value the less agreement of the TPCF with the reliability estimators.

Observations on the Use of Composite Reliability Values

As with any convention one employs, there seems to be a degree of artificality in the composite reliability values, for it would probably be easy to select a convention, by rating perhaps, which would allow the estimators to have a significant degree of promise in all ratings. This, of course, is a most inappropriate means by which to evaluate an estimator. However, results of this research effort seem to indicate that it is not the use of any convention that is posing difficulty, but rather its overuse. As discussed in Appendix E, there were 42 (out of 164) rating-job activity combinations where the convention had to be employed on at least one-third of the technicians in some job activity and rating. Such a high use of a convention can only force the individual performance estimates (in particular the SRE, PRE, and GRE) to more reflect the effect of the convention rather than individual performance.

This observation is further reinforced with the WRE by the frequency with which it must employ the convention. This estimator only employs the convention for the case in which the technician received $\Sigma UE = \Sigma UI = 0$. The case in which the technician did not work at the job activity is ignored by this estimator. Table K-1 reflects the improvement in nonuse of the convention. It is believed that in no small way does considering only the job activities a technician works at and then appropriately weighting those to form a reliability value permit a greater reflection of individual performance and subsequently a greater degree of association with the TPCF.

Conclusions

In conclusion the more promising estimator of electronic maintenance performance as a function of uncommonly effective and uncommonly ineffective performance incidents is a performance estimator similar to the Weighted-Average Reliability Estimate (WRE). This conclusion is made under the assumption that composite reliability values are employed in those job activities in which a man does not work at or received $\Sigma UE = \Sigma UI = 0$. However, even this estimator cannot be recommended for general use within the U. S. Navy at the present time. Only in the EM, ET, FT, and IC ratings is its use presently warranted, while application in the RM, ST, and TM ratings is tenuous. It is suggested that before this estimator is employed in other ratings that the appropriate job task analysis and isolation of job activity factors be completed before a validation effort is attempted in that rating, perhaps similar to the procedures conducted in this report. That it is necessary to perform a validation effort was illustrated in this research effort by the results obtained in the RD rating. It would seem that it is too much to assume that the technique can be automatically applied to other ratings with even similar job activities.



Finally, the results of this research effort do demonstrate that the performance measurement technique being researched does have a large degree of potential for practical application on specific ratings within the U.S. Navy. In particular, the detailed analyses of the Technical Proficiency Checkout Form (TPCF) demonstrated the potential of that instrument to isolate job tasks which may require further training on the part of some technicians. Likewise the Job Performance Questionnaire can also be employed to isolate those individuals of potential detrimental performance levels in terms of excessive uncommonly ineffective incidents of performance. Even though such instruments are completed by the technician's immediate supervisor, as opposed to an unbiased evaluator, those positive qualities of the performance measurement instruments can prove to be most valuable even if one is not interested in deriving an estimate of the probability of effective performance for the technician evaluated. Therefore, short of a procedure which allows for a completely unbiased evaluator, the technique must be given credit for being perhaps the most viable performance measurement technique presently available and of being of greatest potential value to the U.S. Navy.



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APPENDIX A PERFORMANCE EVALUATION INSTRUMENTS



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JOB PERFORMANCE QUESTIONNAIRE

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5.	Capable of knowing relationship of equipme other related equipment with which his rat is concerned.		
6.	Capable of <u>calibrating</u> most of this unit's ment with which his rating is concerned.	equip-	
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DESCRIPTION OF JOB ACTIVITIES

JOB ACTIVITY

DESCRIPTION

- 1. <u>Using Reference Materials</u>—includes the following type of activities:
 - e. use of supporting reference materials
 - b. making out reports
- 2. Equipment Operation--includes the following type of activity:
 - a. operating equipment, electrical and electronics test equipment, and other electronic equipments
- 3. Electronic Circuit Analysis -- includes the following type of activities:
 - understanding the principles of electronic circuitry
 - b. making out failure reports
 - c. keeping records of maintenance usage data
- 4. Personnel Relationships--includes the following type of activity:
 - a. supervising the operation, inspection, and maintenance of electronic equipments
- 5. Electro-safety--includes the following type of activity:
 - using safety precautions on self and equipment
- 6. Instruction--includes the following type of activity:
 - a. teaching others how to inspect, operate, and maintain electronic equipments
- 7. Electro-repair--includes the following type of activity:
 - a. equipment repair in the shop
- 8. Electro-cognition--includes the following type of activities:
 - maintenance and troubleshooting of electronic equipments
 - b. use of electronic maintenance reference materials



MEANINGS OF FLEET ELECTRONICS MAINTENANCE OBJECTIVES

1. Readiness

To maintain efficiently self, subordinate personnel, equipment, and systems in state of readiness consistent with fleet requirements.

2. Performance

To complete any given mission in minimum time with appropriate level of accuracy and reliability.

3. Operation

To obtain optimum system output when equipment is operated, i.e., output characterized by precision and variability appropriate to mission.

4. Safety

To carry out duties with maximum protection for men and equipment consistent with mission.

5. Preparation

To prepare for personnel requirements of present and future equipment, systems, and situations through use of training programs, maintenance of high morale, etc.



APPENDIX B DATA COLLECTION PROCEDURES



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PROCEDURE EMPLOYED IN DATA COLLECTION

The data collection effort originated by requesting of CINCPACFLT and CINCLANTFLT that ships (Destroyer-type vessels) and men (in the electronic maintenance ratings - EM, ET, FT, IC, RD, RM, ST, and TM) be considered for participation in this project. It was requested that initially a project coordinator at an appropriate fleet command be assigned in order to coordinate the efforts of the project researchers from NAVPERSRANDLAB and ship liaison officers. Favorable response from both fleet commands resulted in the assignment of the following project coordinators:

LCDR G. W. Dunne
Flag Secretary
Cruiser-Destroyer Flotilla Nine
FPO San Francisco, CA 96601
(located at the U.S. Naval
Station, San Diego, CA)

RDCS C. J. Masterson COMCRUDESLANT Staff CRUDESLANTFLT FPO New York, N. Y. (located at the U.S. Naval Station, Newport, RI)

The duties and responsibilities of the project coordinator may be outlined as:

- identifying the ships available for participating in the project.
- seeking the assignment of a liaison officer (usually the senior electronics materials officer) aboard each available ship to coordinate the data collection aboard that ship.
- 3. scheduling the meetings between the ship liaison officers and the project researchers.
- 4. making appropriate checks on the progress of the data collection effort aboard each ship.
- 5. collection and reviewing for completeness of the data forms before mailing all forms to NAVPERSRANDLAB.

Once the project researchers had established the ships and liaison officers of those ships that were participating in the project, an initial meeting was arranged between the ship liaison officers and the project researchers. At this meeting the researchers explained the purpose of the project and outlined the duties and responsibilities of a ship liaison officer for completing the data collection effort aboard his ship. It was emphasized that it would be the ship liaison officers responsibility to perform all phases of the data collection effort aboard ship.



Each liaison officer was given a set of instructions (Appendix C) which outlines in detail the steps he was to perform in order to complete the data collection effort aboard his ship. Essentially it was required that the liaison officer select immediate supervisors of men involved in electronic maintenance activities in the EM, ET, FT, IC, RD, RM, ST, and TM ratings who may participate in the project. Each liaison officer was also given a set of Roster Forms (a sample of a Roster Form can be found in Appendix C, page C-8). One Roster Form was completed by the liaison officer for each supervisor participating in the project and it includes a list of the men the supervisor is to evaluate. A technician occurred on this list if that supervisor had been his immediate supervisor for at least the past two months, length of the evaluation period, and the technician at that time was at most a petty officer second class in one of the aforementioned ratings. In order to achieve the most reliable information and not to place too much of a burden on any one supervisor, the supervisor was limited to evaluating no more than seven men - the seven, or less men, he would be most knowledgable about with respect to their on-the-job performance.

At the second and final meeting, the Roster Forms were returned to the project researchers who gave a Xerox copy of these forms to the ship liaison officer and the project coordinator. Using the Roster Forms the project researchers completed packages of Performance Evaluation Forms, one such package for each supervisor. Each package contained a set of Instructions for the Supervisor for Completing Performance Forms (Appendix D), a Xerox copy of the appropriate Roster Form listing the men that the supervisor is to evaluate, and the appropriate number of sets of Performance Evaluation Forms (discussed in the section, Data Collection Instruments), one set for each technician the supervisor was to evaluate.

Upon returning to his ship, the liaison officer held a meeting with the supervisors to explain the purpose of the project and to describe how they were to complete the sets of Performance Evaluation Forms. Once the supervisors had completed their evaluations, the liaison officer submitted all forms to the Administrative Officer aboard his ship who provided the additional demographic information on the men evaluated. After this exercise the ship liaison officer returned all forms to the project coordinator. The project coordinator, once he received the completed forms from all the ships, mailed all forms to MAVPERSRANDLAB.



APPENDIX C

INSTRUCTIONS FOR SHIP LIAISON OFFICER

NOTE: The instructions on pages C-3 thru C-12 (pages -1 thru -5 of the text of the instructions and pages -1 and -1 thru -4 for enclosures (1) and (2) respectively) of this report are paginated for publication. For economy of reproduction, the page numbers were not changed for this report.



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NAVAL PERSONNEL RESEARCH AND DEVELOPMENT LAPORATORY Washington, D. C.

Performance Evaluation Measures
Information to Ship Liaison Officers

Background

The general purpose of the present study is to further investigate an economical and practical method for providing feedback information regarding the readiness of Naval electronically oriented technical personnel (specifically EM, ET, FT, IC, RD, RM, ST, and TM ratings) for completing their assigned mission. Essentially, a result of the study may include the application of a performance evaluative technique which will allow continuous and quantitative answers to questions such as: What is the current level of effectiveness of the maintenance personnel in a given rating, ship, or squadron? How does the maintenance personnel effectiveness level of a given rating, ship, or squadron compare with that of other ratings, ships, or squadrons?

In a previous Office of Naval Research study, a unique performance measurement technique was developed which the Naval Personnel Research and Development Laboratory is now concerned with replicating in order to establish its validity and to determine the practicality of this concept. Basically, eight dimensions of the electronics maintenance ratings were defined: (1) using reference materials, (2) equipment operation, (3) electronic circuit analysis, (4) personnel relationships, (5) electrosafety, (6) instruction, (7) electro-repair, and, (8) electro-cognition. On the basis of a supervisor observing his men performing on the job, he indicates the number of times that he saw unusually effective or unusually ineffective performance demonstrated on each of the eight job dimensions. Pased upon the estimates of unusually effective ineffective behavior, meaningful measures of technician effectiveness can be established. Further, the individual technician effectiveness values can be further treated to form effectiveness indices for ratings, ships. and squadrons.

Because this effort is primarily concerned with replicating a prior study, any information provided by ship personnel will be used only for esearch purposes by the Naval Personnel Research and Development Laboratory and only to serve as a statistical data base for validating the results of that prior study. Furthermore, any comparisons made between ratings, ships, or squadrons will be used only for research purposes.



Suggested Ship Liaison Officer Assistance Plan

The steps outlined below are those as envisioned by the project researchers which you may execute in order to insure an efficient data collection effort aboard your ship.

Completion of the Roster Forms

Either a project researcher or the project coordinator at CREDESLANT has given you a set of Roster Forms (enclosure (1)). To complete the Roster Forms you are to select men involved in electronic maintenance activities in the FM, ET, FT, IC, kD, RM, ST, and TM ratings and immediate supervisors of those men aboard your ship. Generally the men being evaluated will be petty officers second class and below. The supervisors may be either enlisted personnel or officers but must be the immediate supervisors of the men they are to evaluate. They should have known the men they are evaluating for at least the past two months. Not all supervisors of men in the previously mentioned ratings will qualify for participation in this project, however, it is desired that all individuals in the above ratings participate. Once you have completed the Roster Forms, go to each supervisor to participate in this project and have him verify that he is the immediate supervisor of the men that you have selected for him on the respective Roster Form.

Once the Roster Forms have been verified by the supervisors, send a copy of these forms to:

Mr. Bernard Rafacz
Naval Personnel Research
and Development Laboratory
Room 3315, Bldg. 200
Washington Navy Yard
Washington, D. C. 20390

He is your contact man at the Maval Personnel Research and Development Laboratory. If you have any problems or questions concerned with the implementation of the instructions, it may be possible for you co contact him (AUTOVON 288-4457) or the project coordinator.

He will prepare the packages of Performance Evaluation Forms for mailing to you. One such package is a blue folder containing a set of Instructions to the Supervisor on the left inside cover and sets of performance evaluation forms on the right. On the front of each package is a label with the name of the supervisor who is to complete the sets of performance evaluation forms contained therein on the men for whom he is their immediate supervisor.



Briefing Each Supervisor

Once you have received the Packages of Performance Evaluation Forms from the Naval Personnel Research and Development Laboratory, you are now ready to distribute these forms to the supervisors participating in this project. This may be done in one of two ways: 1) calling a meeting of all supervisors so that you may distribute the packages and brief them as to their responsibilities all at one time, or, 2) meet individually with each supervisor and instruct him on a personal basis as to his responsibilities in completing the experimental forms.

It is your choice as to which procedure to follow, however, for ships with only several supervisors participating, the latter method may be the most convenient.

When meeting with the supervisor you should instruct him on the same subject material as you were originally briefed either by the project researcher or the project coordinator at CRUDESLANT. In particular, you should guide the supervisor through his set of Instructions, going over the three SAMPLE pages of performance evaluation forms. Important things to mention to the supervisor about the three forms in each set of Performance Evaluation Forms are outlined below.

1. The Personnel Identification Information Form (PIIF)

The purpose of this form is to provide various background information on the man being evaluated. The supervisors should first of all supply the information at the top of this form. the background information to the right of the cross-hatched area is to be supplied by the supervisor. Please tell him that he must do this to facilitate the job of the Administrative Officer aboard his ship. Because he is normally in everyday contact with the men he is evaluating, it should not be too difficult to obtain this information directly from the man being evaluated. The information ross-hatched area will be supplied by the to the left of th ther aboard your ship and so the supervisor Administrative should not conc. himself with this section of the PIIF. If for some reason he cannot obtain some of the background information from the man being evaluated, please have the Administrative Officer supply this information at a later date. Particular instructions for completing this form are given in the Supervisor's Instruction Package. Please be sure that he understands these instructions and the correct procedure for entering the requested information on the blank PIIF Op-Scan Sheet in each set of experimental forms.



2. The Technical Proficiency Checkout Form (TPCF)

This form is designed to estimate the level of technical complexity at which a man is able to work without direct supervision. The supervisor should provide the information at the top of the TPCF. The eight task descriptions are listed in ascending order of difficulty with task No. 8 identified as the most difficult task. The supervisor is simply to record an "X" in the CHECKED OUT box for each task which he feels the ratee is capable of performing on his cwn without direct supervision." Otherwise he should record an "X" in the NOT CHECKED OUT box. The supervisor must be certain there is an "X" in a box opposite each task description.

3. The Job Performance Questionnails (JPQ) ANSWER SHEET

The JPQ ANSWER SHEET is undoubtedly the most difficult form we are asking the supervisor to complete. However, the researchers have tried to facilitate the understanding of what information is requested by giving explicit samples to each question to be answered. In particular the JPQ ANSWER SHEET records estimates of the total number of uncommonly effective (UE) and uncommonly ineffective (UI) incidents of performance the supervisor has observed over a prior two month period on the man being evaluated. Read page 7 of the Supervisor's Instructions aloud to every supervisor and be sure that he understands what he is expected to record on the JPO ANSWER SHEET. At the top of page 9 is a definition of what is generally meant as an uncommonly effective or uncommonly ineffective per-However, the supervisor should not adhere strictly to this definition. He would best knew for his particular working area what such performances are. He should use the researchers definition only as a guide to conceptualize upon the extremes of performance.

Page 13 of the Supervisor's Instructions is a SAMPLE of a JPQ ANSWER SHEET found in each set of performance evaluation forms. The supervisor must review the table on page 9 together with the sample in order to understand where and how to record his estimates of the total number of UE and UI incidents of performance on the JPQ ANSWER SHEET. Please review this table with him. Also remark on the NOTE at the bottom of page 9.

QUESTION (c) on page 11 of the Supervisor's Instructions must be answered in column (c) for each job activity. Please point this out to the supervisor.

Three remaining questions given on pages 11 and 12 of the Supervisor's Instructions are to be answered in the lower corner provided on the JPQ ANSWER SHEET. Finally the supervisor is to place his social security account number in the space provided on each JPQ ANSWER SHEET he completes.



Instruct the supervisor that he will have approximately one week to complete all forms for the men he is to evaluate. It is important that he realizes that the Op-Scan sheets (the PIIF and JPQ ANSWER SHEET) will be machine processed and therefore it is essential that they remain as much in their original condition as is possible. Once the supervisor completes the forms, he must return them to you in the original folder, as you submitted them to him. You should review the forms as they are submitted and check that all of the information that is requested has been provided. Please be sure that each form of each set contains the supervisors name and the technician's name in the proper locations.

4. Briefing the Administrative Officer

It is now necessary that you submit all packages to the Administrative Officer aboard your ship in order that the Personal Identification Information Form (PIIF) be completed for each man evaluated. Enclosure (2) is a set of Instructions for the Administrative Officer. You will have to request his assistance in order that the information to the left of the cross-hatched area on each PIIF be completed. All of this information may be found in the personnel jacket of the man being evaluated. Complete details for providing this information re contained within the Instructions to the Administrative Officer. To insure that the forms are properly completed, you should review these instructions with him.

Once the Administrative Officer has completed all the PIIFs and returned all packages to you, from each set of Performance Evaluation Forms remove the paper clip. (For Op-Scan purposes it is imperative that the forms are not bent.) All completed packages and any extra forms are to be given to the project coordinator. He will send them back to Mr. Rafacz at the Naval Personnel Research and Development Laboratory.



SAMPLE

ROSTER FORM

Name of Liaison Office	rLT George Merklin	Date2 Jan 72
Ship USS ROAN	Homeport Loc	cation Newport
Immediate Supervisor	Abner Smith	Rating/RateEMC
Names of Men to be Eval Tom Henry Jones	Luated	Rating/Rate EM3
George William Klack		RMSN
Robert Larry Lane		ETN2
eie.		etc.
		
		
		
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NAVAL PERSCNNEL RESEARCH AND DEVELOPMENT LABORATOR? WASHINGTON, D.C.

Instructions for the Administrative Officer Completing the Personnel Identification Information Form (PIIF)

Recently your ship and some supervisors of men in the electronic or related ratings - EM, FT, ET, IC, RD, RM, ST, TM - aboard your ship have been selected to participate in a performance evaluation program being conducted by the Naval Personnel Research and Development Laboratory. The men selected are immediate supervisors of personnel in those ratings. The maison officer aboard your ship in charge of coordinating the activities between the researchers and your ship for this program has given each such supervisor a package of materials containing a Supervisor's Instruction Package and sets of Performance Evaluation Forms; one set for each man the supervisor is evaluating. Every set of Performance Evaluation Forms is made up of the following forms:

- 1. Personnel Identification Information Form (PIIF)
- Technical Proficiency Checkout Form (TPCF)
- 3. Job Performance Questionnaire (JPQ) ANSWER SHEET

The supervisor will complete in part a set of Performance Evaluation Forms for each man he is evaluating. Only the Personnel Identification Information Form (PIIF) found in each set is to be completed by the Administrative Officer. In essence, this form records various background information on the man being evaluated. Once the supervisor has completed supplying the information requested on the set of Performance Evaluation Forms, he will return them to the liaison officer who in turn will give them to you, the Administrative Officer, in order that the PIIF in each set be completed. For each PIIF you receive, the supervisor was to have supplied the information to the right of the cross-hatched area on this form. The first thing you should do is to check over this section of the form and see that all of the following information was provided for the man being evaluated:

- a. Technician's <u>Name</u>
- b. " Social Security Number
- c. "Birthdate
- d. " Months Known by Supervisor
- e. " Months on Current Job Assignment
- f. " Months of Active Duty Service (USN* and USNR**)
- g. "Rating and Paygrade

^{*}Record 99 months for USN if the Technician's Months of USN Active Duty Service equals or exceeds 99 months.

^{**}Record 99 months for USNR if the Technician's Months of USNR Active Duty Service equals or exceeds 99 months.

If some information in this section is omitted, please provide it. Use only a #2 pencil when supplying any information on this form. To aid you in completing the PIIFs, page 4 of these instructions contains a sample of a completed PIIF. From this sample you will see that the following information is obtained from the section of the form to the right of the cross-hatched area:

- 1. The supervisor is EMC Abner Smith located at Newport aboard the USS ROAN and completed this form on 7 January 1972.
- 2. This supervisor was evaluating EM3 Tom Henry Jones whose social security account number is 123-45-6789 and he was born in March, 1948. He was known for 13 months by the supervisor, and had been 15 months on his current job assignment. Jones had served no USNR active duty time but had been on USN active duty for 29 months.

One you have reviewed the sample on page 4 for that section, turn your attention to the section of the PIIF to the left of the cross-hatched area. The following table represents all of the information for this section that is to be provided b_{ν} you from the service record for the man being evaluated.

<u>Item</u>	Location in Se	ervice Record
Final Mark A School B " C "	NAVPERS	601-4
Class Standing $ \begin{cases} A & " \\ B & " \\ C & " \end{cases} $	17 11	
Class Size $ \begin{cases} A & " \\ B & " \\ C & " \end{cases} $	· 11	
Years of Civilian Education (Yrs. Ed.)	NAVPERS	60%-3
GCT Score	11	
ARI " MECH "	11 11	
CLER "	11	

In all cases the most recently completed A, B, or C school information is to be provided. If for some reason a technician did not attend some of the schools, leave that section of the PIIF blank for that school. A Final Mark such as 54.23 is recorded as 54.23 on the PIIF. Final Marks of



the form, 90, 82, 77, etc., (i.e., two digit scores) are recorded as 90.00, 82.00, 77.00, etc., respectively. If some schools gave a Final Mark of "satisfactory" or "unsatisfactory", record 99.99 for a satisfactory Final Mark and 00.01 for an unsatisfactory Final Mark.

Once again turn to page 4 of these instructions which contains the previously mentioned sample of a PIIF and to the section to the left of the cross-hatched area. From that section of the sample form, the following table represents the information the Administrative Officer aboard the USS ROAN provided on EM3 Tom Henry Jones:

	Iten	As recorded in Service Record
Final Mark	A School B " C "	77.86 (did not attend B School) 90
Class Standing	A " B " C "	57th - 26th
Class Size	A " B " C "	122 - 27
Years of Civilian	Education (Yrs. Ed.)	12
GCT Score ARI " MECH " CLER "		50 49 56 47

If you have any questions regarding the type or the recording procedure of any of the desired information, please consult the liaison officer for this project. Once all PIIFs in all sets of Performance Evaluations Forms have been completed by you, return all packages to the liaison officer. Thank you for your cooperation and efforts in completing these forms.



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NAME OF THE STATE	
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APPENDIX D INSTRUCTIONS FOR SUPERVISOR

NOTE: The instructions on pages D-3 thru D-15 (pages -1 thru -13) of this report are paginated for publication. For economy of remoduction, the page numbers were not manged for this report.



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NAVAL PERSONNEL RESEARCH AND DEVELOPMENT LABORATORY WASHINGTON, D.C.

Instructions for the Supervisor for Completing Performance Evaluation Forms

You, as a supervisor, have been selected to participate in a performance evaluation program being conducted by the Naval Personnel Research and Development Laboratory. Performance evaluation in the U.S. Navy has served as a guideline for the optimal positioning of manpower and for feedback on Naval school effectiveness. It is felt that you will be the most qualified to give accurate and meaningful information on the performance of the men you supervise. For that reason it is necessary that you make an earnest effort at providing the information requested of you. Without your support and honest efforts, the time and funds put into this research endeavor will have been negated.

You are asked to complete a set of three performance evaluation forms for each petty officer and designated striker - in the EM, ET, FT, IC, RD, RM, ST, and TM ratings only - over whom you have immediate supervision. The titles and descriptions of the three experimental forms are:

- 1. Personnel Identification Information Form (PIIF) this form is concerned with the background data of the man you are evaluating. You will complete part of it and the Administrative Officer aboard your ship will supply the remaining information.
- 2. <u>Technical Proficiency Checkout Form</u> (TPCF) this form records the level of technical complexity at which a man is able to perform without direct supervision.
- 3. Job Performance Questionnaire (JPQ) ANSWER SHEET this form records your estimates of the number of a man's uncommonly effective and ineffective performances that you have observed during a specified time period.

From your ship's liaison officer, who is coordinating this project, you have received a packet of materials containing these instructions and a number of blank copies of the three forms just described. Additional blank forms may also be obtained from the liaison officer if you need them. Together, these forms - the PIIF, TPCF, and JPQ ANSWER SHEET - constitute a full set of Performance Evaluation Forms. You are to fill out one of each type of form - one complete set of three forms - for each man whose performance you evaluate. You will have approximately one week in which to complete the sets of Performance Evaluation Forms for all the men who have been designated for you to evaluate. Please remember to use only a #2 pencil in filling out the Performance Evaluation Forms.



Turn now to pages 4, 6, and 13 of these instructions. On those pages are samples of the three Performance Evaluation Forms. Notice that each of the three sample forms is preceded by a page of instructions. These particular sample forms were completed for the following typical, naval situation.

EMC Abner Smith is the immediate supervisor of seven men aboard the USS ROAN located at Newport. He had been given seven sets of Performance Evaluation Forms, with a Supervisor's Instruction Package, by the liaison officer aboard the USS ROAN. During the week of 7 January 1972, he completed all Performance Evaluation Forms. One of the men he evaluated was EM3 Tom Henry Jones. The samples on pages 4, 6, and 13 are therefore the PIIF, TPCF, and JPO ANSWER SHEET containing information on EM3 Tom Henry Jones furnished by his supervisor, EMC Abner Smith.

Before starting on the Performance Evaluation Forms, study the samples that are included in these instructions. Please do not let your answers be influenced by the information given on the sample Performance Evaluation Forms. The answers given on the sample forms are meant only to be guides in aiding you to complete your forms on the men you evaluate.

All responses to questions are completely CONFIDENTIAL. Any information you provide vill be used only by the researchers of the Naval Personnel Research and Development Laboratory, and only to serve as a statistical data base for arriving at performance estimates of enlisted naval personnel in general.

Your help and cooperation in participating in this project are greatly appreciated by the project researchers at the Naval Personnel Research and Development Laboratory.



Instructions for Completing the PIIF

The Personnel Identification information Form (FIIF) records various background information on the man you are evaluating. Refer to the sample PIIF on page 4 of these Instructions and the DIRECTIONS given there. Noce that you need furnish only the information requested to the <u>right</u> of the cross-hatched area, namely:

Technician's Name a. b. Social Security Number с. Birthdate 11 Months Known by Supervisor d. 11 Months on Current Job Assignment ρ. Months of Active Duty Service (USN* and USNR**) f. Rating and Paygrada g.

It is understood that you will obtain most of this information directly from the man you are evaluating. It is important that this background information be as accurate as is possible. If you doubt the accuracy of any of the information you have obtained from the technician, or if you are unable to obtain it, please ask the Administrative Officer aboard your ship to provide you with the correct information. In addition, you as the supervisor will also fill out the two upper left lines of the PIIF. This includes your name, rating and rate, date, ship and its current homeport location.

From the sample PIIF on page 4, notice that the following information is obtained:

- 1. The supervisor is EMC Abner Smith located at Newport aboard the USS ROAN and completed this form on 7 January 1972.
- 2. This supervisor was evaluating EM3 Tom Henry Jones whose social security account number is 123-45-6789, and he was born in March, 1948. He was known for 13 months by the supervisor, and had been 15 months on his current job assignment. Jones had served no USNR active duty time but had been on USN active duty for 29 months.



^{*}Record 99 months for USN if the Technician's Months of USN Active Duty Service equals or exceeds 99 months.

^{**}Record 99 months for USNR active time if the Technician's Months of USNR Active Duty Service equals or exceeds 99 months.

DIRECTIONS Use a #2 pencil only. Blacken the answer //PATE	19 ± 2
S A M P L E PERSONNEL IDENTIFICATION INFORMATION FORM NAME OF SUPERVISOR Abner Swith DATE 2 January 22 SHIP OR UNIT USS POAT LOCATION DATE 2 January 22 SHIP OR UNIT USS POAT	HIS SECTION IS NOT TECHNICAL TECHNI

Instructions for Completing the TPCF

A sample of a completed Technical Proficiency Checkout Form is shown on page 6, completed by EMC Abner Smith for EM3 Tom Henry Jones. Eight tasks are listed in descending order of difficulty. It is felt that an individual's overall proficiency is directly related to the highest level of task in the set which he can perform without direct supervision.

Beginning with task description NO. 1. (Capable of employing safety precautions ...), place an "X" in the "CHECKED CUT" box if you feel that the man you are evaluating is capable of performing the task on his own without direct supervision. If you feel he is not able to perform the task on his own, place an "X" in the "NOT CHECKED OUT" box. Complete this form by going through the eight tasks described. Be sure to provide the information at the top of the TPCF form, to include your name, rating and rate, full name of the man evaluated, ship and its current homeport location, and the date. Refer to the sample TPCF given on page 6 of these instructions. You will notice that EMC Abner Smith marked EM3 Tom Henry Jones, as being "CHECKED OUT" on tasks 1 through 6, but felt that he was not able to perform tasks 7 and 8 without direct supervision.



SAMPLE

TECHNICAL PROFICIENCY CHECKOUT FORM

NAM	C OF SUPERVISOR /bner Smith RATING/I	RATE EMC	
FUL.	NAME OF MAN EVALUATED Tom Henry Jones		
SHII	OR UNIT USS ROAN LOCATION Newport	DATE 7 d	anuary 72
	TASK DESCRIPTION	CHECKED OUT	NOT CHECKEI OUT
1.	Capable of <u>employing safety precautions</u> on most of this unit's equipment with which his rating is concerned.	\boxtimes	
2.	Capable of <u>replacing</u> most of this unit's equipment with which his rating is concerned.	\boxtimes	
3	Capable of <u>removing</u> most of this unit's equipment with which his rating is concerned.	\boxtimes	
4.	Capable of <u>following block diagrams</u> for most of this unit's equipment with which his rating is concerned.	\boxtimes	
5.	Capable of knowing relationship of equipment to other related equipment with which his rating is concerned.	\boxtimes	
6.	Capable of <u>calibrating</u> most of this unit's equipment with which his rating is concerned.	\boxtimes	
7.	Capable of trouble-shooting/isolated mal- function(s) in most of this unit's equipment with which his rating is concerned.		\boxtimes
8.	Capable of employing electronic principles involved in maintenance of most of this unit's equipment with which his rating is concerned.		\boxtimes

MAKE CERTAIN THERE IS AN "X" IN A BOX OPPOSITE EACH TASK DESCRIPTION



Instructions for Completing the JPQ ANSWER SHEET

The purpose of the Job Performance Questionnaire (JPQ) ANSWER SHEET is to record, for a given individual, your <u>stimates of the total number</u> of uncommonly effective and uncommonly ineffective incidents of performance you have observed during the <u>last two months</u> on each of <u>sight</u> job activities (see page 8).

In previous efforts at evaluating the performance of an individual, such as the reports of Enlisted Performance Evaluation, you mentally "averaged" your observations of excellent and poor incidents of performance for that man in order to arrive at an overall performance estimate. However, in this project we are asking you to focus your attention on only the extremes of performance — uncommonly effective and uncommonly ineffective performances. In addition, we want you to disregard an individual's overall performance and to record for each of the eight job activities the total number of times you have observed the occurrence of uncommonly effective and uncommonly ineffective incidents of performance.

As a general example, on a particular day a man has demonstrated two incidents of uncommonly effective performance while involved in the job activity Equipment Operation and no incident of uncommonly ineffective behavior for that job activity on that day. Sometime later, on perhaps another day, he has demonstrated one uncommonly effective performance and one uncommonly ineffective performance for the job activity Equipment Operation. If in the past two months no other of his duties involved Equipment Operation, then your estimate of the total number of uncommonly effective and uncommonly ineffective incidents of performance, for that job activity, is three uncommonly effective and one uncommonly ineffective incident of performance.

In order to determine exactly what specific types of activities you should consider in estimating whether the man has shown an "uncommonly effective" or an "uncommonly ineffective" performance, refer to the DESCRIPTION of that JOB ACTIVITY as given in the following list:



DESCRIPTION OF JOB ACTIVITIES

JOB ACTIVITY

DESCRIPTION

- 1. Using Reference Materials -- includes the following type of activities:
 - a. use of supporting reference materials
 - b. making out reports
- 2. Equipment Operation--includes the following type of activity:
 - a. operating equipment, electrical and electronics test equipment, and other electronic equipments
- 3. Electronic Circuit Analysis -- includes the following type of activities:
 - a. understanding the principles of electronic circuitry
 - b. making out failure reports
 - keeping records of maintenance usage data
- 4. Personnel Relationships--includes the following type of activity:
 - a. supervising the operation, inspection, and maintenance of electronic equipments
- 5. Electro-safety--includes the following type of activity:
 - using safety precautions on self and equipment
- 6. <u>Instruction</u>—includes the following type of activity:
 - a. teaching others how to inspect, operate, and maintain electronic equipments
- 7. Electro-repair--includes the following type of activity:
 - a. equipment repair in the shop
- 8. Electro-cognition--includes the following type of activities:
 - maintenance and troubleshooting of electronic equipments
 - b. use of electronic maintenance reference materials



Because you are the most knowledgeable of the area you supervise, you will know what incidents of performance can be labeled as uncommonly effective or uncommonly ineffective. It is your standard that is to be used for this trial. However, to further assist you the researchers have provided the following general definitions.

An uncommonly effective performance in a specific job activity is an impressive and/or decisive incident of performance qualitatively above those usually observed. Likewise, an uncommonly ineffective performance in a specific job activity is an impressive and/or decisive incident of performance qualitatively below those usually observed.

With this in mind, turn to the sample of a JPQ ANSWER SHEET on page 13. Under each JOB ACTIVITY are three columns - (a), (b), and (c). Notice that under "Using Reference Materials", column (a), the estimate of the total number of uncommonly effective (UE) performances observed during the past two months for EM3 Tom Henry Jones was two. That is, the supervisor, EMC Abner Smith, estimated that altogether he observed two impressive incidents of performance for EM3 Tom Henry Jones which were qualitatively above those usually observed. Similarly for column (b), the supervisor's estimate of the total number of uncommonly ineffective (UI) performance is one. Disregard part (c) for the moment.

The estimates recorded on the JPQ ANSWER SHEET will be machine processed and must be accurately recorded. The sample on page 13 together with the following table, illustrates the correct procedure for recording one-digit numbers (numbers 0 through 9), and for recording two-digit numbers (numbers 10 through 99), in various positions on the JPQ ANSWER SHEET. Review the table in conjunction with the page 13 sample.

JOB ACTIVITY	No. of UE	No. of UI
Using Reference Materials	. 2	1
Equipment Operation	10	5
Electronic Circuit Analysis	. 14	12
Personnel Relationships	0	0
Electro-safety	0	2
Instruction		
Electro-repair	2	1
Electro-cognition	1	1

NOTE: If the man has not had an opportunity to perform a particular activity, leave that job activity unmarked (as shown for "Instruction" on the sample JPO ANSWER SHEET).

If he has had an opportunity to perform a particular activity, but has not shown any uncommonly effective or ineffective performances, enter a zero (0) for both UE and UI (as shown for "Personnel Relationships" on the sample JPQ ANSWER SHEET).



While observing a technician perform any of the eight JOB ACTIVITIES, you, as his supervisor, may have had many objectives in mind which you felt he should be striving for. However, for the purposes of this trial, we ask you to limit your attention to only the following fleet electronic maintenance objectives:

MEANINGS OF FLEET ELECTRONICS MAINTENANCE OBJECTIVES

1. Readiness

To maintain efficiently self, subordinate personnel, equipment, and systems in state of readiness consistent with fleet requirements.

2. Performance

To complete any given mission in minimum time with appropriate level of accuracy and reliability.

3. Operation

To obtain optimum system output when equipment is operated, i.e., output characterized by precision and variability appropriate to mission.

4. Safety

To carry out duties with maximum protection for men and equipment consistent with mission.

5. Preparation

To prepare for personnel requirements of present and future equipment, systems, and situations through use of training programs, maintenance of high morale, etc.

* * * * *

After recording the estimates of the number of UE and UI performances for each of the eight JOB ACTIVITIES, complete the JPQ ANSWER SHEET by placing, in column (c) for each job activity, your reply to the following question:



QUESTION (c)

Considering this man's <u>overall</u> performance, it is your opinion that the <u>importance</u> of this job activity, as a factor in determining the overall performance of this man, is best described as being:

- 3. of central and primary importance
- 2. a significant factor, but of secondary importance
- of only moderate importance in estimating overall performance
- O. of little or no importance

On the sample JPQ ANSWER SHEET on page 13, QUESTION (c) was answered as 2, 3, 3, 2, 1, 0, 2, 3, respectively for each of the eight JOB ACTIVITIES. For this particular man, his supervisor felt that the job activity, "Equipment Operation," was of primary importance as a factor in determining the man's overall job performance.

* * * * *

Answer the following questions in the block marked QUESTIONS in the lower corner provided on the JPQ ANSWER SHEET.

QUESTION (d)

You were asked to recall the number of UE and UI performances that you have observed for this man during the past two months. You feel that a reasonable time span for evaluation would be:

- 1. two months
- 2. six weeks
- 3. four weeks
- 4. two weeks
- 5. one week

QUESTION (e)

What degree of confidence have you that your <u>estimated</u> number of UE and UI performances for this man are very close to the <u>actual</u> number of such performances that occurred during this time period?

- My estimates are probably very close to the actual numbers.
- There may have been a few UE or UI performances more or less than my estimates.
- 3. Cannot be too sure about my estimates. It was very difficult to recall UE and UI performances.
- The actual number of UE and UI performances could be very different from my estimates. Recalling these UE and UI incidents is too difficult to have any confidence in such estimates.



QUESTION (f)

Aside from his performance, to what extent are this man's efforts on the job devoted to tasks and activities directly related to his rating? Compare this amount of rating-related activity to the average for men in his rating and paygrade.

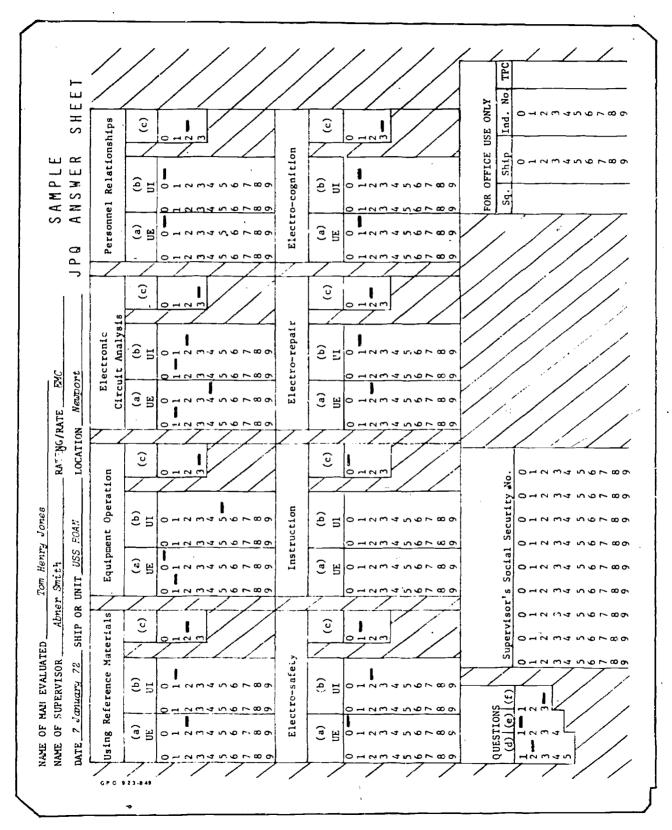
- Involved in <u>definitely more</u> rating-related activities than is usual in his rating and paygrade.
- 2. Involved in about the same rating-related activities as usual in his rating and paygrade.
- 3. Involved in <u>less</u> rating-related activities than is usual in his rating and paygrade.

The above questions were answered as 2, 1, and 3 respectively on the sample of the JPQ ANSWER SHEET of page 13 of these Instructions.

* * * * *

Finally, place your social security account number in the block provided. Also be sure to record your name, rating and rate, ship and its current homeport location, date, and the name of the man you are evaluating at the top of the JPQ ANSWER SHEET.







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APPENDIX E

PROBLEMS IN CALCULATING RELIABILITY RATIOS



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PROBLEMS IN CALCULATING RELIABILITY RATIOS

The purpose of this section is to examine the frequency with which the two cases:

- 1. a technician did not work at a job activity, and
- 2. a technician received $\Sigma UE = 0$ and $\Sigma UI = 0$.

occurred for each rating and job activity across all 21 ships participating in the project. From this one can infer on the extent which any convention for estimating performance in those cases would affect individual SRE, PRE, and GRE values.

Refer to Table E-1 on page E-4. Each square in the table represents the number and proportion of technicians by rating who did not work at a particular job activity or received $\Sigma UE = 0$ and $\Sigma UI = 0$ by their supervisor. Therefore, on job activity Number 1, 25 of the EM's (or 25.8% of the EM's) evaluated either do not work at that job activity (Using Reference Materials) or received $\Sigma UE = 0$ and $\Sigma UI = 0$. This may seem a tolerable level of occurrence of such cases, but when the proportion of such cases exceeds .33, one should begin to consider whether the performance of some individuals is due more to the convention that must be adopted rather than to the individual's own job effectiveness. Of the 64 squares in the table, 42 squares had one-third or more of the men in some rating falling into the two cases for some job activity, 25 squares had one-half or more of the men in some rating falling into the two cases, and most critically, 6 squares had at least 75% of the men in those cases. The RD and RM ratings were particularly notorious for this type of situation occurring. No rating seems to be entirely free of this situation for some job activities. However, some ratings significantly demonstrate this effect for many job activities.



TABLE E-1

NUMBER AND PROPORTION OF TECHNICIANS IN PROBLEM AREAS

b				Rat	ing			
tivity	EM	ET	FT	IC	RD	RM	ST	TM
1	25	38	54	,	50	53	45	6
.	0,258	0,220	0,364	0,155	0,360	0.387	0,296	0,154
2	23	47	49	,	53	42	45	13
٠ [0,237	0,272	0,310	0,121	0,381	0,307	0,296	0,333
3	42	45	75	10	133	ic4	6.0	26
	0,433	0,260	0,487	0,310	0,957	0.45*	0,454	0,667
4	43	#3	93	27	75	77 \	76	10
1	0,443	0,480	0,404	0,466	0,540	0,542	0,500	0,254
5	23	ě3		,	76	50	62	11
	0,237	0,480	0,519	0.121	0,547	0,345	0,408	0,282
6	9.0	116	115	35	100	60	90	20
	0,515	0.671	0,747	0,403	0.710	0,642	0,592	0,513
7	18	48	83	5	137	128	76	. 25
'	0,186	0,277	0,53*	0.086	0.993	0.934	0,500	-0,641
8	35	48	64	12	132	i14	61	. 22
	0,361	0.277	0,416	0,207	0,950	0.832	0,401	0,564
N	umber of	Men Each R	ating				•	·
	97	173	154	58	139	137	152	39



APPENDIX F

THE ADOPTION OF A CONVENTION FOR EACH JOB ACTIVITY AND RATING



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THE ADOPTION OF A CONVENTION FOR EACH JOB ACTIVITY AND RATING

This section discusses the finding that no convention can be adopted per ship that will account for those cases in which a technician either does not work at a particular job activity or received $\Sigma UE = 0$ and $\Sigma UI = 0$ from his supervisor. As an example, observe Table F-1. This table is of the same type as that previously reported on for all 949 men participating in the project (Appendix E, Table E-1). However, Table F-1 is reporting on only one typical ship out of 21 ships in the project. For this ship there were 8 (out of 64) instances where one of those two cases occurred for all men in some job activity and rating. The other ten ships at Location No. 1 demonstrated 16, 7, 8, 10, 13, 12, 5, 5, 5, and 16 (out of 64) instances. The 10 ships at Location No. 2 demonstrated 1, 16, 5, 15, 7, 8, 16, 14, 10, and 15 (out of 64) such instances. Therefore, it is impossible to form an average estimate (or some composite value) per ship for each rating and job activity because, in some ratings and job activities on all ships, there are no technicians who received either ΣUE or ΣUI different from zero.

Derivation of Composite Reliability Values

In order to overcome the aforementioned problem, the convention employed in the report was to develop a composite reliability score to estimate technician performance on those job activities in which the technician received $\Sigma UE = \Sigma UI = 0$ or did not work at that particular job activity.

Let $\Sigma UE(i,j)$ be the sum across all ships at a location of all $\Sigma UE's$ over all men in the ith rating and jth job activity. Similarly the sum of all $\Sigma UI's$ is calculated; denote this sum by $\Sigma UI(i,j)$. The composite reliability score for the ith rating and jth job activity is defined as:

$$R(i,j) = \Sigma UE(i,j)/[\Sigma UE(i,j) + \Sigma UI(i,j)].$$

This particular estimate of job performance provides an "expected" level of effectiveness for a technician in the ith rating and jth job activity (for ships at a particular location). Table F-2 gives the resulting composite reliability values (R(i,j)) for each rating and job activity at each location. For example, from Table F-2 and men at Location No. l (Cruiser-Destroyer Flottilla NINE), R(1,3) is the composite reliability value for EM's on job activity number 3 - Electronic Circuit Analysis - and is given by R(i,3) = .8465. Reiterating, it may be said that for all EM's at Location No. l who have not worked at job activity number 3 or who received $\Sigma UE = 0$ and $\Sigma UI = 0$ for that job activity, their reliability ratio for that job activity is expected to be $r_3 = R(l,3) = .8465$. Similarly this procedure is employed on the other ratings and job activities.



The composite reliability score is an estimate that always can be derived when many ships are involved. However, it does not overcome the implications of the results discussed in Appendix E and their subsequent effect on the estimates SRE, PRE, and GRE.

ь	Rating								
t	EM	ET	FT	IC	RD	RM	ST	TM	
	•	•	5 \	0	0	9	0	. 0	
-	1.667	0.310	0.4173	0.0	0.0	0.946	c.o	0.0	
	6	13	7	o	0	9	٥	1	
	1.5%	5.448	7.583	1.0	<u> </u>	0.990	0.0	3,077	
	5	14	7	3	o	10	. 0	9	
	0.633	0.483	5.583	0.167	n.o	1.000	0.0	0.692	
	6	e	10	3	0	9	o	5	
	1.	1,62:	0.833	1.500	a.c `	0.900	6.5	785	
	6	51	8	0	0	6	6	1	
	1. 13	0.724	6,667	0.1	0.0	0.610	0.0	3.377	
	5	21	9	0	'n	. 10		5	
Ĺ		6.764	11.750	0.0	0.1	1.000	ν, υ	1.3.5	
	5	11	12	1	0	9	v	\$	
_	ورم.د	2,179	1.000	7.167	0.0 -	6.490	0.0	6.692	
	•	16	8	1	o	10	3		
	1.909	1.552	0.667	0.167	n•1 ·	1.000	0.0	ûn2	

Number of Men Each Rating 6 29 12 6 0 10



13

TABLE F-2
COMPOSITE RELIABILITY VALUES

Location No. 1

Job	Rating								
Activity	EM	EΤ	FT	IC	RD	RM	ST	TM	
1	0.8770	0.6831	0.7160	0.7466	0.9257	0.9310	9.8537	0.7333	
2	0.9050	0.7733	0.7802	0.8217	0.9110	0.9497	0.8899	0.8165	
3	0.8465	0.6932	0.7340	0.7981	0.9900	1.0000	0.8652	0.8667	
4	0.8639	0.5987	0.7890	0.7692	0.9390	0.9671	0.8930	0.7733	
5	0.9004	0.7706	0.9107	0.7865	0.9012	0.9712	0.9165	0.7284	
6	0.9333	0.8481	0.8435	0.8039	0.9677	0.9877	0.8571	0.8974	
7	0.8981	0.7872	0.8701	0.8163	1.0000	1.0000	0.9078	0.8269	
8	0.8744	0.7097	0.7771	0.8105	1.0000	9.9949	0.8571	0.7419	

Location No. 2

Job	Rating								
Activity	ÉM	ET	FT	10	RD	RM	ST	TM	
1	0.9101	0.7203	0.8552	0.7879	0.8352	0.#351	0.8547	0.5806	
2	0.8962	0.7110	0.8808	0.8673	0.9017	0.8673	0.8567	0.8n95	
3	0.8342	0.7803	0.8673	0.7719	1.0000	0.8402	0.7794	0.0909	
4	0.9586	0.6729	0.8167	0.9333	0.8214	0.8240	0.7950	0.7317	
5	0.8987	0.8244	0.9032	0.9718	0.8649	0.9667	0.9135	0.6000	
6	0.9530	0.7368	0.9231	0.8095	0.7975	1.0000	0.6759	0.8333	
7	0.9441	0.7866	0.8819	0.8240	0.0000	1.0000	0.8148	0.5714	
8	0.9120	0.7918	0.8827	0.8267	0.7059	1.0000	0.8357	0.4167	



APPENDIX G

A FEW REMARKS ON CURVILINEAR REGRESSION

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A FEW REMARKS ON CURVILINEAR REGRESSION

Let Y be a criterion variable and let X be a predictor variable. If N is the number of observations on each of X and Y, then $Y' = [Y_1, \ldots, Y_N]$ is the row vector of N observations on Y. For a given matrix A, A' will be the transpose of A. In particular one wishes to establish a regression equation for a particular response Y in terms of the variables X, X', X'; i.e. it is desirable to establish which of the three power curves (linear, quadratic, or cubic):

$$Y = \beta_0 + \sum_{i=1}^{p} \beta_i X^i$$
 $p = 1, 2, \underline{or} 3$

best fits the observations obtained on X and Y. The above equations, in terms of the sample observation vectors, can be expressed in matrix notation as:

$$Y = XB + E$$

where Y was defined above. The matrix $X_i = [J, X_1, X_2, X_3]$ where $J' = [1, ..., 1]_{1 \times N}$ and $X_1' = [X_1', ..., X_N]$ for i = 1, 2, 3r 3. There-

fore X_1 is an N x 1 column vector of observations on the predictor variable. $\beta' = [\beta_0, \ldots, \beta_p]$ is the vector of p+1 regression parameters. $E' = [e_1, \ldots, e_N]$ is the vector of errors due to lack of fit in the particular model. One wishes to estimate β such that the error sum of squares is minimized. In particular a <u>least squares</u> estimate β of β is given by

$$\widetilde{\beta} = (X'X)^{-1}X'Y$$

provided the square matrix $X^{1}X$ is nonsingular and the regression problem has been properly expressed. The usual assumption one makes is that E^{1} is distributed with mean $[0, \ldots, 0]_{1\times N}$ and variance-covariance matrix $\sigma^{2}I$

where I is the identity matrix. The term σ^2 is called the common error variance of the observations. The assumption of normality of the error vector is not required in order to obtain the least squares estimates for any of the parameters in the regression equation. Because any assumption of normality for E implies that the observations on X or Y are normally distributed, this report will only be concerned with least squares estimates. One cannot in general discuss normality on X or Y because of some results in this report where it is shown that the distribution of TP scores, SRE, PRE, and GRE are not normally distributed (see Table 17 and page 21). Therefore it is imperative that the reader be aware that while the assumption of E being normally distributed is not required in order to obtain \mathfrak{F} , it is required in order to make tests of hypotheses, as contained in an Analysis of Variance Table. These tests are the usual t- or F-tests and they cannot be applied validly to the sample data collected at either location nor on the combined sample consisting of 949 technicians for the variables SRE, PRE, GRE, and TP score (see, for example, Draper and Smith [7], page 59).



A Least Squares Analysis of the Sample Data

It is possible that a least squares analysis can be attempted independently of the distributional properties of the criterion and predictor variables. For a particular predictor variable (X) and criterion variable (Y), the multiple correlation coefficient (R^2) provides a measure of the proportion of the total variance about the sample mean for the criterion variable explained by a particular regression model. The term R^2 is defined as:

$$R^{2} = \frac{\text{Sum of squares due to regression - Sum of squares due to } \beta_{0}}{\text{Total (corrected) sum of squares}}$$

$$= (\beta_{1}^{2} \chi^{1} Y - \frac{Y^{1} Y}{N}) / (Y^{1} Y - \frac{Y^{1} Y}{N})$$

It should be clear that the larger R^2 is, the better the fitted equation explains variance in the criterion variable. Furthermore, $0 \le R^2 \le 1$, and therefore R^2 = 1 implies a perfect fit. However there are a few problems with this approach (see, for example Draper and Smith [7], page 63). One must weigh the value of R^2 with the least squares estimate (s²) of the common error variance (σ^2) where

$$s^2$$
 = residual mean square
= $(y^*y - \frac{y^*y}{N})$ / (N - p - 1).

Of course, the smaller s^2 is for a particular model under consideration the better the model fits the data. Therefore the approach is to weigh increases in R^2 with decreases in s^2 in order to arrive at the best least squares model for the data.



APPENDIX H

RESULTS OF THE CURVILINEAR REGRESSION ON 949 TECHNICIANS EVALUATED

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THE PREDICTOR VARIABLEIX! IS SEE, AND THE CRITERISM VANIABLEIVE TO SCORE.
```

```
TEST MEANS AND STANDARD DEVIATIONS
                X 0.331
SQUAR 0.214
                                                0.323
                                                 0 288
                x CUFE
                                 0.155
                                                 3 265
                                 5,35,
                                                 4.344
          CORPELATIEN MATRIX
           1 2
                SQUAR
                                 1.000
                                                                0,847
                                                                                0.055
                                 0.953
                                                 1 000
                                                                0,443
                                                                              0,032
                X CUBE
                                 0.887
                                                 0.985
                                                                1.000
                                         $56.0
                                 0.05%
                                                                0.034
                                                                               1,000
          *************************************
          FIRST. SECENC. AND THILD DEGREE PREVNEHIALS.
          HULTIPLE R SQUARE # 0,303
MULTIPLE R = 0.055
          N.D.F.1 = 1
N.D.F.2 =
                               947
           FOR ANALYSIS OF VARIANCE ON R .
                                                          2,650
          BEIA WEIGHTS
            0.055
          CONTRIBUTIONS TO MULTIPLE CORRELATION AND REGRESSION FACTOR LOADINGS. 2ND COLUMN)
         SOURHED BETA WEIGHTS
                                             1.000
         0.003
8 #EIGHTS
            3.397
          INTERCEPT CONSTANT # 5,219
          ************************
         MULTIPLE R SQUARE # 0,007

MULTIPLE R = 0.086

N.D.F.1 = 2

N.D.F.2 = 946

F FOR ANALYSIS OF VARIANCE ON R #
         BETA WEIGHTS

0.264 -0.220

CONTRIBUTIONS TO MULTIPLE CORRELATION
           SOUARED BETA HEIGHTS
0.070 0.048
8 HEIGHTS
        FOR ANALYSIS OF VARIANCE ON R # 14.091
         BETA WEIGHTS
         -1.492 -3.627 2.276
-1.492 -3.627 2.276
CONTRIBUTIONS TO -ULTIPLE CORRELATION
AND RECRESSION FACTOR LOADINGS, 2ND COLUMN)
1 x 0.002 0.205
2 SQUAR 20.316 0.155
3 x CURE 0.077 0.144
                 X CUBE
                                 0.077
         SQUARED BETA WEIGHTS 2,226 13.157 3,
                               5,179
         B HEIGHTS
         13.805 -29.457 20.108
INTERCEPT CONSTANT .
                                        4.869
          ANGVA TABLE FOR POLYNOMI, L.
                                            *****************************
        REDUCTION DLE TO LINEAR FIT, WITH 1 DF = 0,003
RESIDUAL 5.5. 0.997 DF = 947 RESIDUAL M.S. =
F FOR LINEAR FIT = 2.850 /
        REDUCTION DLE TO GENERAL QUADRATIC FIT WITH DF 2, #
REDUCTION M.5. # 0,004
RESIDUAL 5.5. # 0,993 DF 0 946 RESIDUAL M
F FOR QUACHATIC FIT 4 3.537
HEDUCTION DUE TO QUADRATIC TERM ALONE, WITH 1 DF, 0
F FOR QUADRATIC TERM ALONE # 4,215
                                                         946 RESIDUAL H.S. .
                                                                                           0,001
        HEDUCTIAN QUE TO GENERAL CUBIC FIT WITH OF S. .
                                                                             0:043
        HEDUCTION H.S. # 0.014
RESIDUAL S.S. # 0.957 DF #
                                                        945 RESIDUAL M.S. =
                                                                                           0.001
FOR GENERAL CUBIC FIT a 14,091

DUCTION DUE 13 CUBIC TERM ALONE WITH 1 DF a , 0.035

FOR CUBIC TERM ALONE a 34,943
```



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& VARIABLES-PRE AND TH SCHREE 249 DESERVATIONS.
                   PALYMORIAL FITTING FOR
                   THE PREDICTOR VARIABLE(X) IS DRF. AND THE CRITERION VARIABLE(Y) IS TO SCORE;
                  TEST MEANS AND STANDARD DEVIATIONS
1 X 0.558 0.46
2 SQUAR 0.477 0.386
                                                                                                                         0.467
                                   ' CLEE
                     3
                                                                                 0.420
                                                                                5.350
                                                                                                                          2,340
                  COURELATIEN MATRIX
                                                                                                                                                                  0,934
                                   X
SQUAR
                                                                                 1.000
                                                                                                                          0.976
                                                                                                                                                                                                            0.100
                                                                                                                                                                                                            0.063
                                                                                0.976
                                                                                                                         1.000
                                                                                                                                                                  0,959
                                   X CUEE
                                                                0,100
                   4 Y 0.100 0.063 0.036
                                                                                                                                                                                                            1,000
                  HIRST SECUND, AND THIRD DEGREE POLYNOMIALS, TO THE PROPERTY OF THE PROPERTY OF
                  F FOR ANALYSIS OF VARIANCE ON R W
                                                                                                                                                        9,576
                  BETA WEIGHTS
                 BETA METURES
0.100
0.100
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND RECHESSION FACTOR LOADINGS, 2ND COLUMN)
0.010
1.000
                  SOUARED BETA WEIGHTS
                        0.010
                  8 HEIGHTS
0.576
                  INTERCEPT CONSTANT # 5.028
                0,036
                 N.D.F.1 = 2
N.D.F.2 = 946
F FOR ANALYSIS OF VARIANCE ON R =
                                                                                                                                                   17,505
                HETA HEIGHTS

0.817 -0.734

CONTRIBUTIONS TO MULTIPLE CORRELATION

AND REGHESSION FACTOR LOADINGS. 2
                                                                                                                                             ZND COLUMN)
                1 X 0.082
2 SQLAR 0.046
SQUARED BETA HEIGHTS
3.067 0.939
8 HEIGHTS
4.700 -4.431
INTERCEPT CONSTANT *
                                                                                                       4.838
               INTERCEPT CONSTANT # 4.838
#ILLESTIBLE RESURE # 0.036
MULTIPLE R = 0.189
MULTIPLE R = 0.189
MU.D.F.1 # 3
M.D.F.2 # 945
F FOR ANALYSIS OF YARIANCE ON R #
META WEIGHTS
               11.665
                4.341 -3.495 -2.590
INTERCEPT CONSTANT +
                                                                                                      4,841
                ANDVA TABLE FOR POLYNOMIALS
RECUCTION DLE TO LINEAR FIT, WITH 1 DF P 0,010
RESIDUAL S.S. 9,999 DF 947 RESIDUAL M.S.
FOR LINEAR FIT # 9,576
              HEDUCTION DLE TO GENERAL QUADRATIC FIT WITH DF 2, meduction M.S. 0.018
RESIDUAL S.S. 0.964 DF 0.946 RESIDUAL S.F. FOR QUARRATIC FIT = 17.508
REDUCTION LUE TO QUADRATIC TERM ALONE, WITH 1 DF, p. FOR QUARRATIC TERM ALONE = 25.195
                                                                                                                                                   946 RESIDUAL M.S. .
               REDUCTION DUE TO GENERAL CUBIC FIT WITH DF 3, # 4EBUCTION M.S. # 0.012
RESIDUAL S.S. # 0.964 DF # 945 RESID
                                                                                                                                                                                                      6:034
                                                                                                                                                 945 RESIDUAL MIS. .
AR GENERAL CUBIC FIT # 11,865 UCTION EUE TO CUBIC TERM ALBNE WITH 1 DF # , AL CUBIC TERM ALBNE # 0,016
```

H-4

. 0

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POLYNOMIAL FITTING FOR
                                          2 VARIABLES-GRE AND IR SCORE. 949 ONSERVATIONS.
 THE PREDICTOR VARIABLE(x) IS ORE. AND THE CRIPERION VARIABLE(Y) IS THE SCOPE.
   TEST MEANS AND STANDARD DEVIATIONS
1 X 0,927 0.14
2 SQUAR 0,879 0.17
                                                0.145
0.177
0.210
          X CLEE
                              0.839
                              5.330
  CORRELATION FATRIX
                              1.000
                                                0.956
                                                                  0.895
0.996
1,000
                                                                                    0.047
          SQUAR
                              0.956
                                                1.000
                                                                                    0.0AL
          X CUEE
                              0.895
                                                0,986
                                                                                    0.077
                             0.087
                                                0.081
                                                                                    1.000
            *******************
  FIRST, SECCHE, AND THIRD DEGREE POLYNOMIALS.
  MULTIPLE R SQUARE = 0,008
  MULTIPLE R = 0.007
N.O.F.1 = 1
N.O.F.2 = 1
P.O.F.2 = 1
F.FOR ANALYSIS OF VARIANCE ON R W
  BETA WEIGHTS
     0.087
  CONTRIBUTIONS TO MULTIPLE CORRELATION AND RECRESSION FACTOR LOADINGS, 2ND COLUMN)
  SQUARED BETA HEIGHTS
                                           1.900
     0.008
 8 WEIGHTS
  INTERCEPT CONSTANT .
                                       4.053
 #ULTIPLE R SQUARE # 0,008

#ULTIPLE R = 0.987

N.D.F. 1 = 2

N.D.F. 2 = 946
    FOR AVALYSIS OF VARIANCE ON R .
 F FOR ANALYSIS B) YERRALD

BETA WEIGHTS

0.103 -0.017

CONTRIBUTIONS TO HULTIPLE CORRELATION

AND REGRESSION FACTOR LOADINGS. SND COLUMN)

X 0.009 0.998
X 0.009
2 SQLAR #0,001
SQUARED BETA WEIGHTS
D.011 5,000
8 WEIGHTS
1,663
 1.663 -0.225
INTERCEPT CONSTANT .
                                     4,006
 FOR ANALYSIS OF VARIANCE ON R # 2,798

FOR ANALYSIS OF VARIANCE ON R # 2,798

BETA WEIGHT

0.601 -1.542 1.005

CONTRIBUTIONS TO MULTIPLE CORRELATION

AND RECRESSION FACTOR (BADINGS. 2N) COLUMN:

1 X 0.057 0.923

0.867
3 X CUBE 0
SQUARED BETA HEIGHTS
0,438 2,378
                           10.125
                                              0.867
                             0.077
                                              0.816
   0.438 2.378 1,010
HEIGHTS
  10.690 -20.332 11.213
 INTERCEPT CONSTANT .
                                     3.914
  ANSVA TABLE FOR POLYNOMIALS
REDUCTION DIE TO LINEAR FIT, WITH 1 DF = RESIDUAL S.S. = 0.992 DF 4 947 F 1 CR LINEAR FIT 4 7,161
                                                                       0.008
  REDUCTION DUE TO GENERAL QUADRATIC FIT HITH OF 2; .
REDUCTION P.S. = 0.004
RESIDUAL S.S. = 0.972 OF = 946 RESIDUA
F FOR QUACHATIC FIT = 3.589
REDUCTION CUE TO QUADRATIC TERM ALONE, HITH 1 DF,
F FOR QUACRATIC TERM ALONE = 0.024
                                                                                                 0.001
****************************
REDUCTION DUE TO GENERAL CUBIC FIT WITH DE 3. .
REDUCTION M.S. .
                                0.003 T
0.991 DF 4
 RESIDUAL S.S.
                                                         945 RESIDUAL M.S. .
```

POR CUBIC TERM ALONE THE 1 OF ...

H-5

0.001

100.0

```
SUSRE.
                                                                                                                                                                                             949 ORSFRYATIONS.
       THE PREDICTOR VARIABLE(X) IS HER, AND THE CRITERION VARIABLE(Y) IS TO SCORE.
       TEST MEANS AND STANDARD DEVIATIONS
                                                               0.603
                       SGUAR
                                                                                                     0.231
                     X CABE
                                                               0.290
                                                                                                     0.227
                                                              5.350
                                                                                                     2.340
      CORRELATION PATRIX
                                                              1.000
                                                                                                    0.970
                                                                                                                                         0,920
                                                                                                                                                                               0.257
                      SQUAR
                                                               0.970
                                                                                                     1.000
                                                                                                                                                                               0.247
                                                                                                                                         0.986
                     X CLEE
                                                               0.920
                                                              0.257
                                                                                                    0,247
                                                                                                                                         0.229
                                                                                                                                                                               1.000
                                        FIRST, SECENC, AND THIRD DEGREE POLYNOHIALS,
       F1.822512328282828282
      MULTIPLE R SGUARE # 0,066
HULTIPLE R # 0.257
     N.D.F.1 = 1
N.D.F.2 = 947
F FOR ANALYSIS OF VARIANCE ON R 4
    SETA MEIGHTS

0.297

CONTRIBUTIONS 19 MULTIPLE CORRELATION

AND REGRESSION FACTOR LOADINGS, 2ND COLUMN;

1 X 5004ARED BETA HEIGHTS
          0.066
    B WEIGHTS
     INTERCEPT CONSTANT .
                                                                        3,571
   SQUARED BETA WEIGHTS
0.086 0.001
8 WEIGHTS
   3.371 -0.384
INTERCEPT CONSTANT =
                                                                                 3,473
    MULTIPLE R SCUARE 9 0,071
MULTIPLE R = 0.266
N.O.F.1 = 3
N.O.F.2 = 945
FOR ANALYSIS OF VARIANCE ON R =
  HETA WEIGHTS
-0.337 1.972 -1.012
CONTRIBUTIONS TO HULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS. 2
                                                                                                                   2ND COLUMN)
                                              #0.087
0.389
                                                                                                0.965
                       SQLAR
                       X CUBE
                                                          .0.231
                                                                                                0.859
 SOUARED BETA HEIGHTS
0.113 2.472 1,
B HEIGHTS
                                                         1,024
  -3.865 13.953 -10.453
INTERCEPT CONSTANT =
                                                                              4.245
     ANGVA TABLE FOR POLYNGHIALS
    REDUCTION DUE TO LINEAR FIT, WITH 1 DF a
RESIDUAL S.S. = 0.934 DF a 947
FOR LINEAR FIT = 67,017
    REDUCTION DUE TO GENERAL QUADRATIC FIT WITH DF 2.
REDUCTION N.S. = 0.033
RESIDUAL S.S. = 0.934 DF = 946 RES
F FOR QUACRATIC FIT = 33.519
REDUCTION TUE TO QUADRATIC TERM ALONE, WITH 1
F FOR QUADRATIC TERM ALONE = 0.005
  0,055 a representation of the services of the 
0.024'
0.929 DF •
                                                                                                                       945 RESIDUAL M.S. .
                                                                                                                                                                                                        0,001
```

FOR GENERAL CUBIC FIT .

EDUCTION OUE TO CUBIC TERM ALONE WITH & OF # .

4,698

H-6

APPENDIX I

RESULTS OF THE CURVILINEAR REGRESSION ANALYSIS BY RATING



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THE PREDICTOR VARIABLE (X) IS SRE, AND THE CRITERION VARIABLE (Y) IS TP SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                                                                                                                             0.302
                                                                            0.413
                        SOUAR
                                                                             0.260
                       X CURE
                                                                             0.185
                                                                                                                              0.258
                                                                            6.000
                                                                                                                              1.953
    CORRELATION MATRIX
                                                                            1.000
                                                                                                                              0.945
                                                                                                                                                                             0.867
                                                                                                                                                                                                                              0.365
                        SQUAR
                                                                                                                                                                                                                             0.364
0.364
                                                                                                                             1.000
                                                                             0.945
                                                                                                                                                                              0.980
                                                                             0.467
                        X CUBE
                                                                                                                                                                              1.000
                                                                             0.365
                                                                                                                              0.364
                                                                                                                                                                                                                               1.000
    FIRST. SECOND. AND THIRD DEGREE PULYNOMIALS.
     MULTIPLE R SQUARE # 0.133
MULTIPLE R = 0.365
N.D.F.1 = 1
N.D.F.2 = 95
F.FOR ANALYSIS OF VARIANCE DN R =
  BETA WEIGHTS
0,365
  CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS: 2ND COLUMN;
3 1.000
  NO NE NE STATE O SOUARED BETA WEIGHTS
  0.133
B WEIGHTS
   INTERCEPT CONSTANT # 5.027
    MULTIPLE R SQUARE P
NULTIPLE R = 0.369
N.D.F.1 = 2
N.D.F.2 = 94
                                                                                           0.136
N.D.F.2 = 94
F FOR ANALYSIS OF VARIANCE ON R = 7.420
BETA WEIGHTS
0,196 0.179
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS: 2ND COLUMN)
1 X 0.071 0.987
2 SQUAR 0.065 0.985
 POUMPED RETA WEIGHTS
0.038 0.032
R WEIGHTS
1.265 1.246
INTERCEPT CONSTANT
                                                                                                5+148
  MULTIPLE R SOHARE = 0.419
N.O.F.1 w 3
N.O.J. 2 93
                                                                                              0.1/6
  F FOR ANALYSTS OF VARIANCE ON R =
F FOR MINISTER PROPERTY WITH THE PROPERTY OF T
SQUAR 10 X CHRF +0, SQUARED BETA WEIGHTS 14-274
R WEIGHTS -7.246 26.735 -18.221 INTERCEPT CONSTANT =
```

```
ANOVA TARLE FOR POLYNCHIALS

REDUCTION DUF TO LINEAR FIT, WITH 1 DF = 0.133

RESIDUAL S.S. = 0.867 OF = 95 RESIDUAL M.S. = 0.009

F FOR LINEAR FIT = 14.560

REDUCTION DUF TO GENERAL QUADRATIC FIT WITH DF 2, = 0.136

REDUCTION M.S. = 0.008

RESIDUAL S.S. = 0.864 DF = 94 RESIDUAL M.S. = 0.009

F FOR QUADRATIC FIT = 7.420

REDUCTION DUF TO GUADRATIC TERM ALONE, WITH 1 DF = 0.003

F FOR QUADRATIC TERM ALONE = 0.375

REDUCTION DUF TO GENERAL CUBIC FIT WITH DF 3, = 0.176

REDUCTION M.S. D 0.059

RESIDUAL S.S. = 0.4824 DF = 93 RESIDUAL M.S. = 0.009

FOR GENERAL CUBIC FIT = 6.615

DUCTION DUF TO CUBIC TERM ALONE WITH 1 DF = 0.0040

OR CUBIC TERM ALONE = 4.460
```

THE PREDICTOR VARIABLE (X) IS PRE. AND THE CRITERION VARIABLE (Y) IS TO SCORE.

```
TEST HEARS AND STANDARD DEVIATIONS
                                  0.441
                                                  0.370
             SQUAR
             X CURE
                                  0.539
                                                      0.359
                                                      1.953
                                  4,000
    CORRELATION HATRIX
                                  1.000
                                                      0.979
                                                                         0.942
             SQUAR
                                                                                             0.747
                                 0.979
                                                      1.000
                                                                                             0.297
             X CUSE
                                                                          1.000
                                                                                              0.322
                 Y
                                                      0.297
                                                                                             1.000
    FIRST. SECONO. AND THIRD REGREE PULYNOMIALS.
    MULTIPLE R SQUARE 0.001
MULTIPLE R 0.247
N.D.F.1 = 1
N.D.Fs2 = .95
F.FOR ANALYSIS OF VARIANCE ON R =
                                                                    6.163
    RETA WEIGHTS
   CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS, 2ND COLUMN)
1 2 0.061 1.000
    SOUARED BETA WETOHTS
   0.061
A WEIGHTS
   1.303
INTERCEPT CONSTANT
                                         5.112
    MULTIPLE R SQUARE = MULTIPLE R = 0.365 N.D.F.1 = 2 N.O.F.2 = 94
                                       0.133
   F FOR ANALYSTS OF VARIANCE ON R .
BETS WEIGHTS
  BETS WEIGHTS

-1.033 1.367

CONTRIBUTIONS TO MULTIPLE CORRELATION

AND REGRESSION FROTOR LOADINGS. 2ND COLUMN;

X =0.255 0.676

D.388 0.883
  SOUAR 0.388
SOUARED BETA WEIGHTS
  1.066 1.769
B WEIGHTS
  -5:453 7.664
INTERCEPT CONSTANT =
                                           5.516
  MULTIPLE R SQUARE = MULTIPLE R = __0.365
N.D.F.1 = __3
N.D.F.2 = __93
                                      0.133
MULTIPLE 3
N.O.F.1 = 3
N.O.F.2 = 93
F. FOR ANALYSIS OF VARIANCE ON R = 4.774
BETA MEIGHTS
=1.006 1.241 0.041
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS: 2ND COLUMN;
1 K =0.248 0.676
0.368 0.863
 3 X CURF 0.00
SQUARED BETA IGHTS
1.012 1.539 0.002
B REIGHTS
 -5.313 6.647 0.
INTERCEPT CONSTANT =
                              0.224
                                          5.515
  ANOVA TABLE FOR POLYNOHIALS
REDUCTION DUE TO LINEAR FIT, WITH 1 OF = 0.061
RESIDUAL 5.7. = 0.939 OF = 95 RESIDUAL H.S. •
FOR LINEAR FIT = 6.163
  REDUCTION DUE TO GENERAL QUADRATIC FIT WITH OF 2. .
REDUCTION M.S. = 0.067
RESIDUAL S.S. = 0.867 DF =
FOR QUADRATIC FIT = 0.867 DF = 94 RESTOUAL FOR QUADRATIC FIT = 1.238 REOLICTION OUF TO QUADRATIC TERM ALONE. WIT4 1 CF4 = FOR QUADRATIC TERM ALONE = 7.867
                                                 71867
REGUICTION DUE TO GENERAL CUBIC FIT WITH DF 3. =

REGUICTION M.S. = 0.044

RESIONAL 5.S. = 0.767 DF = 93 RESIONAL
                                                               93 RESTOUAL M.S. .
                                                                                                         0.609
```

REPORT OF THE STATE OF THE STAT

THE PREDICTOR VARIABLE(X) IS GRE, AND THE CRITERION VARIABLE(Y) IS TO SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                                                                    0.900
                                                                                                              0.051
                        SQUAR
                         X CUSE
                                                                     6.900
                                                                                                               1.953
      CORRELATION MATRIX
                                                                                                                                                        0.997
0.999
                                                                     1.000
                                                                                                               0.999
                                                                                                                                                                                                  106.0
116.0
156.0
                        SOUÂR
                                                                    0.999
                                                                                                              1.000
                        X CUBE
                                                                                                                                                        1.000
                                 Y
                                                                     0.301
                                                                                                               0.311
                                                                                                                                                                                                   1.000
      FIRST. SECOND. AND THIRD DEGREE PULYNOMIALS.
      MULTIPLE R SQUARE = 0.090
MULTIPLE R = 0.301
N.O.F.1 = 1
N.O.F.2 = 95
F FOR ANALYSIS OF VARIANCE ON R =
                                                                                                                                             9.435
     BETA WETGHTS
            0,301
    CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS. 2ND COLUMN:

0.090 1.090
    SQUARED RETA WEIGHTS
           0.090
    A MEIGHTS
    11.507
INTERCEPT CONSTANT .
                                                                               -5.043
     MULTIPLE R SQUARE = 0.154
MULTIPLE R = 0.393
N.D.F.1 = 2
N.D.F.2 = 94
F FOR ANALYSIS OF VARIANCE ON R =
    BETA WEIGHTS
   BETA MEIGHTS

-5.585 5.891

CONTRIBUTIONS TO MULTIPLE CORRELATION

AND REGRESSION FACTOR LOADINGS, 200 COLUMN;

T x -1.679 0.764

2 SQUAR 1.833 0.793
   SOUAR STOUER SOUARED BETA YEIGHTS 31.196 34.747 B WEIGHTS
   EST.EST PIR.TS
INTERCO TRADASTANT ...
                                                                                      97.486
    HULTIPLE R SQUARE W
                                                                              0.421
   N.D.F.2 = 3
   N+D.F+2 = 93
F FOR ANALYSIS OF VARIANCE ON R =
 F FOR ANALYSIS OF THE STATE OF 
                         $ 39,221
$QUAR -82,947
$ CUBE 44,947
 A CUSE 44
SQUARED RETA WEIGH'S
                                                                                                           0.683
   *********
  8 WFIGHTS
  867,907=======2117.081
  INTERCEPT CONSTANT = =1400.053
    ANOVA TABLE FOR POLYNONIALS
 REDUCTION OUF TO LINEAR FIT. WITH 1 OF # 0.090
RESIDUAL 5.5. = 0.910 OF # 95 RESIDUAL M.S. #
F FOR LINEAR FIT # 9.435
REDUCTION DUF TO GENERAL QUADRATIC FIT WITH DF 2. =

RECHOUS S.S. = 0.077

RECHOUS S.S. = 0.846 DF = 94 RESIDUAL F

FOR QUADRATIC FIT = 8.560

REDUCTION DUF TO GUADRATIC TERM ALONE, WITH 1 DF = F

FOR QUADRATIC TERM ALONE = 7.081
                                                                                                                                         94 RESIDUAL H.S. .
REDUCTION DUE TO GENERAL CUBIC FIT WITH DF 3. ...
REDUCTION N.S. ... 0.074
RESIDUAL 5.5. ... 0.779 DF ... 93 RESID
```

93 RESTOUAL M.S. .

0.067

FOR GENERAL CUBIC FIT = 8.794

ENCENTRAL CUBIC TERM ALONE WITH 1 OF + .

POR CUBIC TERM ALONE = 7.987

0.008

EM RATING

```
THE PREDICTOR VARIABLE(X) IS WEE, AND THE CRITERION VARIABLE(Y) IS TO SCORE,
          0.185
                                                    0.225
          A Y A TRIM
                                  6.000
                                                                    0.918
                 X 1.000
SQUAR 0.963
                                                                                     1.492
n.504
n.481
                                                    0,963
                                                   1.000
                 X CHAE
                                   0.018
                                                                    1.000
                                   0.492
                                                    0.504
                                                                                     1.006
          FIRST, SECONO, AND THIRD REGREE PULYNOMIALS.
          MIN TIPLE R SQUARE = 0.42

MULTIPLE R = 0.492

N.O.F.1 = 1

N.O.F.2 = 95

F FOR ANALYSIS OF VARIANCE ON R =
          AFTE WEIGHTS
          CONTRIBUTIONS TO MULTIPLE CORRELATION
          AND REGRESSION FACTOR LOADINGS: 2ND COLUMN;
1 x 0.242 1.000
SOUARED RETA WEIGHTS
            0.242
          A WEIGHTS
             5,194
          INTERCEPT CONSTANT .
                                          2.639
          MULTIPLE R SOHARE 0.254
MULTIPLE R 0.504
N.O.F.1 = 2
N.O.F.2 = 94
            FOR ANALYSIS OF VARIANCE ON R .
                                                            16.024
          BETA WEIGHTS
0.095 0.411
          CONTRIBUTIONS TO MULTIPLE CORRELATION

AND REGRESSION FACTOR LOADINGS: 2ND CQLUMN;

X 0.047 0.976

> SQUAR 0.207 0.999

SQUARED RETA MEIGHTS
            U.009 0.149
WEIGHTS
          1.012 3.573
[NTERCERT CONSTANT = 3:728
          MULTIPLE R SQUARE # 0.300
MULTIPLE R 0.548
N.O.F.1 3
          N.O.F.Z = 93
F FOR ANALYSIS OF VARIANCE ON R =
                                                              13.273
          BET & WETGHTS -1.475 5.100 -3.210
         -1.475 5.100 -3.210
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS. 2ND COLUMN;
1 1 -0.726 0.899
2 SQUAR 2.568 0.920
3 Y CURF -1.543 0.878
SOUARED RETA WEIGHTS
2 176 26 412 10 207
          2.176 26.012 10.307
B WF1GHT5
          -15.571 44.319 -26.561
INTERCEPT CONSTANT -
           ANDVA TABLE FOR POLYNOMIALS
          PERDUCTION DUE TO LINEAR FIT. HITH 1 DF # 0.242

PERIDUAL S.R. = 0.758 DF = 95 RESIDUAL N.S. = FOR LINEAR FIT = 30.345
           PERUCTION DUF TO GENERAL QUADRATIC FIT WITH DF 2. .
         REDUCTION H.S. = 0.127

RESTOUAL 5.5. = 0.746 DF = 94 RESTOUAL
F FOR QUADRATIC FIT = 16:024
REDUCTION OUF TO CUADRATIC TERM ALONE, WITH 1 CF. =
F FOR QUADRATIC TERM ALONE = 1.532
                                                               94 RESIDUAL H.S. 4
         REDITCTION DUE TO BENERAL CUBIC FIT WITH DE 31 .
         RESIDUAL 5.5. = 0.100
RESIDUAL 5.5. = 0.700 OF =
                                                              93 RESTOUAL H.g. =
```

THE PREDICTOR VARIABLEIX) 13 SRE. AND THE CRITERION VARIABLEIY) 15 TF SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                                                                              0.168
                        SQUAR
                                                                                                                                 0.149
                        X CURE
                                                                              0.040
  CORRELATION WATREX
                                                                              1.000
                                                                                                                                  0.909
                                                                                                                                                                                    0.778
                                                                                                                                                                                                                                       0.318
                        SQUAR
                                                                              0.909
                                                                                                                                 1.000
                                                                                                                                                                                    0.963
                                                                                                                                                                                                                                       0.205
                         X CUBE
                                                                                                                                                                                                                                       1.126
                                                                                                                                                                                    1.000
   4 Y 0,318 0.205 0.128
                                                                                                                                                                                                                                       1.000
   FIRST, SECOND, AND THIRD REGREE PULYNOHIALS.
  MULTIPLE R SQUARE # 0.121
MULTIPLE R # 0.318
N.O.F.1 * 1
N.O.F.2 * 171
F FOR ANALYSIS OF VARIANCE ON R #
   BETÄ WEIGHTS
  0.316
CONVERNATIONS TO MULTIPLE CORRELATION
               NO REGRESSION FACTOR LOADINGS: 2ND COLUMN;
                                                                                0,101
   SQUARED RETA WEIGHTS
  0.101
8 WEIGHTS
   INTE CEPT CONSTANT . 6.057
  MULTIPLE R SOMARE 0.14

MULTIPLE R SOMARE 0.14

MULTIPLE R SOMARE 1.70

MULTIPLE R 1.70

M.O.F.1 2

M.O.F.2 3

TO ANALYSIS OF MACTANOR
                                                                                                0.142
  N.O.F.Z = 170
F FOR ANALYSIS OF VARIANCE ON R =
                                                                                                                                                                   14.021
   BETA WETGHTS
  BETA WEIGHTS

0.799 -0.4=9

CONTRIBUTIONS TO MULTIPLE CORRELATION

AND REGRESSION FACTOR LOADINGS: ZND COLUMN:

| x 0.241 0.844
| 2 SQUAR -0.099 0.545

SOUARED RETA VEIGHTS

0.576 0.235

B WFIGHTS

4 A00 -5-703
           0.759
   6.498 -5.797
INTERCEPT CONSTANT -
                                                                                                            5.835
    *****************
  MULTIPLE R SCHARE = 0.142
MULTIPLE R = 0.377
N.O.F.1 = 3
N.O.F.2 = 169
F FOR ANALYSIS OF VARIANCE ON R =
NOVE TO THE CONTROL OF THE COLUMN OF TH
  T CURE 0.019
SOUARED BETA WEIGHTS
0.714 0.944 0.023
B WEIGHTS
                                                                             2.129
    7.235 -8.477 2.
INTERCEPT CONSTANT -
                                                                                                            5,817
```

```
ANOVA TABLE FOR POLYNCHIALS

REDUCTION OUF TO LINEAR FIT. WITH 1 DF = 0.101

REDUCTION OUF TO GENERAL GUADRATIC FIT LITH OF 2. = 0.005

F FOR LINEAR FIT = 19.191

REDUCTION W.S. = 0.071

RESIDUAL S.S. = 0.088 OF = 170 RESIDUAL W.S. = 0.005

F FOR GUADRATIC FIT = 14.021

REDUCTION DUE TO GUADRATIC TERM ALONE, WITH 1 OF = 0.041

F FOR GUADRATIC TERM ALONE = 8.058

REDUCTION DUE TO GENERAL CUBIC FIT WITH OF 3. = 0.148

REDUCTION M.S. = 0.047

RESIDUAL S.S. = 0.855 OF = 169 RESIDUAL W.S. = 0.005

FOR GENERAL CUBIC FIT = 9.323

FOR GENERAL CUBIC FIT = 9.323

FOR GENERAL CUBIC FIT = 9.323

FOR GENERAL CUBIC FIT = 0.078
```

```
TEST MEANS AND STANDARD DEVIATIONS
                                                                        0.404
                                                                                                               0.352
                                SQUAR
                                                                        0.287
0.213
6.514
                                                                                                                0.288
                               X CUBE
                                                                                                               0.248
                COPRELATION MATRIX
                               X 1.000
SQUAR 0.967
                                                                                                               0.967
                                                                                                                                                    0.904
                                                                                                                                                                                          0.366
0.361
0.343
                                                                                                                1.000
                                                                                                                                                     1.000
                                X CURE
                                                                        0.904
                                                                                                               0.982
                  A Y 0.366 0.351 0.343
                                                                                                                                                                                           1.000
                FIRST, SECOND, AND THIRD DECREE PULYNOMIALS.
               MULTIPLE R SQUARE 1 134
MULTIPLE R 0.3A'
N.O.F.1 1 1
N.O.F.2 171
F FOR ANALYSIS OF VARIANCE ON R .
                BETS WEIGHTS
               O.366
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS. 2ND COLUMN)
1 x 0.174 1.000
                SOUARED RETA WEIGHTS
                0.134
R WEIGHTS
                1.854
INTERCEPT CONSTANT = 5.765
               MULTIPLE R SQUARE = 0.135
MULTIPLE R = 0.367
N.O.F.1 = 2
N.O.F.2 = 170
F FOR ANALYSIS OF VARIANCE ON R =
              F FOR ANALYSIS OF VARIANCE ON R = 13.222

BETA WEIGHTS
0.260 0.109

CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS: 2ND COLUMN)
1 x 0.095 0.997
2 SQUAR 0.039 0.984

SQUARED RETA WEIGHTS
0.069 0.012

R WEIGHTS
1.319 0.477
                                                                                                                                       13.222
               1.319 0.477
INTERCEPT CONSTANT = 5.787
                  ****************
             MULTIPLE R SONARE = N.O.F.1 = 3
                                                                                   0.135
                  .O.F.2 = 169
FOR ANALYSIS OF VARIANCE ON R =
                                                                                                                                     8.749
              BETA WEIGHTS

0.364 -0.137 0.148

CONTRIBUTIONS TO MULTIPLE CORRELATION
                        AND REGRESSION FACTOR LOADINGS: 2ND COLUMN)

X 0.133 0.997

R SQUAR -0.049 0.983
                                    X 0.133
59UAR -0.049
X CURF
                                                                          0.051
                                                                                                             0.935
             SOUTRED BETA WEIGHTS
                    0.133 0.019 0.022
WEIGHTS
                     1.845
              1.845 -0.847 1.068
INTERCEPT CONSTANT -
                                                                                          5:783
                 ANOVA TABLE FOR POLYNCHIALS
             REDUCTION DUF TO LINEAR FIT, WITH 1 DF = 0.134
RESIDUAL 5.5. = 0.866 DF = 171 RESIDUAL M.S. = 0.003
'F FOR LINEAR FIT = 26.423
                 REDUCTION DUF TO GENERAL QUADRATIC FIT WITH DF 2. m
             REDUCTION H.4. = 0.067
HESTOUAL S.S. = 0.465 DF = 170 RESTOUAL
F FOR QUADRATIC FIT = 13-222
REDUCTION OUF TO GUARRATIC TERM ALONE = 0.152
F FOR QUADRATIC TERM ALONE = 0.152
                                                                                                                                     170 RESTOUAL M.S. .
                                                                                                                                                                                                  0.001
              REDUCTION OUR TO GENERAL CUBIC FIT WITH DF 3.
             REDUCTION H.S. . RESIDUAL 5.5. .
                                                                                   0.045
                                                                                    0.865 DF .
                                                                                                                                    169 RESTOUAL M.S. .
                                                                                                                                                                                                                        0.005
ERICIT GENERAL CURIC FIT = 8.769

CONTROL OF THE CURIC TERM ALONE WITH 1 OF T . 0.000

CONTROL OF THE PROPERTY OF THE CONTROL OF THE CONTROL
```

POLYMONIAL FITTING FOR

X SQUAR

X CURE

SQUAR

X CUBE

RETA WEIGHTS

0.258 R WEIGHTS

MILTIPLE R SQUARE = 0.508 N.O.F.1 = 1 171 F.F.2 = 171

SOURED BETA WEIGHTS

CORRELATION MATRIX

TEST MEANS AND STANDARD DEVIATIONS 0.796

N.O.F.2 = 171 F FOR ANALYSIS OF VARIANCE ON R =

0.740 6.514

1.000

0.960

0.493

0.508 FIRST. SECOND. AND THIRD DEGREE POLYNOHIALS.

0.458

0.408.
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS. 2ND COLUMN)

♦

1-9

ET RATING

```
169 RESTOUAL H.S. .
F FOR GENERAL CHAIC FIT # 20.966
REDUCTION OUF TO CURIC TERM ALONE WITH 1 DF # .
F FOR CURIC TERM ALONE # 1.274
```

2 VARIABLES-BRE AND TO SCORE. 173 OBSERIATIONS.

0.412

THE PREDICTOR VARIABLEIX) IS ORE, AND THE CRITERION VARIABLE (Y) IS TO SCORE,

0.893 0.984 1.000 0.495

59.470

0.222 0.245

0.960 1.000

0.512

1.000

THE PREDICTOR VARIABLE(X) IS WRE, AND THE CRITERION VARIABLE(Y) IS TO SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                                                                                                    0.372
                                                                                                                                                                             0.210
                               SQUAR
                               X CHAE
                                                                                                         6.514
                                                                                                                                                                              1.784
   CORRELATION MATRIX
                                                                                                                                                                                                                                                 0.932
0.987
                                                                                                                                                                                                                                                                                                                    0.445
0.413
0.390
                               RAUGE
                                                                                                                                                                              0.977
                                                                                                                                                                              1.000
                                                                                                          0.932
                                * CUBE
                                                                                                                                                                                                                                                  1.000
     FIRST. SECOND. AND THIRD REGREE PULYNUMIALS.
  FIRST SECOND. AND THIRD AFFREE PARTIES AND THIRD AFFRE PARTIES AND THE PARTIES AND THE PARTI
   BETS WEIGHTS
  BFTS WEIGHTS

0.445

CONTRIBUTIONS TO MULTIPLE CORRELATION

AND PEGRESSION FACTOR LUADINGS. 2ND COLUMN)

7 0.198 1.000
              0.198
   R WEIGHTS
3.789
INTERCEPT CONSTANT =
                                                                                                                                       4,344
           MULTIPLE R SOHARE = 0.439
MULTIPLE R = 0.457
N.0.F.1 = 2
N.0.F.2 = 170
F FOR ANALYSIS OF VARIANCE ON R =
    BETA WETGHTS
   0.913 -0.479
CONTRIBUTIONS TO MULTIPLE CORRELATION
                   AND RESERVEN FACTOR LOADINGS: 2ND COLUMN)

1 X 0.406 0.914

2 SOUAR -0.198 0.904
    SQUARED RETA METCHTS
   0.229
B WEIGHTS
  R WEIGHTS
7,769 =3,862
INTERCEPT CONSTANT = $2,478
ETHERSTELLE STANFORD TO THE TOTAL TO THE TOTAL TOTA
   MULTIPLE R SQUARE = MULTIPLE R = 0.458
N.O.F.1 = 3
N.O.F.2 = 169
1.651 1.875 0.285
8 WEIGHTS
     10.937 -10.877 4.602
INTERCEPT CONSTANT #
                                                                                                                                             3.103
     ANDVA TABLE FOR POLYNCHIALS
    MEDUCTION DUE TO LINEAR FIT. WITH 1 OF # 0.198

DESTRUCTED 171 RESIDUAL M.S.

F FOR LINEAR FIT # 42.260
         REDUCTION DUE TO GENERAL QUADRATIC FIT WITH OF 2. .
   REDUCTION N.S. = 0.104
RESTOUAL S.S. = 0.791 OF = 170 RESTOUAL
F FOR QUADRATIC FIT = 22.416
REDUCTION DUF TO QUADRATIC TERM ALONE. WITH 1 OF. =
F FOR QUADRATIC TERM ALONE = 2.261
                                                                                                                                                                                                                         170 RESTOUAL M.S. #
```

REDUCTION DUE TO GENERAL CUBIC FIT WITH DF 3. #
REDUCTION N.S. = 0.070
RESIDUAL S.S. = 0.790 DF = 149 RESI

RIC FOR GENERAL CURIC FIT = 14.942 "
DUCTION DUE TO CURIC TERM ALONE WITH 1 OF = 4

0.011

149 RESTOUAL M.S. .

I-10

THE PREDICTOR VARIABLE(X) IS SHE, AND THE CRITERION VARIABLE(Y) IS TO SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                                    0.103
0.177
0.124
            SQUAR
                                                           0.255
            X CURE
                                    6.240
  CORRELATION MATRIX
                                    1.000
                                                           0.944
                                                                                 0.981
                                                                                                        n.346
            SQUÂR
                                                           1.000
                                    0.044
            X CUBE
                                    0.871
                                                                                  1.000
                                                                                                        n.294
                                    0.346
                                                           0.322
                                                                                 0.294
                                                                                                        1-006
  FIRST. SECOND. AND THIRD REGREE PULYNOHIALS.
  ********************
  MULTIPLE R SOHARE = 0.119
MULTIPLE P = 0.34A
N.O.F.1 = 1
N.O.F.2 = 152
F FOR ANALYSIS OF VARIANCE ON R =
  BETA WEIGHTS
 BET'S MELUTIONS

0.346

CONTRIBUTIONS TO MULTIPLE CORRELATION

AND REGRESSION FACTOR LUMBINGS. 2ND COLUMN)

1 X 0.110 1.000
 0.119
R WFIGHTS
2.410
INTERCEPT CONSTANT =
                                              5.509
 0.120
 N.D.F.2 = 15t
F FOR ANALYSTS OF VARIANCE ON R =
F FOR ANALYSTS OF BETA WEIGHTS

0.382 **-0.078

CONTRIBUTIONS TO MULTIPLE CORRELATION

AND REGRESSION FACTOR LOADINGS: 2NO COLUMN;

1 0.132 0.992

COUAR -0.012 0.992
 SOUARED BETA VEIGHTS
0.146 0.001
R WFIGHTS
 INTERCEPT CONSTANT .
                                           5.457
HULTIPLE R SQUARE = 0.120

MULTIPLE R = 0.346

N.D.F-1 = 3

N.D.F-2 = 150

F FOR ANALYSIS OF VARIANCE ON R =
F FOR ANALYSTS 0.

BFTA WEIGHTS
0.339 0.043 -0.083
CONTRIGUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS: 2ND COLUMN;
1 X 0.117 0.999
0.027 0.932
0.050
3 X CURF -0.024
SOUARED RETA WEIGHTS
0.115 0.007 0.007

0.115 0.007 0.007

R WEIGHTS

2.364 0.665 -0.730

INTERCEPT CONSTANT =
```

```
ANOVA TARLE FOR POLYNOMIALS

PERIUCTION DUE TO LINEAR FIT, WITH 1 OF = 0.119

RESIDUAL 5.5. 0 0.881 OF = 152 RESIDUAL M.4. 0 0.006

F FOR LINEAR FIT = 20.623

PERIUCTION DUE TO GENERAL QUADRATIC FIT WITH DF 2. = 0.120

REDUCTION W.5. = 0.060

PESTIDUAL 5.5. 0 0.880 DF = 151 RESIDUAL M.5. = 0.006

F FOR QUADRATIC FIT = 10.289

REDUCTION DUE TO CUADRATIC TERM ALONE. WITH 1 FF. = 0.000

F FOR QUADRATIC TERM ALONE = 0.027

REDUCTION FOUF TO GENERAL CUBIC FIT WITH DF 3. = 0.120

REDUCTION M.5. = 0.060

PESTIDUAL 5.5. = 0.060

OR GENERAL CUBIC FIT = 0.006

OR GENERAL CUBIC FIT = 0.000
```

THE PREDICTOR VARIABLE(X) IS PRE- AND THE CRITCHION VARIABLE(Y) IS TO SCORE.

```
TEST HEARS AND STANDARD DEVIATIONS
                                      0.568
                                                               0.393
           SQUÂR
                                                               0.347
           X CURE
                                      0.410
                                      6.240
 CORRELATION MATRIX
                                                               0.976
                                                                                        0.931
                                                                                                                 0.264
0.296
                                      1.000
           SQUÂR
                                      0.976
                                                               1.000
                                      0.931
                                                               0.987
                                                                                        1.000
           X CURE
                                                                                                                 0.112
                                      0.264
                                                               0.296
                                                                                         0.312
                                                                                                                  1.000
 FIRST. SECOND. AND THIRD REGREE PULYNOMIALS.
HILTIPLE R SOMERE = 0.070

HULTIPLE R = 0.264

N.O.F.1 = 1

N.O.F.2 = 152

F. FOR ANALYSIS OF VARIANCE ON R = 11.394

RETA WEIGHTS
0.264

CONTRIBUTIONS TO PULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS, .2ND COLUMN)
1

X 0.070 1.000
 ************************
 SOUTRED RETA WETCHTS
0.070

R WETCHTS
1.372

INTERCEPT CONSTANT = 5
                                                5.460
                                0.100
0.317
 MULTIPLE R SOHARE =
MULTIPLE R SOUTHER = Using Williple R = 0.317
N.O.F.1 = 2
N.O.F.2 = 151
F FOR ANALYSIS OF VARIANCE ON R =
FEOR ANALYSIS OF VARIANCE ON R = 8.408
RETA MEIGHTS
=0.924 0.877
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS; 2ND COLUMN;
1 X =0.138 0.834
2 SOUAR 0.239 0.934
SOUARED RETA WFIGHTS
0.275 0.652
B WFIGHTS
=2.724 4.520
INTERCEPT CONSTANT = 5.635
 8 WF 10m15

-2.724 4.520

INTERCEPT CONSTANT = 5.635
 MULTIPLE R SOWARE = 10,104
MULTIPLE R = 0.322
N.O.F.1 = 3
N.O.F.2 = 150
F FOR ANALYSIS OF VARIANCE ON R =
                                                                                  5.784
F FOR ANALTSIS OF SETA HEIGHTS 0.394 -1.375 1.303 CONTRIBUTIONS TO MULTIPLE CORRELATION AND REGRESSION FACTOR LOADINGS: 2NG COLUMN; 7 X 0.104 0.820 A.018
               50UAR
                                      -0.407
                                                               0.910
               X CURE
                                                               1.968
                                       0.406
 SOUARED BETA WEIGHTS
SOUMHED HETA WEIGHTS
0.155 1.492 1.497
B WEIGHTS
2.048 -7.499 7.657
INTERCEPT CONSTANT =
                                                   5.606
```

```
ANOVA TABLE FOR POLYNCHIALS

REDUCTION DUE TO LINEAR FIT, WITH 1 DF = 0.070
RESIDUAL S.S. = 0.030 DF = 152 RESIDUAL M.S. = 0.006
F FOR LINEAR FIT = 11.394

REDUCTION OUF TO GENERAL QUADRATIC FIT WITH DF 2. = 0.100
REDUCTION M.S. = 0.050
PESIDUAL S.S. = 0.000 DF = 151 RESIDUAL M.S. = 0.006
F FOR QUADRATIC FIT = 8.408
REDUCTION OUE TO GUADRATIC TERM ALONE, WITH 1 DF = 0.030
F FOR QUADRATIC TERM ALONE = 5.114

REDUCTION M.S. = 0.035
RESIDUAL S.S. = 0.096 DF = 150 RESIDUAL M.S. = 0.006

FOR GENERAL CUBIC FIT = 5.784
IFOUCTION OUF TO CURIC FIT ALONE WITH 1 DF = 0.007
```

```
2 VARIADLES-GRE AND IP SCORE- 154 CHSERVATIONS.
POLYNOMIAL FITTING FOR
                                                                                                                       FT DATING
THE PREDICTOR VARIABLEIX IS GRE, AND THE CRITERION VARIABLEIY) IS TRESCORE.
TEST MEANS AND STANDARD DEVIATIONS
                       0.932
0.873
                                       0.071
       SQUAR
       X CURE
                        0.622
                                       2.042
                        6.240
CORRELATION MATRIX
                                       n 2997
                                                      0.991
       SQUAR
                        1.000
                                                                      0.374
                                       1.000
                                                                      0.380
                        0.997
       3 RUBE
                        0.491
                                                       1.000
                                                                      0.342
                        0.374
                                       0.384
                                                       0.382
                                                                      1.000
FIRST. SECOND. AND THIRD NEGREE PULYNOHIALS.
MULTIPLE R SOHARE # 0.140
MULTIPLE R # 0.374 °
N.D.F.1 # 1
N.D.F.2 # 152
F FOR ANALYSIS OF VARIANCE DN R #
BETA WEIGHTS
   0.374
SOUARED RETA WEIGHTS
0.140
R WEIGHTS
10.418
INTERCEPT CONSTANT # #3.840
 MULTIPLE R SQUARE = MULTIPLE R = 0.385
                             0.148
0.858 1.699

##FIGHTS
-26.793 21.215
INTERCEPT CONSTANT =
                               12.681
MULTIPLE R SQUARE = 0.105

MULTIPLE R = 0.406

N.O.F.1 = 3

N.O.F.2 = 150

F FOR ANALYSIS OF VARIANCE ON R =
3 X CURE -10,607
SQUARED RETA WEIGHTS
615,3702789,744.769,824
B METGHTS
717.524 R59.498-334.275
INTERCEPT CONSTANT =
                             199.144
  ANOVA TARLE FOR POLYNOMIALS
  PENUCTION DIF TO LINEAR FIT. HITH 1 OF THE RESIDUAL S.S. B 0.860 DF THE 152
                                                  152 RESTOUAL M.S.
RESIDUAL S.S. . 0.860
F FOR LINEAR FIT . 24.721
                            0.860
REDUCTION DUE TO GENERAL QUADRATIC FIT WITH DF 2. =

REDUCTION M.S. = 0.074

PESIDUAL S.S. = 0.852 DF = 151 RESIDUAL M

F FOR QUADRATIC FIT = 13.156

REDUCTION DUE TO GUADRATIC TERM ALONE: WITH 1 DF: = -

F FOR QUADRATIC TERM ALONE = 1.508
                                                  151 RESTUUAL M.S. .
REDUCTION DUE TO GENERAL CUBIC FIT WITH DF 3. a REDUCTION M.S. = 0.055
                                                                                                                          I-13
  RESTOUAL S.S. .
                             0.835 DF .
                                                  150 " RESIDUAL M.S.
```

0.006

FOR GENERAL CURIC FIT = 9.871
DISCRICTION DUF TO CURIC TERM ALONE WITH 1 DF = ,

THE PREDICTOR VARIABLE(X) IS MRE. AND THE CRITERION VARIABLE(Y) IS TO SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                                        0.A20
0.420
                                                                  0.190
           SQUAR
                                        0.100
            X CURE
CORRELATION MATRIX
                                                                                             0.919
0.984
1.000
                                                                                                                       0.430
            SQUAR
                                        0.970
                                                                   0.986
            X CUBE
                                        0.419
                                                                  0.388
                                                                                                                        1.100
  FIRST. SECOND, AND THIRD REGREE PULYNOMIALS.
     MULTIPLE R = 0.430

Natiple R = 0.430

Natiple R = 152

FEOR ANALYSIS OF VARIANCE ON R =
F FOR ANALTSIS OF THE RETA MEIGHTS
0.430
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS. 2ND COLUMN)
0.185
1.000
 0.185
B WEIGHTS
                                                     3:366
 INTERCEPT CONSTANT .
MULTIPLE R SOURRE = 0.200

MULTIPLE R = 0.448

N.O.F.1 = 2

N.O.F.2 = 151

F. FOR ANALYSIS OF VARIANCE ON R = 18.907

BETA WEIGHTS

0.923 -0.508

CONTRIBUTIONS TO MULTIPLE CORRELATION

AND REGRESSION FACTOR LGADINGS+ 2ND COLUMNS

1 0.397 0.961

SOUARED RETA WEIGHTS

SOUARED RETA WEIGHTS
SOUARED RETA MEIGHTS:
0.452 0.758
R WFIGHTS
9.943 -4.432
 MULTIPLE R SQUARE = 0.264

MULTIPLE R = 0.451

N.D.F-1 = 3

N.D.F-2 = 150

F FOR ANALYSIS OF VARIANCE ON R =
BETA MEIGHTS

0,328 0.004 -0.928

CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS 2ND COLUMN)

1 X 0.141 0.953

A.859
 3quar 0,382
3 % CUBE ~0,319
SOUAREO BETA WEIGHTS
0,107 0,949
                                                                   0,859
                                                                   n.762
 B WEIGHTS
3.531 9.745 #8.994
INTERCEPT CONSTANT #
                                                       2.811
```

```
ANOVA TABLE FOR POLYNCMIALS

REDUCTION DUF TO LINEAR FIT, WITH 1 OF = 0.185

PERIDUAL S.S. = 0.815 OF = 152 RESIDUAL M.S. = 0.005

F FOR LINEAR FIT = 34.532

PERUCTION DUF TO GENERAL QUADRATIC FIT WITH DF 2, = 0.200

REDUCTION M.S. = 0.100

PERIDUAL S.S. = 0.800 OF = 151 RESIDUAL M.S. = 0.005

F FOR QUADRATIC FIT = 18.967

REDUCTION DUF TO CUADRATIC TERM ALUNE, WITH 1 OF = 0.015

F FOR QUADRATIC TERM ALONE = 2.860

REDUCTION DUE TO GENERAL CURIC FIT WITH OF 3. = 0.204

REDUCTION M.S. = 0.68

PERIOUAL S.S. = 0.796 OF = 150 RESIDUAL M.S. = 0.005

FOR GENERAL CURIC FIT = 12.786

FOR CURIC TERM ALONE = 0.635
```

I-14

THE PHENICTOR VARIABLEIXE IS SHEE AND THE CRITERION VARIABLEIVE IS TR. SCORE.

```
TEST HEARS AND STANDARD DEVIATIONS
                                                                              0.266
                                                                                                                                 0.280
                        SOUAR
                                                                               0.14E
0.096
5.552
                                                                                                                                   0.220
                        X CHRE
                                                                                                                                   2.371
  CORRECATION HATRIX
                                                         1.000
                       SQUAR
                                                                                                                                   0.941
                                                                                                                                                                                     0.846
                                                                                                                                                                                                                                        0.361
                                                                              0.941
                                                                                                                                   1.000
                                                                                                                                                                                                                                        0.345
                       A CUBE
                                                                                                                                                                                     1.000
                                                                               0.161
                                                                                                                                                                                                                                        1.000
    FIRST. SECOND. AND THIRD REGREE PULYNOHIALS.
  HULTIPLE & SQUARE = 0.1
                                                                                                0.131
 HULTIPLE R = 0.3A1
N.D.F.1 = 1
N.O.F.2 = 36
          O.F.> = 55
For analysts of variance on R =
  PFTA WEIGHTS
 0.361
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS: 2ND COLUMN;
0.131 1.000
  SOUARED RETA WEIGHTS
  0.13L
A METGHTS
           3.042
  INTERCEPT CONSTANT .
                                                                                                      4.738
 MULTIPLE R SQUARE = 0.131
MULTIPLE R = 0.362
N.O.F.1 = 2
N.O.F.2 = 55
  N.D.F.2 = 55
F FOR ANALYSIS OF VARIANCE ON R =
F FOR ANALYSIS CONTRACTOR AND RESONAND COLUMN)

AND REGRESSION FACTOR LUADINGS+ 2ND COLUMN)

TO COLUMN COLU
 AND REGRESSION FACTOR

J X 0.116

2 SOUAR 0.015

SOUARED BETA VEIGHTS

0.102 0.002

8 WFIGHTS

2.711 0.474

INTERCEPT CONSTANT =
                                                                                                       4 - 761
 MULTIPLE R SQUARE = 0.133
MULTIPLE R = 0.364
N.D.F.1 = 3
N.D.F.2 = 54
F FOR ANALYSIS OF VARIANCE ON R =
F FOR ANALYSTS OF BETA WEIGHTS O.063 0.49F =0.423 CONTRIBUTIONS TO MULTIPLE CORRELATION AND REGRESSION FACTOR LOADINGS. 2ND COLUMNS X 0.023 0.99Z 0.241 0.947 0.241 0.947
 SOUARED RETA WETCHTS
0.004 0.487 0.179
8 WEIGHTS
0.534 7.440 -5.306
INTERCEPT CONSTANT =
                                                                                                          4.814
```

```
ANOVA TARLE FOR POLYNCHIALS

DENUCTION DUF TO LINEAR FIT. WITH 1 OF = 0.131

RESIDUAL S.C. = 0.869 DF = 56 RESIDUAL M.S. = 0.016

F FOR LINEAR FIT = 8.407

REDUCTION DUF TO GENERAL QUADRATIC FUT WITH DF 2, = 0.131

REDUCTION M.S. = 0.065

UFCIONAL S.C. = 0.869 DF = 55 RESIDUAL M.S. = 0.016

F FOR GUADRATIC FIT = 4.126

REDUCTION DUF TO GUADRATIC TERM ALONE, WITH 1 OF = 0.000

F FOR GUADRATIC TERM ALONE = 0.014

REDUCTION DUF TO GENERAL CUBIC FIT WITH DF 3, = 0.133

REDUCTION M.S. = 0.044

OF FOR GENERAL CUBIC FIT WITH DF 3, = 0.133

REDUCTION M.S. = 0.044

OF FOR GENERAL CUBIC FIT = 2.755

PAGE OF TO CUBIC TERM ALONE WITH 1 OF = 0.002
```

..........

I-15

THE PREDICTOR VARIABLE(X) IS PRE, AND THE CRETERION VARYABLE (Y) IS TO SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                              0.414
                                                  0.399
         SQUAR
         X CURE
                              5.552
                                                  2.371
 CORRELATION MATRIX
                              1.000
                                                  0.976
                                                                     0.934
                                                                                        0.322
         SQUAR
                              0. 776
                                                  1.000
                                                                                        0.350
         X CURE
                              0 . .. 36
                                                  0.989
                                                                     1.000
                                                                                        0.362
                                                  0.350
                                                                                        1.000
 FIRST, SECOND, AND THIRL NEGREE PULYNOMIALS.
 MULTIPLE R SQUARE = MULTIPLE R = 0+32
N+D.F+1 = 1
                          0.322
 N.D.F.2 = 1
N.D.F.2 = 56
F FOR ANALYSIS OF VARIANCE ON R =
 P. TA WEIGHTS
 CONTRIBUTIONS TO MULTIPLE CORRELATION AND REGRESSION FACTOR LOADINGS. 2ND COLUMN; 3 x 0.104 1.000
 SQUARED RETA HEIGHTS
0.104
R WEIGHTS
1.914
INTERCEPT CONSTANT
                                      4:580
 ***************
 MULTIPLE R SQUARE .
                                   0.131
MULTIPLE R = 0.362
N.D.F.2 = 2
N.D.F.2 = 55
F FOR ANALYSIS OF VARIANCE ON R =
F FOR ANALYSIS UNBETA WEIGHTS

-0.417 0.747

CONTRIBUTIONS TO MULTIPLE CORRELATION

AND REGRESSION FACTOR LOADINGS 2ND COLUMN)

1 -0.134 0.891

0.265 0.968
0.174 0.973
R WFIGHTS
+2.475 4.972
INTERCEPT CONSTANT =
                                      4.773
 ******************
MULTIPLE R SOJARE = 0.134
MULTIPLE R = 0.367
N-D.F.1 = 3
N.D.F.2 = 54
F FOR ANALYSIS OF VARIANCE ON R =
BETA WEIGHTS

0.378 -1.251 1.246

CONTRIBUTIONS TO PULTIPLE CORRELATION

AND REGPESSION FACTOR LOADINGS. 2ND COLUMN;

I 0.122 0.879

2 SQUAR -0.438 0.955
3 X CUBF O
SOUARED BETA WEIGHTS
                                                 0.987
0.143 1.565 1.552
B WEIGHTS
                           0.666
2.245 -A.JT6 0.6
INTERCEPT CONSTANT =
```

```
ANOVA TABLE FOR POLYNCHIALS
REDUCTION DUE TO LINEAR FIT. WITH 1 DF
         RESIDUAL S.S. . FOR LINEAR FIT .
                                 0.896 DF .
                                                          56 RESTOUAL M.S. .
                                     6.493
          REDUCTION DUE TO GENERAL QUADRATIC FIT WITH OF 2. .
         REDUCTION H.S. = 0.065
RESIDUAL S.S. = 0.869 OF =
         F FOR QUADRATIC FIT = 4:142
REDUCTION DUF TO QUADRATIC TERM ALONE, WITH 1 DF. =
F FOR QUADRATIC TERM ALONE = 1:709
         REDUCTION DUE TO GENERAL CUBIC FIT WITH OF 30 4
         REDUCTION H.R. ..
RESTOUAL S.S. ..
                                    0.045
                                     0.866 DF =
                                                          54 RESTOUAL H.S. &
                                                                                          0.016
ERIC FOR GENERAL CURIC FIT = 2.796
EDUCTION OUF TO CURIC TERM ALONE WITH 1 DF = .
```

THE PREDICTOR VARIABLE(X) IS ORE. AND THE CRITERION VANIABLE(Y)- IS THE SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                                0.886
          SOUAR
                                0.411
                                                     0.214
         M CURE
                                0.751
                                                     0.261
 CORRELATION MATRIX
                                1.000
                                                     0.954
                                                                         0.901
0.989
                                                                                              0.387
          SOUAR
                                                     1.000
          X CUBE
                                0.901
                                                     0.989
                                                                         1.000
                                                                                              0.381
                                0.387
                                                     0.388
                                                                                               1.000
 FIRST, SECOND. AND THIRD REGREE PULYNOHIALS.
  MULTIPLE R SOLARE # 0.149
MULTIPLE R = 0.387
N.O.F.1 = 1
N.O.F.2 = 56
FEGR ANALYSIS OF VARIANCE ON R =
 CETÀ WEIGHTS
0.387
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS. 2ND COLUMN)
7 X 0.149 L.000
 T X 0
SQUARED BETA WEIGHTS
 0.149
B WFIGHTS
    5.706
 INTERCEPT CONSTANT # 0+494
MULTIPLE R SOMARE = 0.154

MULTIPLE R = 0.392

N.O.F.1 = 2

N.O.F.2 = 55

F. FOR ANALYSIS OF VARIANCE ON R =
                                                                    4.995
DETS HEIGHTS

0.178 0.719

CONTRIBUTIONS TO MULTIFLE CORRELATION

AND MEGRESSION FACTOR LOADINGS: 2ND COLUMN)
            X 0.069
SQUAR 0.085
                                                    0.966
                                                    0.991
                                 0.085
SOUARED RETA WEIGHTS
0.032 0.048
R WEIGHTS
2.625 2.419
[NTERCEPT CONSTANT = ]
MULTIPLE R COHARE = 0.169
MULTIPLE R = 0.411
N.D.F.1 = 3
N.D.F.2 = 54
F FOR ANALYSIS OF VARIANCE ON R =
BETA HEIGHTS

1.813 -4.777 3.467

CONTRIBUTIONS TO FULTIPLE CORRELATION

CONTRIBUTIONS TO FUTURE LOADINGS. 2
      AND PEGRESSION FACTOR LOADINGS: 2ND COLUMN)

1 0.701 0.939
2 SQUAR -1.853 0.944
3 X CURF 1.322
SOUARED BETA WEIGHTS
3.287 22.747 12.023
B WEIGHTS
                                                    0.927
76.763 -52.748 31.515
INTERCEPT CONSTANT =
                                          0+953
```

```
ANOVA TARLE FOR PCLYNCHIALS
       REDUCTION DUE TO LINEAR FIT. WITH 1 DF = 0.149

PESTOUAL S.S. = 0.851 DF = 56 RESTOUAL M.S. 9

F FOR LINEAR FIT = 9.838
        ********************************
         REDUCTION DUE TO GENERAL QUADRATIC FIT WITH DF 2. .
       REDUCTION M.S. = 0.077

RESTRIBLE S.S. = 0.846 DF = 55 RESTRUAL
F FOR QUADRATIC FIT = 4.995

REDUCTION DUF TO GUADRATIC TERM ALONE. WITH 1 DF. =
F FOR QUADRATIC TERM ALONE = 0.280
                                                                 55 RESTOUAL M.S. .
       REDUCTION DUF TO DENERAL CUBIC FIT WITH DE 3. .
     REDUCTION M.S. #
                                       0.054
                                        n.A31 OF =
                                                                 54 RESIDUAL M.S. .
ERIC OR GENERAL CURIC FIT = 3.668

GROUND ON CHAIC TERM ALONE WITH 1 OF = .
```

1.012

..................

I-17

THE PREDICTOR VARIABLE (X) IS WEE, AND THE CRITERION VARIABLE (Y) IS THE SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                                          0,542
0,729
0,216
                                                                    0.190
               SQUAR
                                                                   0.207
0.189
2.371
               X CURE
    CORRELATION MATRIX
                                                                                            0.942
0.988
             SQUÂR
                                                                   1.000
                                          0.982
                                                                                                                      0.3A2
              X CUBE
                                          0.942
                                                                                            1.000
    FIRST. SECOND. AND THIRD REGREE PULYNOMIALS.
      *************
  MULTIPLE R SOHARE = 0.108
MULTIPLE R = 0.434
N.D.F.1 = 1
N.D.F.2 = 56
F FOR ANALYSIS OF VARIANCE DN R =
BETS WEIGHTS
0.434
  RETA MELOTICO AND AND AND REGRESSION FACTOR LOADINGS. 2ND COLUMNI NO REGRESSION FACTOR LOADINGS. 2ND COLUMNI NO REGRESSION FACTOR LOADINGS.
  0.166
B WFIGHTS
  5.426
INTERCEPT CONSTANT .
  ***************
 MULTIPLE R SQUARE = 0.241
MULTIPLE R = 0.491
N.O.Fe1 = 2
N.O.Fe2 = 55
  N.D.F.Z = 55
F FOR ANALYSIS OF VARIANCE DN R =
 F FOR ANALYSTS OF VARIANCE UN N = BETA METCHTS
1.616 -1.204
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LDADINGS: 2ND COLUMN)
1 0.701 0.884
2 SQUAR -0.460 0.776
 SOUARED RETA VEIGHTS
2.610 1.449
B WEIGHTS
 20.199 -13.RAG
INTERCEPT CONSTANT # -0
                                                 -0.852
MULTIPLE R SOHARE = 0.492
MULTIPLE R = 0.541
N.O.F.1 = 3
N.O.F.2 = 54
F.OR ANALYSIS OF VARIANCE DN R =
BETA WEIGHTS
5.460 -9.934 5.016

CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LPADINGS: 2ND COLUMN;
1 X 2.370 0.803
2 SQUAR -3.796 0.706
                                     -3.796
1.718
? SQUAR =3.796
3 × CURF 1.718

SQUARED BETA WEIGHTS
29.815 98.487 25.158
8 WEIGHTS
68.262=113.948 63.029
INTERCEPT CONSTANT = -7.572
                                                                0.633
```

```
ANOVA TABLE FOR POLYNCHIALS

REOLCTION DUF TO LINEAR FIT, WITH 1 DF = 0.53

RESIDUAL S.S. = 0.812 OF = 56 RESIDUAL N.S. = 0.014

F. FOR LINEAR FIT = 12.997

REDUCTION DUF TO GENERAL QUADRATIC FIT WITH DF 2, = 0.241

REDUCTION M.S. = 0.121

PESIDUAL S.S. = 0.759 DF = 55 RESIDUAL M.S. = 0.014

F. FOR QUADRATIC FIT = 8.746

REDUCTION DUF TO QUADRATIC TERM ALONE, WITH 1 OF, = 0.053

F. FOR QUADRATIC TERM ALONE = 3.636

REDUCTION DUF TO GENERAL CUBIC FIT WITH DF 3, = 0.292

REDUCTION M.S. = 0.708 DF = 54 RESIDUAL M.S. = 0.013

FOR GENERAL CUBIC FIT = 7.441

CITON DUF TO CURIC TERM ALONE WITH 1 OF = , 0.051

ERIC GROWERAL CUBIC FIT = 7.441

CITON DUF TO CURIC TERM ALONE WITH 1 OF = , 0.051
```

POLYNOMIAL FITTING FOR 2 VARIABLES-SHE AND TO SCORE 139 OBSERVATIONS.

```
FOR ANALYSIS OF VARIANCE ON R = 6.961

BETA WEIGHTS
-1.325 1.145

CONTRIBUTIONS TO PULTIPLE CORRELATION
AND REGRESSION FACTOR LUADINGS. 2ND COLUMN;
1 x 0.223 -0.551
2 SQUAR -0.130 -0.360

SOUÁREO RETA WEIGHTS
1.758 1.665

B WEIGHTS
-5.788 6.550
 -5.788 6.549
INTERCEPT CONSTANT =
```

3 % CUBE -0.17 SOUARED BETA WEIGHTS 1.133 19.677 11.102 8 WEIGHTS 4.647 =24.844 22.430 INTERCEPT CONSTANT =

MULTIPLE R SQUARE = 0.00
MULTIPLE R = 0.305
N-D.F-1 = 2
N-D.F-2 = 136

N.D.F.R = 136 F FOR ANALYSIS OF VARIANCE ON R =

ANOVA TABLE FOR POLYNCHIALS PERIOUAL S.S. = 0.972 OF = .37 RESIDUAL N.S. F FOR LINEAR FIT = 3.979

REDUCTION OUF TO GENERAL QUADRATIC FIT WITH OF 2. . REDUCTION US. 0.044
REGULATION M.S. 0.046
REGULATION M.S. 0.907 DF 136 RESIDUAL
F FOR QUADRATIC FIT 0.9061
REDUCTION OUF TO QUADRATIC TERM ALONE 10.690
F FOR QUADRATIC TERM ALONE 9.690 136 RESIDUAL N.S. .

REDUCTION DUE TO GENERAL CUBIC FIT WITH DF 3. a REDUCTION M.S. a 0.044RESIDUALS :: . . 0.869 OF . 135 RESTOUAL M.S. .

5.972

ERIC GENERAL CURIC FIT = 6.401
TION OUF TO CURIC TERM ALONE WITH 1 DF = .

I-19

RD RATING

```
THE PREDICTOR VARIABLE(X) IS PRE. AND THE CRITERION VARIABLE(Y) IS TH SCORE.
```

```
TEST HEARS AND STANDARD DEVIATIONS
                            SQUAR
                                                                 0.475
                                                                                                                        0.473
    T X CURE
A Y
CORRELATION MATRIX
                                                                                                                                                 0.444
                                                                                          0.425
                                                                                         2.406
                            X
SQUAR
                                                              1.000
                                                                                                                                                                                                          0.982
0.997
                                                                                                                                                 0.994
                                                                                                                                                                                                                                                             *0.215
                                                                                       0.994
                                                                                                                                                1.000
                            X CUBE
                                                                                                                                                                                                          1.000
                                                                                -0.715
                                                                                                                                                                                                                                                                   1.000
    FIRST. SECOND. AND THIRD REGREE PULYNOMIALS.
   MULTIPLE R SOUARE 0.046
MULTIPLE R n n.215
N.O.F.1 = 1
N.O.F.2 = 137
F FOR ANALYSIS OF VARIANCE ON R =
    BETT WEIGHTS
   I X G
SOUARED RETA WEIGHTS
   0.046
B WFIGHTS
  . =0.710
INTERCEPT CONSTANT =
  MULTIPLE R SCHARE # 0.046
MULTIPLE R # 0.215
N.D.F.1 # 2
N.D.F.2 # 136
F FOR ANALYSIS OF VARIANCE ON R #
 SOUAR O. SOUARED BETA MEIGHTS O. 62 O. 600 R MEIGHTS MEIGHTS MEIGHTS MEIGHTS MEIGHTS MEIGHT FOR CONSTANT MERCEPT CONSTANT MER
                                                                                                          3.143
      MULTIPLE R SCHARE = 0.093
MULTIPLE R = 0.304
N.O.F.1 = 3
N.O.F.2 = 135
F FOR ANALYSIS OF VARIANCE ON R =
F FOR ANALYSIS OF VARIANCE UN H = 4.592
BETA WEIGHTS
10.549 -25.973 15.252
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS: 2ND COLUMN;
1 4 -2.263 -0.705
2 SOUAR 5.527 -0.701
3 X CUBE +3.171
SOUAREO BETA WEIGHTS
                                                                                                                                        -0.663
111.283 670.981 232.636
R WEIGHTS
34.915 -88.888 53.730
INTERCEPT CONSTANT #
                                                                                                                   3:132
```

```
ANOVA TABLE FOR POLYNCHIALS

PROUCTION BUF TO LINEAR FIT, WITH 1 DF = 0.046

RESIDUAL S.S. = 0.054 DF = 137 RESIDUAL M.S. = 0.007

F FOR LINEAR FIT = 6.612

GENUCTION DUF TO GENERAL QUADRATIC FIT WITH DF 2. = 0.046

REDUCTION M.S. = 0.054 OF = 136 RESIDUAL M.S. = 0.007

F FOR QUADRATIC FIT = 3.282

REDUCTION OUF TO GUARRATIC TERM ALONE, WITH 1 CF. = 0.000

F FOR QUADRATIC TERM ALONE = 0.000

REDUCTION LUF TO GENERAL CUBIC FIT WITH OF 3. = 0.003

REDUCTION M.4. = 0.031

REGIDUAL S.4. = 0.907 DF = 135 RESIDUAL M.S. = 0.007

FOR GENERAL CURIC FIT = 4.592

OUCTION DUF TO CURIC TERM ALONE WITH 1 DF = 10.0047

FOR CURIC TERM ALONE = 6.926
```

THE PREDICTOR VARIABLE(X) IS GRE. AND THE CRITERION VARIABLE(Y) IS TO SCORE.

```
TEST HEARS AND STANDARD DEVIATIONS
1 % 0.969 0.054
2 SQUAR 0.943 0.091
                                                    0.054
          X CURE
                                                    0.119
                                0,919
                                2.406
 CORRELATION MATRIX
                                                    0.992
                                                                        0.972
0.994
1.000
                                1.000
                                                                                             0.021
          SQUAR
                                                    1.000
                                                                                           0.006
          X CUBE
                                0.972
                                                    0+006
                                                                                             1.000
                                1.ó21
                                                                       -0.009
 FIRST: SECOND: AND THIRD REGREE PULYNOMIALS.
                          0.0%0
 MULTIPLE R SOHARE = 0.000
MULTIPLE R = 0.001
N.O.F.1 = 1
N.O.F.2 = 137
F.FOR ANALYSIS OF VARIANCE ON R =
                                                                   0.060
 BFT& WEIGHTS
 BFTA WELDTIS

0.021

CONTRIBUTIONS TO MULTIPLE CORRELATION

AND REGRESSION FACTOR LOADINGS: 2ND COLUMN;

9.000 1.000
 SOURRED BETA WEIGHTS
 0.000
B MEIGHTS
 0.603
INTERCEPT CONSTANT =
                                        2.222
 MULTIPLE R SOHARE = 0.013
MULTIPLE R = 0.113
N.O.F.1 = 2
N.O.F.2 = 136
F FOR ANALYSIS OF VARIANCE ON R =
PERCEPT CONSTANT .
                                        -7:602
 000,0200000000000000000000
MULTIPLE R SQUARE 0.047
MULTIPLE R 0.217
N.O.F.1 3
N.O.F.2 135
F FOR ANALYSIS OF VARIANCE ON R =
F FOR ANALYSIS OF BETA WEIGHTS

39.917 -85.574 46.272
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS: 2ND COLUMN)
T X 0.833 0.096
-0.550 0.030
3 ULAR -0.551

3 K CUAR -0.216

SOUARED BETA METGHTS

593.4047322.9862141.134

8 METGHTS
152,2510000000 610,244
INTERCEPT CONSTANT # -280.674
```

```
ANOVA TABLE FOR POLYNCHILLS

PROUCTION OUF TO LINEAR FIT, WITH 1 DF = 0.000

REGIOVAL S.S. = 1.000 DF = 137 RESTOUAL M.S. = 0.007

F FOR LINEAR FIT = 0.060

PROUCTION OUF TO GENERAL QUADRATIC FIT WITH DF 2. = 0.013

REDUCTION M.S. = 0.006

PESCIOVAL S.R. = 0.987 DF = 136 RESTOUAL M.S. = 0.007

F FOR QUADRATIC FIT = 0.073

REDUCTION OUF TO GUADRATIC TERM ALONE. WITH 1 DF = 0.012

F FOR QUADRATIC TERM ALONE = 1.686

REDUCTION DUE TO GENERAL CURIC FIT WITH DF 3. = 0.047

REDUCTION M.S. = 0.016

RESTOUAL S.S. = 0.953 DF = 135 RESTOUAL M.S. = 0.007

FOR GENERAL CURIC FIT = 2.218

FOR CURIC TERM ALONE = 4.858
```

THE PREDICTOR VARIABLE(X) IS WRE. AND THE CRITERION VARIABLE(Y) IS TO SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                                 0.417
                                                     0.183
           SQUAR
           X CUBE
                                 2.406
                                                      1.504
   CORRELATION HATRIX
                                1.000
                                                                         0.949
                                                     0.983
                                                                                            -0.051
-0.075
           SQUAR
                                 0.943
                                                     1.000
           X CUPE
                                 0.949
                                                     0.990
                                                                          1.000
                                                                                            -0.085
                                                   -0.075
                              -0.051
                                                                        -0.084
                                                                                              1.000
    FIRST. SECOND. AND THIRD REGREE PULYNOMIALS.
   MULTIPLE R SOHARE = 0.043
MULTIPLE R = 0.05%
N.O.F.1 = %
M.O.F.2 = 137
F FOR ANALYSIS OF VARIANCE ON R =
  BETS WEIGHTS
  -0.051
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS: ZND COLUMN)
1 0.003 -1.000
SOUARED BETA WEIGHTS
  0,003
R W/IGHTS
=0,438
  INTERCEPT CONSTANT .
                                       3.076
  ----
 MULTIPLE R SOLARE = 0.021
MULTIPLE R = 0.145
N.O.F.1 = 2
N.O.F.2 = 136
F FOR ANALYSIS OF VARIANCE ON R =
  MULTIPLE A SQUARE .
F FOR ANALYSTS OF

BFTA WEIGHTS

0.669 -0.773

CONTRIBUTIONS TO MULTIPLE CORRELATION

AND REGRESSION FACTOR LOADINGS: 2ND COLUMN;

1 X -0.034 -0.354

-0.055 -0.521
 I X =0.
2 SOUAR
SOUARED RETA VETOMTS
0.448 0.538
R WETOMTS
 5.715 -5.744
INTERCEPT CONSTANT -
                                      1.368
 MULTIPLE R SQUARE 0 0.001
MULTIPLE R = 0.204
MULTIPLE R = 0.204
N.O.F.1 = 3
N.O.F.2 = 135
F FOR ANALYSIS OF VARIANCE ON R =
F FOM ANALTSIS OF THRESTEE OF SETTA WEIGHTS
3.072 46.309 3.337
CONTRIBUTIONS TO PULTIFIE CORRELATION
AND REGRESSION FACTOR LOADINGS, 2ND COLUMN;
1 X -0.158 -0.252
2 SQUAR 0.463 -0.371
                                                                  1 . 9A A
3 X CURE -0.284
SOUARED BETA VETONTS
9.439 40.947 11.134
R WETONTS
   26.233 -44.617 22.482
 INTERCEPT CONSTANT .
  ANOVA TABLE FOR POLYNCHIALS
REDUCTION DUF TO LINEAR FIT, WITH 1 OF 0.003
RESTDUAL 5.5. 0.997 OF 137 RESTOUAL M.S. 0.007
F FOR LINEAR FIT 0.362
  REDUCTION DUF TO GENERAL GUADRATIC FIT WITH OF 2. .
REDUCTION M.R. = 0.010
RESIDUAL S.S. = 0.979 OF =
F FOR 2040RATIC FIT = 1.457
r FOR QUADRATIC FIT = 1457
REDUCTION DUF TO CUADRATIC TERM ALONE, WITH 1 DF. =
F FOR QUADRATIC TERM ALONE = 2.548
                                                                136 RESTOUAL M.S. .
```

0.014 0.959 OF . 0.041

135 RESTOUAL M.S. .

REDUCTION DUE TO GENERAL CUBIC FIT WITH OF 3. .

RESTOUAL 5.5.

POLYNOMIAL FITTING FOR

SOHAR

X CURE

CORRELATION MATRIX

¥

SOUAR

X CURE

BFT& WEIGHTS

TEST MEANS AND STANDARD DEVIATIONS 954.0 107.0

0.424

1.000

0.960

0.211

MULTIPLE R SOHARE = 0.044

MULTIPLE R = 0.211

N.D.F-1 = 1

N.G.F-2 = 135

F FOR ANALYSIS OF VARIANCE ON P =

CONTRIBUTIONS TO PULTIPLE CORRELATION

TO THE O. 044 SOURCE RETAINS

0.044
R WFIGHTS
1.228
INTERCEPT CONSTANT #

0.028 0.146 WFIGHTS -0.974 2.181 INTERCEPT CONSTANT -

MULTIPLE R SOLIARE = 0.068
MULTIPLE R SOLIARE = 0.260
N.O.F.1 = 3
N.O.F.2 = 133
F FOR ANALYSIS OF VARIANCE ON R =

FIRST, SECOND, AND THIRD DEGREE PULYNOMIALS.

AND REGRESSION FACTOR LOADINGS+ 2ND COLUMN)

3:681

3.968

2 VARIABLES-SHE AND IP SCORE- 137 GHSENVATIONS.

0.211

n.234

1.000

THE PREDICTOR VARIABLE(X) IS SHE, AND THE CRITERION VARIABLE.Y) IS IF SCOPE.

0.907 0.988 1:000 0.244

3.348

9.356

0.960

1.000

0.244

I-23

RM RATING

THE PREDICTOR VARIABLE(X) IS PHE, AND THE CRITERION VARIABLE(Y) IS TP SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                         0.400
0.756
                                                   0.341
           SQUAR
                                                   0.342
           X CHUE
                                0.719
                               4.453
   CORRELATION MATRIX
           A 1.000
                                                   0.985
                                                                      0.953
                                                                                          0.156
                                                  1.000
                                                                                         0.164
           X CUBE
                               0.953
                                                                      1.000
                                                   0.104
                               0.154
                                                                      0.167
               *************************
  FIRST. SECOND. AND THIRD REGREE PULYNONIALS.
   0 · 024
  MULTIPLE R SQUARE = 0.024
MULTIPLE R = 0.156
N.O.F.1 = 1
N.O.F.2 = 139
F FOR ANALYSIS OF VARIANCE ON R =
  BFT& WEIGHTS
 CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS. 2ND COLUMN;

1 2 0.024 1.000
  SOUARED RETA WEIGHTS
  0.024
R WEIGHTS
  INTERCEPT CONSTANT = 3:747
  *************
 MULTIPLE R SQUARE = 0.1A6
N.O.F.1 = 2
N.O.F.2 = 134
                                    0.028
 N.O.F.2 = 134
F FOR ANALYSIS OF VARIANCE ON R =
BETA MEIGHTS
-0.159 0.320
CONTRIBUTIONS TO MULTIPLE CORRELATION
      AND REGRESSION FACTOR LOADINGS, 2ND COLUMN;

1 x =0.025 0.942

2 SOUAR 0.052 0.986
 SOUAPED BETA WEIGHTS
0.025 0.142
B WEIGHTS
 =0.894 1.861
INTERCEPT CONSTANT = 3.807
 ********************
MULTIPLE R SOHARE 0.021
MULTIPLE R 0.177
N.O.F.1 = 7
N.O.F.2 = 133
F FOR ANALYSIS OF VARIANCE ON R =
F FOR ANALYSIS OF VARIANCE ON R = 1.428
BETA WEIGHTS
1.952 -4.746 2.710
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS, 2ND COLUMN;
1 X 0.290 0.886
2 SQUAR -0.712 0.927
3 X CURF 0.453 0.947
3 X CURF 0.453
SDUARED RETA VETONTS
3.430 18.890 7.345
B WEIGHTS
                                                0.947
10.440 -24.648 15.610
INTERCEPT CONSTANT = 3.795
```

dest by ERIC

THE PREDICTOR VARIABLE(X) IS GRE, AND THE CRITERION VARIABLE(Y) IS TO SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                          0.997
                                             0.007
         SQUAR
                                             0.013
         X CHRE
                           0.992
                            4.453
                                             1.925
  CORRELATION MATRIX
                          1.000
                                             1.000
                                                             1.000
                                                                              0.041
         SQUAR
                           1.000
                                             1.000
                                                             1.000
                                                                             0.946
         X CUSE
                           1.000
                                             1.000
                                                             1.000
                           0+041
                                            0.044
                                                             0.046
                                                                              1.000
           FIRST. SECOND. AND THIRD REGREE PULTNOHIALS.
  *********************
 MULTIPLE R SQUARE D.022

MULTIPLE R = 0.041

N.O.F.1 = 1

N.O.F.2 = 135

F FOR ANALYSIS OF VARIANCE ON R =
  BETA MEIGHTS
 n,041
CONTRIBUTIONS TO PULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS: 2ND COLUMN)
1 7 0.002 1.000
 0.002
R WF[GHTS
 12.035
INTERCEPT CONSTANT =
                                -7.551
  MULTIPLE & SQUARE .
                                0.140
 MULTIPLE R .
                      0.374
 N.O.F.2 4
 N.O.F.2 4 134
F FOR ANALYSIS OF VARIANCE ON R =
BETA WEIGHTS
-51.053 51.05
CDNTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS, 2ND COLUMN;
1 -2.092 0.116
2.231 0.117
 BETA WEIGHTS
 SOUTRED BETA HETGHTS
 R WEIGHTS
 PROPERTY CONTENT - 7386-344
 MULTIPLE R SQUARE = 0.372
MULTIPLE R = 0.372
Non-F-1 w 3
Non-F-2 = 133
                                 0.139
N.O.F.2 = 133
F FOR ANALYSIS OF VARIANCE ON R =
F FOR ANALYSIS UP 3

BFTA WEIGHTS

-15.047 19.928 15.563

CONTRIBUTIONS TO MULTIPLE CORRELATION

AND REGRESSION FACTOR LOADINGS. 2ND COLUMNS

7

-1.036
0.853
0.117
0.853
SOUPPED BETA WEIGHTS
278,327 381,346 242,393
A METGHTS
###10013
#######2921.4971980.930
INTERCEPT CONSTANT # 5796.982
```

```
ANGVA TARLE FOR POLYNCHIALS
                                                                                                                                                                                              2888888888888888888
                               HENUCTION DIF TO LINEAR FIT, WITH 1 DF = 0.002
RESTOUAL S.S. = 0.498 DF = 135 RESTOUAL M.S. =
F FOR LINEAR FIT = 0.227
                                   REDUCTION DUF TO GENERAL QUADRATIC FIT WITH DF 2, .
                            REDUCTION HOS TO DENERAL MOMENTALIC FIT THE DESTRUCTION HOS TO 0.000 DF TO 134 RESTOUAL FOR QUADRATIC FIT TO 10.000 MITH 1 PF TO FOR QUADRATIC TERM ALONE WITH 1 PF TO FOR QUADRATIC TERM ALONE WITH 1 PF TO FOR QUADRATIC TERM ALONE TERM ALONE WITH 1 PF TO THE MOMENTAL TER
                                                                                                                                                                                                                                                  134 RESTOUAL H.S. .
                              REDUCTION DUF TO GENERAL CURIC FLT WITH OF 3. #
                            REDUCTION H.S. .
                                                                                                                                                   0.046
                                RESTOUAL S.S. .
                                                                                                                                                                                                                                                133 RESIDUAL M.S. .
FOR GENERAL CURIC FIT = 7-131

POLICTION OUF TO CURIC TERM ALONE WITH 1 OF = .

FOR CURIC TERM ALONE = -0+156.
```

I-25

THE PREDICTOR VARIABLE (X) IS DRE. AND THE CRITERION VARIABLE (Y) IS THE SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                           0.632
           PAUDZ
                                   0.453
          X CURE
                                   0.342
  CORRELATION MATRIX
                                                         1.925
                        1.000
          X
SOUAR
                                                        0.956
                                                                              0.898
                                                                                                    0.766
                                                        1.000
                                   0.956
                                                                                                    0.390
          x CUBE
                                  0.498
                                                         0.390
                                  0.366
 FIRST, SECOND. AND THIRD OFFIEE PULYNOHIALS.
 FIRST, SECOND, AND THIRD OF GREE POLYNOM

THE STREET STREET

MULTIPLE R SOLIARE # 0.134

MULTIPLE R # 0.366

N.D.FY! = 1

N.D.F.2 = 139

FERR ANALYSIS OF VARIANCE ON R # 2
 BETA WEIGHTS
 CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS. 2ND COLUMN;

1 X A.134 1.000
 T X A SOUARED RETA WEIGHTS
 *0.134
B-WEIGHTS
     3.037
 INTERCEPT CONSTANT .
                                            2.533
  **************
 MULTIPLE R SQUARE = MULTIPLE R = 0+391
MULTIPLE R = 0.391
N.D.F.1 = 7
N.D.F.2 = 134
F. FOR ANALYSIS OF VARIANCE ON R = 12.103
RPTA WEIGHTS
-0.609 0.475
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS. 2ND COLUMN;
1 X =0.033 0.935
0.186 0.998
SOUARED RETA VEIGHTS
0.008 0.226
R MEIGHTS
-0.739 3.430
INTERCEPT CONSTANT =
                                            3 276
 ****************
MULTIPLE R SQUARE =
MULTIPLE R = 0.392
N.D.F.1 = 3
N.D.F.2 = 133
                                       0.154
N.D.F.2 = 133
F FOR ANALYSIS OF VARIANCE ON R =
BETA WEIGHTS
-0.260 0.959 -0.324
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS 2ND COLUMNS
4 -0.095 0.933
BETA WEIGHTS
            SOUAR -0.095
                                0,374
                                                       0.996
SOUARED RETA VETONTS
                                                       0.987
SOURCE HEIR VEIGHTS
0,068 0.919 0.105
H WEIGHTS
-2.159 7.320 -2.472
INTERCEPT CONSTANT =
                                             3:347
```

```
ANOVA TARLE FOR POLYNCHIALS

DEDUCTION DUF TO LINEAR FIT. WITH 1 DF = 0.134

RESIDUAL S.S. = 0.866 DF = 135 RESIDUAL N.S. = 0.006

F FOR LINEAR FIT = 20.845

REDUCTION DUF TO GENERAL QUADRATIC FIT WITH DF 2; = 0.153

REDUCTION M.S. = 0.077

RESIDUAL S.S. = 0.867 DF = 134 RESIDUAL N.S. = 0.006

F FOR QUADRATIC FIT = 12.103

REDUCTION DUF TO CUADRATIC TERM ALONE, WITH 1 CF; = 0.019

F FOR QUADRATIC TERM ALONE = 3.046

REDUCTION DUF TO GENERAL CUBIC FIT WITH DF 3; = 0.154

REDUCTION DUF TO GENERAL CUBIC FIT WITH DF 3; = 0.154

REDUCTION DUF TO GENERAL CUBIC FIT WITH DF 3; = 0.006

FOR GENERAL CUBIC FIT = 8.042

PLOTION DUF TO CUBIC FIT = 8.042

PLOTION CUBIC TERM ALONE = 0.086
```

2 VARIABLES-SHE AND TP SCORE- 152 CRSERVATIONS.

THE PHEDICTOR VARIABLE(X) IS PRE- AND THE CRITERION VARIABLE(Y) IS TO SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                              0.452
                                                    0.390
         SQUAR
                                0.385
         X CURE
                                                    2.425
 CORRELATION MATRIX
                                1.000
                                                                         0.923
         SQUAR
                                                    0.972
                                                                                             0.058
                                                    1.000
                                                                                             0.055
                                                                         1.000
         X CUBE
                                                                                             0.065
                                0.050
                                                                                             1.000
 FIRST. SECOND. AND THIRD DEGREE PULYNOMIALS.
 FIRST SECOND. AND THIRD DEGREE PUL
HULTIPLE R SDIARE = 0.073

MULTIPLE R = 0.058

N.D.F.1 = 1

F. FOR ANALYSIS OF VARIANCE ON R =
F FOR ANALTS, ...

BETA WEIGHTS
0.058

CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS; 2ND COLUMN)
3 0.003 1.000
 SOUARED BETA WEIGHTS
    0.003
WEIGHTS
    0,362
 INTERCEPT CONSTANT =
                                          5 · 5 R •
  MULTIPLE R SQUAPE = 0.058
N-0.F-1 = 2
N-0.F-2 = 149
                                        0.063
N.O.F.2 = 149
F FOR ANALYSIS OF VARIANCE ON R = 0.255
8FT4 WEIGHTS
0.083 -0.025
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS: 2ND COLUMN;

| X 0.005 U.995
| SQUAR +0.001
GRESGION FA

I X 0

2 SOUARE DETAINEIGHTS
0,007 0.001

B WEIGHTS
0,514 =0.001

INTERCEPT
 0.51: =0.170
INTERCEPT CONSTANT =
                                           5:577
 MULTIPLE R SCHARE = 0.162
MULTIPLE R = 0.320
N.O.F.1 = .3
N.O.F.2 = 148
F FOR ANALYSIS OF VARIANCE ON R =
 BFTA WEIGHTS
4,864 -11.756
                             7.177
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS. 2ND COLUMN;

X 0.283 0.182
2 SQUAR -0.648 0.172
3 X CURE 0.468
SOUARED RETA WEIGHTS
                                                    0.204
23,659 138.218
8 WFIGHTS
                            51.513
                           50.161
 INTERCEPT CONSTANT .
                                          5.410
```

```
ANOVA TARLE FOR POLYNCHIALS

GENUCTION OUF TO LINEAR FIT. WITH 1 OF = 0.103

RESIDUAL S.S. = 0.997 OF = 150 RESIDUAL M.S. = 0.007

FEAR LINEAR FIT = 0.509

DECIDIAL S.S. = 0.902

DECIDIAL S.S. = 0.997 DF = 149 RESIDUAL M.S. = 0.007

FEOR GUADRATIC FIT = 0.259

REDUCTION DUF TO GUADRATIC TERM ALONE. WITH 1 OF = 0.000

FEAR QUADRATIC TERM ALONE = 0.005

REDUCTION DUF TO GENERAL CUBIC FIT WITH DF 3. = 0.102

REDUCTION DUF TO GENERAL CUBIC FIT WITH DF 3. = 0.102

REDUCTION M.S. = 0.034

RESIDUAL S.S. = 0.898 DF = 148 RESIDUAL M.S. = 0.006

GENERAL CUBIC FIT = 5.634

RECIDIAL TERM ALONE = 16.339
```

ext Provided by ERIC

148 RESTOUAL N.S. .

I-29

Z VARIAULES-GRE AND TP SCORE. 152 ORSERVATIONS.

PALISHIAL FIFTING FOR

RECUCTION DUE TO GENERAL CUBIC FIT WITH DF 3. #
RECUCTION H.S. # 0.059
PECIDUAL 5.5, # 0.822 OF # 148 RESID

ERICCTION OUT TO CURTO TERM ALONE WITH 1 OF = ;

(3)

THE PREDICTOR VARIABLE(X) IS MAE. AND THE CRITERION VARIABLE(Y) IS TO SCORE.

0.207

0.236

TEST MEANS AND STANDARD DEVIATIONS 0.403

SQUAR

X CURE CORRELATION MATRIX

904.0

```
0.921
                                          1.000
                                                              0.973
                                                                                                      0.234
0.265
0.263
                    X
SQUAR
                                                              1.000
                    X CHRE
                                          150.0
                                          0.234
                                                              0.265
                                                                                  0.263
                                                                                                      1.000
            FIRST. SECOND. AND THIRD DEGREE PULYNOMIALS.
             MULTIPLE R SOHARE = 0.055
MULTIPLE R = 0.234
N.O.F.1 = 1
N.O.F.2 = 150
F FOR ANALYSIS OF VARIANCE ON R =
                                                                             8.707
            BETA MEIGHTS
            0.234
CONTRIBUTIONS TO PULTIPLE CORRELATION
              IND REGRESSION FACTOR LOADINGS. 2ND COLUMN;
            SOURRED RETA WEIGHTS
            0.055
B WFIGHTS
2.746
            INTERCEPT CONSTANT # 4:128
            MULTIPLE R SOHARE = 0.000

MULTIPLE R SOHARE = 0.000

MULTIPLE R = 0.283

N.O.F.1 = 2

N.O.F.2 = 149
            F FOR ANALYSIS OF VARIANCE ON R = 6.467
            PERTA WEIGHTS

-0.426 0.479

CONTRIBUTIONS TO FULTIPLE CORRELATION

AND REGRESSION FACTOR LOADINGS: 2ND COLUMN;

1 40.100 0.829

2 SOUAR 0.160 0.936
           SOUAPED BETA WEIGHTS
0.182 0.441
8 WEIGHTS
-4.999 6.937
INTERCEPT CONSTANT =
                                                    5.982
             MULTIPLE R SOHAPE = 0.123

MULTIPLE R = 0.351

N.O.F.1 = 3

N.O.F.2 = 149

F FOR ANALYSIS OF VARIANCE ON R =
          F FOR ANALYSIS OF BETA WEIGHTS

-2,679 6.135 -3.314

CONTRIBUTIONS TO MULTIPLE CORRELATION

AND REGRESSION FACTOR LOADINGS, 2ND COLUMN;

T -0.628 0.668

2 SQUAR 1.624 0.754

2 X CURE -0.873 0.751
            B WEIGHTS -31.415
            -31.415 62.653 -34.060
INTERCEPT CONSTANT =
             ANNYA TARLE FOR POLYNCHIALS
            PERUCTION DUE TO LINEAR FIT, WITH 1 DF = 0.055

RECIDUAL S.C. = 0.945 DF = 150 ...RESIDUAL M.S. &
F FOR LINEAR FIT = 6.707
              REDUCTION DUE TO GENERAL GUADRATIC FIT WITH OF 2. =
            REDUCTION M.S. = 0.040

PECIDUAL S. C. = 0.920 DF = 149 RESIDUAL.

F FOR GUADRATIC FIT = 6.467

REDUCTION DUF TO GUADRATIC TERM ALONE, WITH 1 DF, =

F FOR GUADRATIC TERM ALONE = 4.050
                                                                       149 RESTOUAL, M.S. .
            REDUCTION DUE TO GENERAL CURIC FIT WITH DE 3. .
            REGULTION M.S. . RESIDUAL S.S. .
                                         0.041
0.877 DF =
                                                                           148 RESIDUAL M.S. . 0.006
FOR GENERAL CURIC FIT # 6.925

ERIC EDUCTION DUF TO CURIC TERM ALONE WITH 1 DF # . 0.643

ERIC FOR CURIC TERM ALONE # 7.296
```

I-30

THE PREDICTOR VARIABLE(X) IS SHE, AND THE CRITERION VAHIABLE(Y) IS TH SCORE.

TEST MEANS AND STANDARD DEVIATIONS

```
0.173
                                         0.194
       SQUÂR
       X CUBE
                          0.035
                          5.282
CORRELATION MATRIX
                                                           0.791
0.972
1.000
       X 1.000
SQUAR 0.905
                                          0.905
                                                                            0.198
                                          1.000
                                                                           0.052
       X CUBE
                          0.791
                                                          -0.057
                                                                            1.000
FIRST. SECOND. AND THIRD REGREE PULYNOMIALS.
0.039
N.D.F.2 = 37
F FOR ANALYSIS OF VARIANCE ON R = 1.503
BET# WEIGHTS
BETA WELDALS
0,198
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS. 2ND COLUMN)
T Y 0.039 1.000
0.039
R WFIGHTS
2.024
INTERCEPT CONSTANT =
                                  4.932
MULTIPLE R SCHARE = 0.129
MULTIPLE R = 0.359
N.O.F-1 = 2
N.O.F-2 = 36
  FOR ANALYSIS OF VARIANCE ON R =
BFTA WEIGHTS
0.835 -0.764
O.835 -0.7ñ4
CONTHIGUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS. 2ND COLUMN)
T 0.165 0.501
P SOUAR -0.036 0.144
SOUARED BETA MEIGHTS
0.698 0.496
B WFIGHTS
B.558 -10.348
INTERCEPT CONSTANT = 4.490
MULTIPLE R SOHARE = 0.217
MULTIPLE R = 0.466
N.O.F.1 = 3
N.O.F.2 = 35
..................
N.O.F.2 = 35
F FOR ANALYSIS OF VARIANCE DN R = 3.230
0.215 9.285 7.028
8 WFIGHTS -4.749 44.744 +48.424
           S 44.764 +48.424 4:802
INTERCEPT CONSTANT .
 ANOVA TABLE FOR POLYNCHIALS
REDUCTION DUF TO LINEAR FIT. WITH 1 DF # 0.039

DESIDUAL 5.5. # 0.961 DF # 37 RESIDUAL H.S. #
F FOR LINEAR FIT # 1.503
 REDUCTION DUE TO GENERAL GUADRATIC FIT WITH DE 2. .
REDUCTION M.S. = 0.064

RECIDUAL S.C. = 0.871 OF = 36 RESTOUAL
F FOR QUADRATIC FIT = 2.656

RECUCTION OUF TO GUADRATIC TERM ALCHE, WITH 1 FF. =
F FOR QUADRATIC TERM ALONE = 3.700
                                                       36 RESTOUAL M.S. .
                                                                                0.090
 REDUCTION OUE TO GENERAL CURIC FIT WITH OF 3+ .
REDUCTION M.S. = 0.072
RESTOUAL S.S. = 0.783 DF =
                                                     35 RESIDUAL M.S. D
                                                                                          0.022
```

0.084

ERECTOR OF THE STATE OF THE STA

P FOR ANALYSIS OF VARIANCE ON R = 1.857

BFTA WEIGHTS
-0.502 0.7AR

CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS, 2ND COLUMN;
1 -0.122 0.796
2 SOUAR 0.216 0.918

7 SOUAR 0.
SOUAREO BETA WEIGHTS
0.252 0.591
R WEIGHTS
-2.821 5.754
INTERCEPT CONSTANT 4.879 MULTIPLE R SOMARE = 0.15
NULTIPLE R = 0.39
N.O.F.1 = 3
N.D.F.2 = 35 0.153

T E 0.049 SOUARED RETA WEIGHTS

1.366 INTERCEPT CONSTANT =

0,059 B WFIGHTS

N.D.F.2 = 35 F FOR ANALYSIS OF VARIANCE ON R = N.D. ...

FERR ANALYSTS OF

BETA WEIGHTS

-3.073 6.7=6 -3.575

CONTRIBUTIONS TO PULTIFLE CORRELATION
AND REGRESSION FACTOR LOADINGS: 2ND COLUMN;

J X -0.750 0.622

2 SOUAR 1.906 0.717

3 X CURE -1.002 0.716

3 X CURE +1.00 SQUAREO RETA WEIGHTS 9.506 46.644 12.784 8 WEIGHTS

=17.309 45.448 =27.737 INTERCEPT CONSTANT =

ANOVA TABLE FOR POLYNCHIALS REDUCTION DUF TO LINEAR FIT. WITH 1 DF # 0.059
RESIDUAL 5.5. # 0.941 DF # 37 RESIDUAL M.S. #
F FOR LINEAR FIT # 2.329

REDUCTION DUE TO GENERAL QUADRATIC FIT WITH DE 2. . FOR QUADRATIC TERM ALONE = 1.363 36 RESTOUAL M.S. . 0.034

REDUCTION DUE TO BENERAL CUBIC FIT WITH DF 3+ = REGULCTION H. R. 0.051
RESIDUAL 5. R. 0.847 DF = 35 RESI 35 RESIDUAL M.S. .

R GENERAL CURIC FIT = 2.112

ENTING OUF TO CURIC TERM ALONE WITH 1 OF ...

TM RATING

I-32

0.024

THE PREDICTOR VARIABLE(X) IS GRE. AND THE CRITERION VARIABLE(Y) IS TO SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                                                                                                     0.A71
0.788
0.720
                                  SQUAR
                                                                                                                                                                    0.210
                                  X CUBE
                                                                                                     5.282
                                                                                                                                                                     1.946
        CORRELATION MATRIX
                                                                                                                                                                     0.952
                                                                                                                                                                                                                                  0.864
                                                                                                                                                                                                                                                                                                 0.186
                                                                                                     0.952
                                                                                                                                                                  1.000
                                                                                                                                                                                                                                  0.984
                                  X CURE
                                                                                                     0.186
                                                                                                                                                                   0.170
                                                                                                                                                                                                                                  0.163
                                                                                                                                                                                                                                                                                                  1.000
       FIRST, SECOND, AND THIRD DEGREE PULYNOMIALS.
           MULTIPLE R SOHARE = 0.035
MULTIPLE R = 0.186
N.D.F.1 = 7
N.D.F.2 = 37
F. FOR ANALYSTS OF VARIANCE ON R =
    F FOR ANALTSTS OF TOTAL STATE OF TOT
                                                                                                                                                                                                               1 - 327
    0.035
B WEIGHTS
2.131
     INTERCEPT CONSTANT = 3,426
      MULTIPLE R SCHARE = 0.035

MULTIPLE R = 0.187

N.D.F.1 = 2

N.D.F.2 = 3A

F FOR ANALYSIS OF VARIANCE ON R =
F FOR ANALYSIS OF VARIANCE ON R = 0.656
BFTA WEIGHTS
0.257 =0.075
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LCADINGS: 2ND COLUMN;
1 x 0.048 0.993
7 SOUAR =0.013 0.907
SOUARED BETA VEIGHTS
0.066 0.066
B WEIGHTS
2.944 =0.754
  2.944 =0.7A4
INTERCEPT CONSTANT =
                                                                                                                          3.272
MULTIPLE R SCHARE = 0.085

MULTIPLE R SCHARE = 0.085

MULTIPLE R = 0.293

N.O.F.1 = 3

N.O.F.2 = 35

F FOR ANALYSIS OF VARIANCE ON R =
                                                                                                                                                                                                            1.082
PFOR ANALYSIS OF VARIANCE ON R = 1.082

BET4 WEIGHTS
3.436 -9.087 5.97]

CONTRIBUTIONS TO FULTIFLE CORRELATION
AND REGRESSION FACTOR LUADINGS: 2ND COLUMN;
1 1 0.639 0.639
2 SQUAR -1.529 0.584
3 X CURF 0.974 0.560
T X CURF 0.974
SOUARED RETA WEIGHTS
                                                                                                                                                              0.560
11.607 80.7A8 35.655

R WFIGHTS

39.356 -84.877 48.463

INTERCEPT CONSTANT
                                                                                                                             2.981
```

```
ANDVA TABLE FOR POLYNCHIALS

REDUCTION DUF TO LINEAR FIT, WITH 1 DF # 0.035

PESIDUAL S.S. = 0.065 OF = 37 RESIDUAL M.S. = 0.026

PEOLETION DUF TO GENERAL QUADRATIC FIT WITH DF 2, # 0.035

REDUCTION DUF TO GENERAL QUADRATIC FIT WITH DF 2, # 0.035

REDUCTION M.S. = 0.065 OF = 36 RESIDUAL M.S. = 0.027

PEOLICTION DUF TO GUADRATIC TERM ALCNE, WITH 1 CF. = 0.001

PEOLICTION DUF TO GENERAL CUBIC FIT WITH DF 3, # 0.003

REDUCTION DUF TO GENERAL CUBIC FIT WITH DF 3, # 0.005

REDUCTION DUF TO GENERAL CUBIC FIT WITH DF 3, # 0.005

REDUCTION M.S. = 0.028

RESIDUAL S.S. = 0.015 OF = 35 RESIDUAL M.S. # 0.026
```

I-33

THE PREDICTOR VARIABLE(X) IS WHE, AND THE CRITERION VARIABLE(Y) IS TO SCORE.

```
TEST MEANS AND STANDARD DEVIATIONS
                                   0.496
                                                          0.194
0.193
0.155
            SQUAR
     3
            X CHBE
                                     0.176
                                     5.282
                                                           1.986
   CORRELATION MATRIX
                                                                                 0.930
0.986
1.000
                                    1.000
                                                           0.978
                                                                                                        0.403
            SQUAR
                                    0.978
                                                           1.000
                                                                                                        0.419
            X CUBE
                 ¥
                                    0.403
                                                           0.419
                                                                                  0.409
                                                                                                        1.000
                        ***********************
  FIRST, SECOND, AND THIRD DEGREE PULYNOMIALS.
   ***********************************
  MULTIPLE R SOMERE = 0.162
MULTIPLE R = 0.403
N.O.F.1 = 1
N.O.F.2 = 37
F FOR ANALYSIS OF VARIANCE ON R =
  BETE WEIGHTS
      0.403
  CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LGADINGS: 2ND COLUMN;
1 x 0.162 1.000
  SOURRED BETA VEIGHTS
  0.1AZ
R WEIGHTS
      4.127
  INTERCEPT CONSTANT =
                                                3.234
  ************
 MULTIPLE R SOHARE # 0.1/7

MULTIPLE R = 0.421

N.O.F.1 = 7

N.O.F.2 = 36

F FOR ANALYSIS OF VARIANCE ON R =

RETA MELOUSE
 F FOR ANALYSIS OF VARIANCE ON R = 3.867
BETA WEIGHTS
-0.161 0.876
CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS: 2ND COLUMN;
1 x -0.065 0.958
2 SOUAR 0.242 0.997
SOUARED RETA NEIGHTS
0.026 0.332
8 METGHTS
 -1.644 5.018
INTERCEPT CONSTANT .
                                               4.424
 MULTIPLE R SCHARE = 0.2
MULTIPLE R = 0.473
N.O.F.1 = 3
N.O.F.2 = 35
                                          0.224
 N.D.F.Z = 35
F FOR ANALYSIS OF VARIANCE ON R =
F FOR ANALYSIS DE FERRE SETTA WEIGHTS

-3.771 8.778 -4.667

CONTRIBUTIONS TO MULTIPLE CORRELATION
AND REGRESSION FACTOR LOADINGS: 2ND COLUMN;

X -1.520 0.852

P SOUAR 7.650 0.886

R X CUBF -1.907 0.864
3 X CUBF -1,90
SCUARED RETA WEIGHTS
14,222 75,433 21,785
B WEIGHTS
 #39.622 89.446 #56.064
Thtercept constant #
                                            9.017
```

```
ANOVA TARLE FOR POLYNCHIALS
  RESTOUAL S.S. = 0.836 DF = 37 RESTOUAL M.S. = FOR LINEAR FIT = 7,172
   REDUCTION DUF TO GENERAL QUADRATIC FIT WITH DF 2. .
 REDUCTION DUF TO GENERAL CURDING.

REDUCTION M.S. = 0.088

PESIDUAL S.S. # 0.823 OF = 36 RESIDUAL

F FOR QUADRATIC FIT = 3.867

REDUCTION DUF TO GUADRATIC TERM ALONE #ITH 1 OF: #

F FOR QUADRATIC TERM ALONE = 0.634
                                                                    36 RESIDUAL M.S. .
                                                                                                    0.014
 RECUICTION OUE TO GENERAL CUBIC FIT WITH OF 3, m
RECUICTION M. C. . 0.075
RESIDUAL S.C. . 0.776 DF . 35 RESID
                                                                    35 RESIDUAL M.S. .
                                                                                                               0.022
O OR GENERAL CURIC FIT .
```

3.362

ERICHETION DIP TO CURIC TERM ALONE WITH 1 DF = .

I - 34

APPENDIX J

JOB TASK CONDITIONAL AND JOINT FREQUENCIES BY RATING

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JOB TASK CONDITIONAL AND JOINT FREQUENCIES BY RATING

This Appendix will be devoted to presenting the Job Task Conditional and Joint Frequency Matrices by rating for the data collected on the Technical Proficiency Checkout Form (TPCF). The matrices are symetric about their main diagonal and therefore only the upper triangular portion will be presented. Denote an arbitrary entry in the i^{th} row and j^{th} column ($i=1,\ldots,8$) as $a_{i,j}$, read left to right and top to bottom for rows and columns respectively.

For the Job Task Conditional Frequency Matrices, the entries on the main diagonal, $a_{i,j}$ (i=1, ...,8), are the number of technicians CHECKED OUT on the (9-i)th job task, but NOT CHECKED OUT on an any more difficult task. The entry off the main diagonal, $a_{i,j}$ (i=1, ..., 8; j*i) is the number of technicians CHECKED OUT on the (9-j)th task given they are CHECKED OUT at most on the (9-i)th task. For example, from page J-4, there were 17 EM's who were CHECKED OUT at most on Job Task No. 7, i.e., they were NOT CHECKED OUT on Job Task No. 8. Of those 17 men 4, 13, 14, 17, 17, and 17 were CHECKED OUT on Job Task No.'s 6, 5, 4, 3, 2, and 1 res, pectively.

In the Job Task Joint Frequency Matrices, the entries on the main diagonal, $a_{i,i}$ (i=1, ..., 8), are the number of individuals CHECKED OUT on the ith task, regardless of their performance on any of the other tasks. The entries off the main diagonal, $a_{i,j}$ (i=1, ..., 8; j>i), are the number of technicians CHECKED OUT on both the ith and jth job task. For example, from page J-4, 75 EM's were CHECKED OUT on Job Task io. 5, 42 on Job Tasks No.'s 5 and 6, etc.



FLECTRICIAN'S MATE (EM) RATING (N = 97)

Conditional Frequencies

- 1	\sim	h		•	•	2.0
ı I	u	b	T	a	`	1 .

				•			
1	. 5	. 3	4	5	. 6	7	R
55	52	50	50	50	33	46	55
17	17	17	14	× 13	4	17	
. 6	4	5	5	6	6		
6	6	6	5	6	~		
3	3	3	3				
6	5	6	,				
0	0						
2							

The number of technicians not CHECKED OUT on any job task is 2

Joint Frequencies

	Job Task										
1	2	. 3	4	5	6	7、	8				
95	B7	87	77	75	43	63	55				
	87	54	73	70	40	63	5?				
		87	71	69	40	62	50				
. •			77	69	4 0	56	50				
				. 75	42	56	50				
		•	a.		43	35	33				
	•					63	46				
							55				



ELECTRONICS TECHNICIAN (ET) RATING (N = 173)

Conditional Frequencies

			Jol	b Task			
н	7	6	5	4	3	È	1
132 :	114	£ A	113	131	127	123	131
	15	5	6	15	15	11	15
		10	5	8	1 n	7	9
		,	2	5	2 .	1	5
				9	a '	4	, * A
		\			2	ĩ	, . 1
		N. Comment				0	0
		Ì					ج ا

The number of technicians not CHECKED OUT on any job task is 1

Joint Frequencies

	Job Task										
1	s	3	4	5	6	7	8				
158	146	160	163	126	95	129	131				
	147	146	143	117	A9	118	123				
,		164	159	124	95	124	127				
	,	1	165	125	94	129	131				
				126	88	108	113				
	1	هر			96	83	81				
-		*		,		129	114				
							132				



FIRE CONTROL TECHNICIAN (FT) RATING (N = 154)

Conditional Frequencies

			Jo	b Task			
· A	7	6	5	4	3	2	. 1
102	92	74	86	101	94	90	. 102
	Я	4	6	8	A	8.	Ą
		14	9	12	14	12	14
			15	14	12	10	14
				Ą	4	4	7
					2	1	5
						0	0
							,

The number of technicians not CHECKED OUT on any job task is 4

Joint Frequencies

	Job Task										
1	. 2	3	4	5	6	7	8				
145	125	134	141	115	93	1,00	102				
	125	124	121	101	H6	94	90				
		134	128	108	91	96	94				
	•		143	114	90	99	101				
				116	76	86	86				
					93	76	74				
			••		`	100	92				
			. 4				102				



INTERIOR COMMUNICATIONS ELECTRICIAN (IC) RATING (N = 58)

Conditional Frequencies

			Job	Task			
Ä	7	6	5	4	. 3	2	1
21	27	20	23	27	26	. 26	27
	6 °	1	6	5	6	6	6
		5	4	5	4	4	5
			8	7	6	4	в
	ı			• . 3	3	2	. 3
	,				4	3	4
	,			•		0	0

The number of technicians not CHECKED OUT on any job task is 1 $\,$

Joint Frequencies

Job Task

8		7	. 6	5	4	3	S	:
27))) 	33	26	41	44	. 49	45	57
26		. 32	2 5	36	38	45	45	
26		32	25	38	41	- 49		
27		32	53	3 7	44			
23		29	25	41				
20	•	21	54					
27		33						,
27								



RADARMAN (RD) RATING (N = 139)

Conditional Frequencies

	, ,		Job	Task			
н	• 7	6	5	4	3	5	1
			-	4		49	
56	, 7	6	25	14	4	. 5	26
	- 1	1	1	1	1	1	_ 1
		1	1	1	n	0	1
		,	70 -	31	30	16	70
		•	v	3	S	2	3
•					1	1	1
						2	. 2
							32

The number of technicians not CHECKED OUT on any job task is 3

Joint Frequencies

F		Task		`			
1	2	3 .	4	5	6	7	8
136	27	38	50	97	A	8	26
	27	23	10	21	3	1	5
		36	23	` 34	. 5	2	4
	•	,	50	47	7	6	. 14
-	•		•	97	a	8	. 25
	• .				, в	3	6 0
	·					8	7
		4		j.			26



RADIOMAN (RM) RATING (N = 137)

Conditional Frequencies

166	Taci	,
ממנו	1051	•

			•	•			
a	7	6	5	4 ,	3	2	1
13	10	Ų	13	13	13	12	13
	17	В	17	17	11	11	17
		29	24	55	28	24	59
			55	32	40	34	55
				4 .	3	3	4
				₹	5	4 ·	. 5
						n	0
							12

The number of technicians not CHECKED OUT on any job task is 2

Joint Frequencies

Job Task

1	2	3	4	5	6	7	8
135	, F 8	100	нв	113	46	27	13
	MIE	ಕರ	. 69	81	36	20	15
		100	76	. 91	40 3	21	13
			9.9	83	39	27	13
				113	45	27	13
					46	17	9
						27	10
							13

SONAR TECHNICIAN (ST) RATING (N = 152)

Conditional Frequencies

Job Task

Д	7	. 6	<i>△</i> 5	4	3	. 2	1
но	64	74	. 69	78	77	77	80
	13 -	11	Q	12	. 13	13	13
		- 18	16	10	15	14	18
	. 2	5.	27	15	10	10	. 24
				2	1	1	. 2
					, 3	. 3	3
						0	0

The number of technicians not CHECKED OUT on any job task is 5

Joint Frequencies

Job Task

8	7	6	5	4	3	2	1
80	77	103	118	115	119	118	144
77	76	98	99	108	118	118	
. 77	76	99	100	102	119		
٠ 78	76	94	101	117			,
69	69	92	121	•			
74	75	103					
64	77						
0.0							



TORPEDOMAN'S MATE (TM) RATING (N_= 39)

Conditional Frequencies

			Job	Task			
Ą	7	6	5	4	3	5	1
н	6	۴	8	8	8	8	8
	12	5	12	11	12	12	12
		3	3	3	3	3	3
	•		11	4	6	6	10
				5 ·	3	2	5
					n	0 _	0
	_					. 0	0
	-			*			0

The number of technicians not CHECKED OUT on any job task is 0

Joint Frequencies

	-		Job	Task			
1	5	3	4	5	6	7	8
34	3 0	31	31	33	14	18	А
	31.	31	27	29	14	ĭ8	8
		32	28	29	14	18	8
		0	31	26	14	17	8
				34	14	18	8
	•				14	_10	6
						18	6
							۵



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APPENDIX K
THE UTILITY OF THE WRE

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THE UTILITY OF THE WRE

The purpose of this section is to demonstrate the utility of the WRE as an estimator of individual performance. Essentially it is a useful type of estimator in that it is not dependent on a convention to be adopted for the case wherein a man did not work at particular job activity. As such the convention need only provide reliability ratios for those job activities for which the man being evaluated received $\Sigma UE = 0$ and $\Sigma UI = 0$ from his supervisor.

Each square in Table 1 represents a breakdown of Table E-1 into the number (and proportion) of men who did not work at a particular job activity and those men who received $\Sigma UE=0$ and $\Sigma UI=0$. For example, on job activity Number 1, 19 (19.6% of the EM's) received $\Sigma UE=0$ and $\Sigma UI=0$ and 6 (6.2% of the EM's) did not work at that job activity. The composite reliability values need then only be employed on 19.6% of the men in that rating and job activity rather then on 25% of the men as required by the SRE, PRE, and GRE. More significantly, in the case of RD's and RM's, for example, at most 59% (as compared to 99% for the SRE, PRE, and GRE) of the men in those ratings derive reliability ratios for some job activities from the composite reliability table. Clearly this is a significant improvement which should improve individual performance estimates. The statistical analyses reported in the main text of this paper verified this conjecture.

Derivation of the Weights Employed by the WRE

On the JPQ ANSWER SHEET is Appendix A, page A-4, in column (c) for each job activity the following question was answered by the supervisor on the man he was evaluating:

- QUESTION (c) Considering this man's <u>overall</u> performance, it is your opinion that the importance of this job activity, as a factor in determining the overall performance of this man, is best described as being:
 - 3. of central and primary importance
 - 2. a significant factor, but of secondary importance
 - of only moderate importance in estimating overall performance
 - 0. of little or no importance

The weights (w_i) for the ith job activity are determined by the formula:

If the supervisor recorded the i^{th} job activity as:

of central and primary importance, the weight $w_i = 1.0$

of secondary importance, the weight $w_i = .75$

of moderate importance, the weight $w_i = .5$

of little or no importance, the weight $w_i = .25$.



TABLE K-1

NUMBER AND PROPORTION OF TECHNICIANS IN EACH PROBLEM AREA

				Rati	ng			
Job Act.	EM NZ NW	ET NZ NW	FT NZ UW	IC UZ NW	RD NZ NW	PM NZ NW	ST NZ NW	TM NZ NW
1	19 6 0,196 0,062	38 0	51 5 0,331 0,032	0.155 0.000	46 4 0,331 0,029	34 19 0,248 0,139	39 6 0,257 0,039	6 0 0,154 0,000
2	16 7 0,165 0,072	47 0 0,272 0,00n	48 1 0,312 0,006	7 0 0,121 0.000	52 1 0,374 0.007	27 15 0.197 0.109	36 9 0,237 0,059	13 0 0.333 y.000
3	28 14 0,289 0,144	42 3	70 5 0,455 0,032	17 1 0;293 0,017	71 62 0,501 0,446	47 57 0.343 0.416	53 16 0,349 0,105	20 6 0,513 0,154
4	29 14 0,299 0,144	53 30 0,396 0,173	67 26 0,435 0,169	22 5 0,379 0.086	50 25 0,360 0,180	39 38 0.285 1.277	57 19 0;375 0;125	8 2 0,205 0,851
5	18 5	62 1 0,474 0.006	79 1 0,513 0,006	7 0 0.121 0.000	71 9	39 11 0,285 n,08n	58 4 0,382 0,026	10 1
6	38 12 0,392 0,124	83 33 0.480 0,191	79 36 0,513 0,234	24 11	59 41 0,424 0,295	42 46	. 58 32 0,382 0,211	15 5 0,385 0,128
Ž,	13 5	47 1 0,272 0,006	69 14 0,448 0,091	5 0 0;086 0.00M	50 87 0.367 0.626	54 74 0,394 0,540	49 27	23 2 n,590 0.n51
8	22 13 C,227 0,134	48 0	60 4 0,390 0,026	12 0 0:207 0.008	51 81 0,367 0,583	50 64 0,365 0.467	45 16 -0,296 0,105	18 4

Number of Men Each Rating

97 173 154 58 139 137 152 39

NZ = Number and Proportion of Technicians Who Received $\Sigma UE=0$ and $\Sigma UI=0$

NW = Number and Proportion of Technicians Who Did Not Work at that Job Activity



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5 AUTHORISI (First name, middle in tial, last name)			
Bernard A. Rafacz			
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6 REPORT DATE	78. TOTAL NO O	FPAGES	76. NO OF REFS
June 1973	180		21
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